Pharmacological interventions for recurrent abdominal pain in childhood (Review)


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Pharmacological interventions for recurrent abdominal pain in childhood (Review)

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ABSTRACT

Background
Between 4% and 25% of school-aged children at some stage complain of recurrent abdominal pain (RAP) of sufficient severity to interfere with their daily lives. When no clear organic cause is found, the children are managed with reassurance and simple measures; a large range of pharmacological interventions have been recommended for use in these children.

Objectives
To determine the effectiveness of pharmacological interventions for RAP in children of school age.

Search methods
We searched the Cochrane Central Register of Controlled Trials (CENTRAL), Ovid MEDLINE, Embase, and eight other electronic databases up to June 2016. We also searched two trials registers and contacted researchers of published studies.

Selection criteria
Randomised controlled trials involving children aged five to 18 years old with RAP or an abdominal pain-related functional gastrointestinal disorder, as defined by the Rome III criteria (Rasquin 2006). The interventions were any pharmacological intervention compared to placebo, no treatment, waiting list, or standard care. The primary outcome measures were pain intensity, pain duration or pain frequency, and improvement in pain. The secondary outcome measures were school performance, social or psychological functioning, and quality of daily life.

Data collection and analysis
Two review authors independently screened titles, abstracts, and potentially relevant full-text reports for eligible studies. Two review authors extracted data and performed a 'Risk of bias' assessment. We used the GRADE approach to rate the overall quality of the evidence. We deemed a meta-analysis to be not appropriate as the studies were significantly heterogeneous. We have consequently provided a narrative summary of the results.

Main results
This review included 16 studies with a total of 1024 participants aged between five and 18 years, all of whom were recruited from paediatric outpatient clinics. Studies were conducted in seven countries: seven in the USA, four in Iran, and one each in the UK, Switzerland, Turkey, Sri
Lanka, and India. Follow-up ranged from two weeks to four months. The studies examined the following interventions to treat RAP: tricyclic antidepressants, antibiotics, 5-HT4 receptor agonists, antispasmodics, antihistamines, H2 receptor antagonists, serotonin antagonists, selective serotonin re-uptake inhibitors, a dopamine receptor antagonist, and a hormone. Although some single studies reported that treatments were effective, all of these studies were either small or had key methodological weaknesses with a substantial risk of bias. None of these ‘positive’ results have been reproduced in subsequent studies. We judged the evidence of effectiveness to be of low quality. No adverse effects were reported in these studies.

Authors’ conclusions
There is currently no convincing evidence to support the use of drugs to treat RAP in children. Well-conducted clinical trials are needed to evaluate any possible benefits and risks of pharmacological interventions. In practice, if a clinician chooses to use a drug as a ‘therapeutic trial’, they and the patient need to be aware that RAP is a fluctuating condition and any ‘response’ may reflect the natural history of the condition or a placebo effect, rather than drug efficacy.

**PLAIN LANGUAGE SUMMARY**

Drug treatment of recurrent abdominal pain in children

**Review question**
Do medications improve the pain or other symptoms experienced by children with recurrent abdominal pain (RAP)?

**Background**
Reccurrent abdominal pain in childhood is a term used to describe unexplained episodes of tummy pain for which no cause can be found. The pain is often accompanied by other symptoms, such as diarrhoea or facial pallor. Some researchers have therefore classified different syndromes of unexplained pain according to these other associated symptoms. Recurrent abdominal pain is common in children. It is likely that the underlying cause or trigger differs among children.

**Study characteristics**
We searched the scientific literature worldwide up to June 2016 for research studies of drug treatments for children with RAP. We found 16 studies that met our criteria, examining antidepressants, antibiotics, antihistamines, antispasmodics, a dopamine receptor antagonist, and a hormone treatment. Fourteen studies compared drug treatments to a placebo, and two to usual medical care. The trials were carried out in seven countries: seven in the USA, four in Iran, one in the UK, one in Switzerland, one in Turkey, one in Sri Lanka, and one in India. The studies included a total of 1024 children aged between five and 18 years. All children were recruited from outpatient clinics. Follow-up lasted between two weeks and four months.

**Key results**
This review suggests there is no evidence for the use of medications to improve symptoms or the child’s quality of life. Consequently, if medications are prescribed, this should be done within a well-conducted clinical trial. If a medication is prescribed to a child with RAP, it must be remembered that RAP varies with time, and therefore any improvement or worsening may due to the natural history of the condition rather than a medication response.

**Quality of evidence**
Many of the studies had some weaknesses in their design and how they were reported, therefore the overall quality of the evidence for medications in RAP is low. The studies with better methods included few children and have not been reproduced by other researchers since.
### Summary of findings for the main comparison. Antispasmodics compared to placebo for recurrent abdominal pain

**Antispasmodics compared to placebo for recurrent abdominal pain**

**Patient or population:** school-aged children (5 to 18 years of age) with recurrent abdominal pain  
**Settings:** hospital paediatric outpatient clinics  
**Intervention:** antispasmodic drugs  
**Comparison:** placebo

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| Pain duration (mean pain duration, assessed at 4 weeks) | The mean duration of pain in the control group was 6.17 (± 11.61). | The mean duration of pain in the intervention group was 51.6 (± 23.74). | MD -25.4 (-35.5 to -15.3) | 120 (1) | ☠️☠️☠️☠️ Very low¹ | Asgarshirazi 2015  
No evidence of efficacy |
| Pain improvement (clinician judged, assessed at 2 weeks) | 9 of 21 children in the control group had an improvement in pain. | 15 of 21 children in the intervention group had an improvement in pain. | OR 3.33 (0.93 to 12.01) | 42 (1) | ☠️☠️☠️ Very low² | Kline 2001  
No evidence of efficacy |
| Pain frequency (episodes of pain in 4 weeks, assessed after 4 weeks) | The mean number of episodes of pain in the control group was 21.6 (32.4). | The mean number of episodes of pain in the intervention group was 10.3 (14). | MD 11.3 (2.4 to 20.1) | 132 (1) | ☠️☠️☠️ Very low³ | Narang 2015  
No evidence of efficacy |
| Pain improvement (self reported response to treatment, assessed at 4 weeks) | The response to treatment in the control group was 30.3%. | The response to treatment in the intervention group was 40.6%. | OR 1.6 (0.7 to 3.4) | 115 (1) | ☠️☠️ Low⁴ | Pourmoghaddas 2014  
No evidence of efficacy |

*The basis for the assumed risk (e.g., the median control group risk across studies) is provided in footnotes. The corresponding risk is based on the assumed risk in the comparison group and the relative effect of the intervention (and its 95% CI).  
CI: confidence interval; MD: mean difference; OR: odds ratio; SD: standard deviation
GRADE Working Group grades of evidence

**High quality:** Further research is very unlikely to change our confidence in the estimate of effect.

**Moderate quality:** Further research is likely to have an important impact on our confidence in the estimate of effect and may change the estimate.

**Low quality:** Further research is likely to have an important impact on our confidence in the estimate of effect and is likely to change the estimate.

**Very low quality:** We are very uncertain about the estimate.

---

1 Downgraded for low methodological quality due to single, small study; risk of bias from incomplete outcome data and differential loss of participants between groups. The placebo differed in preparation and dose timing compared to the intervention drug.

2 Downgraded for low methodological quality due to single, small study; risk of bias from selective outcome reporting and short follow-up.

3 Downgraded for low methodological quality due to single, small study; risk of bias from selective outcome reporting and the method of altering the drug doses.

4 Downgraded for single, small study. Not duplicated.

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Cochrane Library
BACKGROUND

Description of the condition

This review is an update of a previously published review in the Cochrane Library on 'Pharmacological interventions for recurrent abdominal pain (RAP) and irritable bowel syndrome (IBS) in childhood' (Huertas-Ceballos 2008a). Recurrent abdominal pain (RAP) is a common problem in paediatric practice. It has been suggested that 4% to 25% of school-aged children at some stage suffer from RAP that interferes with their activities of daily living (Abu-Arefeh 1995; Apley 1958; Faull 1986; Williams 1996; Øster 1972). Recurrent abdominal pain is often regarded as a relatively benign condition, but it is important to note the associated morbidity and anxiety it causes for children and carers. The condition is associated with school absences, hospital admissions, and, on occasion, unnecessary surgical intervention (Scharff 1997; Stickler 1979; Sterdal 2005; Walker 1998). Symptoms sometimes continue into adulthood (Apley 1975). The abdominal pain is commonly associated with other symptoms, including headaches, recurrent limb pains, pallor, and vomiting (Abu-Arefeh 1995; Apley 1958; Faull 1986; Hyams 1995; Stickler 1979; Stone 1970; Øster 1972).

It is generally accepted that RAP in children represents a group of functional gastrointestinal disorders with an unclear aetiology. Children suffer either chronic or recurrent gastrointestinal symptoms that are not explained by a structural, biochemical, or inflammatory process. Apley first sought to define the condition in the 1950s, suggesting that the diagnostic label should be based on the presence of at least three episodes of severe abdominal pain (often, but not necessarily, with associated systemic symptoms) over three months, with no established organic cause (Apley 1958). More recently, an international consensus definition with a symptom-based classification system with specific categories for paediatric presentations has been produced, known as the Rome III criteria (Rasquin 2006). We have used RAP throughout this review as an umbrella term to refer to the five categories included within this classification, which are: functional dyspepsia, irritable bowel syndrome (IBS), abdominal migraine, functional abdominal pain, and functional abdominal pain syndrome. It should be noted that the pain classification for each of the Rome III diagnoses is defined by at least one episode per week for at least two months; this varies from Apley's original definition of RAP (Apley 1958). The Rome classification is not based on known pathophysiological differences between the conditions, but rather on the constellation of clinical features. It remains unclear the extent to which separating children into these categories defines groups that are distinct clinical entities likely to respond differently to treatment. Nonetheless, this classification has been welcomed following the historical use of diverse terms, some implying causation. These include abdominal migraine (Bain 1974; Farquhar 1956; Hockaday 1992; Symon 1986), abdominal epilepsy (Stewens 1970), the irritable bowel syndrome in childhood (Stone 1970), allergic-tension-fatigue syndrome (Sandberg 1973; Sper 1954), neurovegetative dystonia (Peltonen 1970; Rubin 1967), functional gastrointestinal disorder (Drossman 1995), and the irritated colon syndrome (Harvey 1973; Painter 1964).

There is no consensus about which of the numerous proposed causal pathways result in the heterogeneous presentations of chronic abdominal pain, although it has been suggested that physical, emotional, and environmental factors may contribute to the manifestation of unexplained abdominal pain. When considering the diverse proposed mechanisms, it is unsurprising that a variety of treatments have been suggested. The treatment approaches can be grouped into pharmacological, dietary, or psychosocial (psychological, behavioural, or both). Reviews of the effectiveness of psychosocial and dietary interventions for RAP (Huertas-Ceballos 2008b; Huertas-Ceballos 2009), published in 2008 and 2009, have been updated as companions to this updated review (Abbott 2017; Newlove-Delgado in press, respectively).

Description of the intervention

A range of pharmacological treatments have been tried and tested for RAP in childhood: analgesics, dicyclomine (Edwards 1994), pizotifen (Christensen 1995; Symon 1995), herbal extracts (Zhang 1991), and many other drugs (Bain 1974; Worawattanakul 1999). A number of randomised controlled trials have reported on the use of peppermint oil for IBS in adults (Grigoleit 2005), the results of which have been interpreted as suggesting it to be a beneficial intervention. However, an earlier review reached no clear conclusion on efficacy due to poor methodological quality of the included studies (Pittler 1998). Other possible causal factors have been postulated, including food allergies (Poley 1973), reaction to food additives (Anonymous 1984), infectious agents like Helicobacter pylori, and parasitic infestation (Haldenberg 1995; Primelles 1990; Wardhan 1993).

How the intervention might work

The aetiology of abdominal pain-related functional gastrointestinal disorders is unclear. It has been suggested that visceral hypersensitivity (Di Lorenzo 2001; Van Ginkel 2001), autonomic dysfunction (Good 1995), and gut dysmotility may contribute, which may be initiated by an inflammatory, infective, traumatic, or allergic trigger (Mayer 2002; Mills 2001).

Conventional analgesics have been proposed to work by interrupting these abnormal physiological pain responses, which become pathological. Antispasmodics have been proposed to alter gut dysmotility, including peppermint oil, which has antispasmodic actions (Hills 1991). Serotonin (5-hydroxytryptamine) agonists may relieve symptoms by causing vasoconstriction and stimulation of the release of other vasoactive substances, thus inhibiting neurogenic inflammation; this has been found in migraine headaches (Goadsby 2000). Antidepressants treat the associated symptoms, and there is evidence of effectiveness in treating IBS in adults (Ruepert 2011).

Why it is important to do this review

Recurrent abdominal pain in children is very common, and in daily clinical practice there is no consensus about which treatments to offer patients. The approach is therefore inconsistent. It was important to do this review to establish if there is evidence for the effectiveness of pharmacological interventions in children with RAP. This review updates a previous review published in 2008 (Huertas-Ceballos 2008a). Companion reviews addressing the effectiveness of psychosocial, Huertas-Ceballos 2008b, and dietary interventions for RAP; Huertas-Ceballos 2009, are also being updated (Abbott 2017; Newlove-Delgado in press, respectively), so together they can guide clinicians, patients, and their families in treatment decisions.
OBJECTIVES
To determine the effectiveness of pharmacological interventions for RAP in children of school age.

METHODS
Criteria for considering studies for this review
Types of studies
Randomised controlled trials (RCTs).

Types of participants
Children aged five to 18 years with RAP or an abdominal pain-related functional gastrointestinal disorder, as defined by the Rome III criteria (Rasquin 2006).

Types of interventions
Any pharmacological intervention compared to placebo, no treatment, waiting list, or standard care.

Types of outcome measures
Primary outcomes
- Pain intensity
- Pain duration or pain frequency
- Improvement in pain

Secondary outcomes
The following were secondary outcomes, as measured by a validated tool.
- School performance
- Social or psychological functioning
- Quality of daily life

Search methods for identification of studies
Electronic searches
We ran the first literature searches in April 2013 and updated them in April 2014, March 2015, and again in June 2016. We searched the electronic databases and trial registers listed below.

- Cochrane Central Register of Controlled Trials (CENTRAL; 2016, Issue 5) in the Cochrane Library and which includes the Cochrane Developmental, Psychosocial and Learning Problems Specialised Register (searched 10 June 2016).
- Ovid MEDLINE In-Process & Other Non-indexed Citations and Ovid MEDLINE (1946 to current; searched 9 June 2016).
- Embase Ovid (1974 to current; searched 9 June 2016).
- CINAHL Healthcare Databases Advanced Search (Cumulative Index to Nursing and Allied Health Literature; 1981 to current; searched 9 June 2016).
- PsycINFO Ovid (1806 to current; searched 9 June 2016).
- ERIC ProQuest (Educational Resources Information Center; 1966 to current; searched 9 June 2016).
- BEI ProQuest (British Education Index; 1975 to current; searched 9 June 2016).
- ASSIA ProQuest (Applied Social Sciences Index and Abstracts; 1987 to current; searched 9 June 2016).
- AMED Healthcare Databases Advanced Search (Allied and Complementary Medicine; 1985 to current; searched 9 June 2016).
- LILACS (Latin American and Caribbean Literature in Health Sciences; lilacs.bvsalud.org/en; searched 9 June 2016).
- OpenGrey (opengrey.eu; searched 9 June 2016).
- ClinicalTrials.gov (clinicaltrials.gov; searched 9 June 2016).
- World Health Organization International Clinical Trials Registry Platform (WHO ICTRP; apps.who.int/trialsearch; searched 9 June 2016).

We revised the search terms from the original Cochrane RAP reviews (Huertas-Ceballos 2008; Huertas-Ceballos 2009a; Huertas-Ceballos 2009b); we consequently ran the searches for all available years. We used RCT filters where appropriate and imposed no language limits. We translated any non-English language studies identified so that they could be screened and considered for inclusion. The search strategies for each database are reported in Appendix 1.

Searching other resources
We used the Science Citation Index (Web of Science) for forward citation searching to identify papers in which the included reports had been cited, and we checked the reference lists of the included reports to identify any additional studies, including any ongoing or unpublished work. We also contacted researchers who have published studies in this field to ask for details of any relevant trials.

Data collection and analysis
Selection of studies
Two review authors (RAA, AB, TVND, or AEM) independently screened the titles and abstracts of studies identified by the search for relevance. We obtained the full-text reports of any paper that we judged to be potentially suitable for inclusion and then identified studies for inclusion against the Criteria for considering studies for this review. Any disagreements were resolved through discussion with a third review author (JTC).

Data extraction and management
Two review authors (RAA, AB, JTC, TVND, or AEM) extracted data and entered the data into Cochrane’s statistical software Review Manager 5 (Review Manager 2014). All review authors used the same data extraction form. We collected the following data.
Assessment of risk of bias in included studies

We considered the following domains when assessing the risk of bias of included studies:

- selection bias (random sequence generation and allocation concealment);
- performance bias (blinding of participants and personnel);
- detection bias (blinding of outcome assessment);
- attrition bias (incomplete outcome data);
- reporting bias (selective outcome reporting); and
- other sources of bias. We assessed all included studies for other sources of bias that may have altered the estimate of treatment effect, such as differential loss to follow-up, whether the data collection tools were valid, whether there was sufficient power in terms of appropriate sample size, whether baseline parameters were similar, and whether data analyses were appropriate.

Two review authors (RAA, AB, JTC, TVND, or AEM) independently assessed each study. We classified the risk of bias as low, high, or unclear based on the methods detailed in the Cochrane Handbook for Systematic Reviews of Interventions (Higgins 2011a). See Table 1. We considered a trial as having an overall low risk of bias if most of the above domains were assessed as at low risk of bias. We considered a trial as having an overall high risk of bias if several of the above domains were assessed as at high risk of bias or unclear risk of bias.

Measures of treatment effect

Continuous data

For continuous data (e.g. number of days of pain), we analysed means and standard deviations (SDs), if these were available or could be calculated and if there was no clear evidence of skewness in the distribution, and presented these with 95% confidence intervals (CIs).

If different scales were used to measure the same clinical outcome, we combined the standardised mean differences (SMDs) across the studies, and presented these with 95% CIs.

Dichotomous data

For dichotomous data (e.g. pain improved, yes or no), we analysed data using odds ratios (ORs) and presented these with 95% CIs.

Unit of analysis issues

No methods were used as we did not perform a meta-analysis. For methods archived for future updates of this review, please see Appendix 2 and our protocol (Martin 2014a).

Dealing with missing data

We first contacted the original investigators to request any missing data, but we received no responses. We did not impute any values. We did not feel it was relevant to carry out a sensitivity analysis, with and without the missing data, as it was not possible to conduct a meta-analysis. For methods archived for future updates of this review, please see Appendix 2 and our protocol (Martin 2014a).

We collected the proportions of participants for whom no outcome data were obtained and reported them in the assessment of Risk of bias in included studies. We have also provided this information for each study in the 'Risk of bias' tables, beneath the Characteristics of included studies tables. We explored the potential impact of missing data on the findings of the review in the Discussion section.

Assessment of heterogeneity

We anticipated finding considerable heterogeneity between included studies. We assessed clinical heterogeneity by examining the distribution of relevant participant characteristics (e.g. age, definition of RAP) and study differences (e.g. concealment of randomisation, blinding of outcome assessors, interventions or outcome measures).

We did not use our prescribed methods for assessing statistical heterogeneity as we did not perform a meta-analysis. For methods archived for future updates of this review, please see Appendix 2 and our protocol (Martin 2014a).

Assessment of reporting biases

We examined the report of each study to assess for selective outcome reporting. We assessed the study as adequate if it met the following criteria:

- the study protocol was available and all of the study's prespecified (primary and secondary) outcomes that were of interest to the review were reported in the prespecified way; or
- the study protocol was not available, but it was clear that the published reports included all expected outcomes, including those that were prespecified.

Data synthesis

Given the heterogeneity of drug classes and pain measurements used in the included studies, it was not possible to carry out a meta-analysis (DerSimonian 1986). We have therefore provided a narrative description of the results. For methods archived for future updates of this review, please see Appendix 2 and our protocol (Martin 2014a).

Summary of findings

We used the GRADE approach to assess the overall quality of the body of evidence for a specific outcome (Grade Working Group 2013). We used GRADEpro to assess and present the findings in a 'Summary of findings' table (GRADEpro GDT 2015). There were 11 comparisons in total, as we found 11 classes of drugs in the studies. We chose to produce a 'Summary of findings' table for antispasmodic drugs versus placebo (Summary of findings for the main comparison), as this was the most commonly investigated drug class in the studies, with a total of four studies. We presented pain, the primary outcome for this review, in the 'Summary of
findings' table. The measurement of pain varied between studies, as explained in the Types of outcome measures section.

We judged the studies included for each outcome using five criteria: risk of bias, indirectness, inconsistency, imprecision, and publication bias. We used limitations in the design and implementation to assess the overall risk of bias of included studies for each outcome; we downgraded an outcome if the majority of studies had unclear or high risk of bias. We assessed indirectness if a population, intervention, or outcome was not of direct interest to the review. Inconsistency was determined by the heterogeneity of results. If an outcome had a heterogeneity outcome of greater than 70%, we downgraded the quality of the outcome. Imprecision was assessed by the number of participants included in an outcome and by CIs. We downgraded an outcome when only a small number of participants could be included in the analysis or the analysis had wide CIs. Finally, we downgraded for publication bias if studies failed to report outcomes in the published manuscript or if there was a suspicion that null findings had not been published or reported (Schünemann 2011, section 12.2.2).

We gave each outcome a quality marking ranging from 'very low' to 'high'.

- High quality: "further research is unlikely to change our estimate of effect".
- Moderate quality: "further research is likely to have an important impact on our confidence in the estimate of effect and may change the estimate".
- Low quality: "further research is very likely to have an important impact on our confidence in the estimate of effect and is likely to change the estimate".
- Very low quality: "we are very uncertain about the estimate" (Balshem 2011).

Subgroup analysis and investigation of heterogeneity

We used no methods as we did not perform a meta-analysis. For methods archived for future updates of this review, please see Appendix 2 and our protocol (Martin 2014a).

Sensitivity analysis

We used no methods as we did not perform a meta-analysis. For methods archived for future updates of this review, please see Appendix 2 and our protocol (Martin 2014a).

RESULTS

Description of studies

For a full description of the main characteristics of the studies, including details on participants and setting, intervention aspects, and outcome measures, see Characteristics of included studies. See also Characteristics of excluded studies.

Results of the search

For this updated review, we chose to redesign the search strategy in order to include the recognised terms for different types of RAP, as defined by the Rome criteria (Rasquin 2006). Consequently, we ran our searches without date restrictions on each database. The results of the searching and screening process are shown in the PRISMA flow chart (Figure 1). We screened a total of 9649 titles and abstracts, 230 of which we carried forward for further screening at full-text. We excluded 214 reports and included 16 studies.
Figure 1. Study flow diagram.

14,700 records identified through electronic searches

1 record identified from handsearching

5052 duplicates removed

9649 records screened by title and abstract

9419 records excluded

214 full-text reports excluded:
203 excluded as irrelevant.
11 reports excluded with reasons:
- Ineligible study design (4 reports)
- Ineligible population (5 reports)
- Ineligible comparator (2 reports)

230 full-text reports assessed for eligibility

16 studies included in qualitative synthesis

0 studies included in quantitative synthesis (meta-analysis)
Included studies

For details, please see Characteristics of included studies.

We included 16 studies in the review (Asgarshirazi 2015; Bahar 2008; Collins 2011; Heyland 2012; Karabulut 2013; Karunanayake 2015; Khoshoo 2006; Kline 2001; Narang 2015; Pourmoghaddas 2014; Roohafza 2014; Sadeghian 2008; Saps 2009; See 2001; Symon 1995; Zybach 2016), three of which we included in the previous version of this review (Kline 2001; See 2001; Symon 1995).

Participants

The number of participants per study ranged from 12, in Zybach 2016, to 115 (Pourmoghaddas 2014; Roohafza 2014). Eight studies had more than 50 participants (Asgarshirazi 2015; Collins 2011; Karabulut 2013; Karunanayake 2015; Narang 2015; Pourmoghaddas 2014; Roohafza 2014; Saps 2009).

Location

Seven studies were based in the USA (Bahar 2008; Collins 2011; Khoshoo 2006; Kline 2001; Saps 2009; See 2001; Zybach 2016), four in Iran (Asgarshirazi 2015; Pourmoghaddas 2014; Roohafza 2014; Sadeghian 2008), one in the UK (Symon 1995), one in Switzerland (Heyland 2012), one in Turkey (Karabulut 2013), one in Sri Lanka (Karunanayake 2015), and one in India (Narang 2015).

Settings

All 16 studies were conducted in hospital paediatric outpatient clinics.

Study duration

The study duration ranged from two weeks (Collins 2011; Kline 2001; Sadeghian 2008) to four months (Symon 1995). See Characteristics of included studies for details of each study.

Interventions

Pharmacological interventions, drug classes used; tricyclic antidepressants (two studies) (Bahar 2008; Saps 2009), antibiotics (two studies) (Collins 2011; Heyland 2012), antimuscarinics (one study) (Karabulut 2013), 5-HT4 receptor agonists (one study) (Khoshoo 2006), antispasmodics (four studies) (Asgarshirazi 2015; Kline 2001; Narang 2015; Pourmoghaddas 2014), selective serotonin re-uptake inhibitors (one study) (Roohafza 2014), antihistamines (one study) (Sadeghian 2008), H2 receptor antagonists (one study) (See 2001), serotonin antagonist (one study) (Symon 1995), dopamine receptor antagonist (one study) (Karunanayake 2015), and a hormone (one study) (Zybach 2016).

Study design

Three studies were cross-over trials (See 2001; Symon 1995; Zybach 2016). One study had two intervention arms (Asgarshirazi 2015); the second intervention is reported in the accompanying review of dietary interventions (Newlove-Delgado in press). The remaining 12 studies were single-intervention, placebo-controlled trials (Bahar 2008; Collins 2011; Heyland 2012; Karabulut 2013; Karunanayake 2015; Khoshoo 2006; Kline 2001; Narang 2015; Pourmoghaddas 2014; Roohafza 2014; Sadeghian 2008; Saps 2009). All studies were randomised.

Outcomes

Primary outcome

All 16 included trials measured pain; the method of measurement varied. Tools included: visual analogue scale (VAS; 0 to 10), mean pain score and number of days of pain. Some trials used the parental or child report of adequate pain relief as the outcome measure.

Secondary outcomes

One study measured school attendance (Narang 2015). One study measured social and psychological functioning (Roohafza 2014); the authors assessed self rated "depression, anxiety, and somatisation" scores before and after treatment. Two studies measured quality of life (Bahar 2008; Karunanayake 2015). Other outcomes included: gastrointestinal symptoms scale and global assessment of well-being. See Characteristics of included studies for details of outcome measures used in studies.

Excluded studies

We examined 230 full-text reports and excluded 214. Of these, 203 reports were clearly irrelevant. Of the 11 remaining full-text reports, we excluded four due to ineligible study design (Christensen 1995; Cucchiara 1992; Dehghani 2011; Kaminski 2009), five due to ineligible populations (Di Nardo 2011; Everitt 2010; Lloyd-Still 1990; Van Outryve 2005; Yadav 1989), and two because of ineligible comparators (Grillage 1990; Xiao 2013). Please see Characteristics of excluded studies.

Risk of bias in included studies

For a summary, see Figure 2 and Figure 3.
Figure 2. Risk of bias summary: review authors' judgements about each risk of bias item for each included study.

<table>
<thead>
<tr>
<th>Study</th>
<th>Random sequence generation (selection bias)</th>
<th>Allocation concealment (selection bias)</th>
<th>Blinding of participants and personnel (performance bias)</th>
<th>Blinding of outcome assessment (detection bias)</th>
<th>Incomplete outcome data (attrition bias)</th>
<th>Selective reporting (reporting bias)</th>
<th>Other bias</th>
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<td>Asgarshirazi 2015</td>
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</table>
**Allocation**

Of the 16 included studies, we judged seven to be at low risk of bias for random sequence generation (Karunanayake 2015; Khoshoo 2006; Narang 2015; Pourmoghaddas 2014; Roohafza 2014; Sadeghian 2008; Symon 1995), and nine to be at unclear risk, largely because the authors did not explain their method of randomisation (Asgarshirazi 2015; Bahar 2008; Collins 2011; Heyland 2012; Karabulut 2013; Kline 2001; Saps 2009; See 2001; Zybach 2016).

We rated eight studies at low risk of bias because allocation concealment had been achieved (Collins 2011; Narang 2015; Pourmoghaddas 2014; Roohafza 2014; Sadeghian 2008; See 2001; Symon 1995; Zybach 2016), six studies at unclear risk of bias because the authors did not mention allocation concealment (Asgarshirazi 2015; Heyland 2012; Karunanayake 2015; Khoshoo 2006; Kline 2001; Saps 2009), and two studies at high risk of bias because the methods implied that allocation concealment had not been achieved (Bahar 2008; Karabulut 2013).

**Blinding**

We judged 11 studies with clear blinding of participants to be at low risk of bias (Collins 2011; Heyland 2012; Karunanayake 2015; Narang 2015; Pourmoghaddas 2014; Roohafza 2014; Sadeghian 2008; Saps 2009; See 2001; Symon 1995; Zybach 2016). We judged the risk of performance bias as unclear in three studies (Asgarshirazi 2015; Bahar 2008; Kline 2001), and high in two studies where blinding was not achieved (Karabulut 2013; Khoshoo 2006). These
judgements were mirrored in the risk of detection bias, as the primary outcomes were all self or parent reported and therefore relied on blinding.

Incomplete outcome data
Outcome data were largely complete, and we judged 13 of the 16 included studies to have a low risk of attrition bias (Bahar 2008; Collins 2011; Heyland 2012; Khoshoos 2006; Kline 2001; Narang 2015; Pourmoghaddas 2014; Roohafza 2014; Sadeghian 2008; Saps 2009; See 2001; Symon 1995; Zybach 2016). We judged one study to have an unclear risk of attrition bias, as the authors did not report the numerical outcome data or the number of participants completing the study, but rather reported results as a percentage without a total (Karabulut 2013). We judged one further study to have an unclear risk of attrition bias due to insufficient information in the conference abstract and no further information from personal communication with the study authors (Karunanyake 2015). We judged one study to be at a high risk of attrition bias, as there were differential losses to follow-up between the groups (Asgarshirazi 2015).

Selective reporting
We considered the risk of reporting bias to be unclear in two studies (Collins 2011; Khoshoos 2006), and high in three studies (Bahar 2008; Kline 2001; Narang 2015), because the authors did not report results for all the outcomes mentioned in the methods section. We rated the risk of reporting bias as low in 11 studies (Asgarshirazi 2015; Heyland 2012; Karabulut 2013; Karunanyake 2015; Pourmoghaddas 2014; Roohafza 2014; Sadeghian 2008; Saps 2009; See 2001; Symon 1995; Zybach 2016).

Other potential sources of bias
We judged the risk of other potential sources of bias as high in 10 studies (Asgarshirazi 2015; Bahar 2008; Collins 2011; Karabulut 2013; Narang 2015; Sadeghian 2008; Saps 2009; See 2001; Symon 1995; Zybach 2016), and unclear in four studies (Heyland 2012; Karunanyake 2015; Khoshoos 2006; Kline 2001), due to lack of details on baseline characteristics, lack of power calculations, lack of details on how outcomes were measured, and insufficient washout periods in cross-over trials. We rated two studies at low risk of other potential sources of bias (Pourmoghaddas 2014; Roohafza 2014).

For each domain above, where there was insufficient information to make a judgement of high or low risk of bias, we wrote to all trial authors for clarification (Abbott 2014a [pers comm]; Abbott 2014b [pers comm]; Martin 2014b [pers comm]; Martin 2014c [pers comm]; Martin 2016a [pers comm]; Martin 2016b [pers comm]; Martin 2016c [pers comm]; Martin 2016d [pers comm]; Martin 2016e [pers comm]; Martin 2016f [pers comm]; Martin 2016g [pers comm]). As we received limited response, we assigned ratings of unclear risk of bias for these studies in these domains.

Effects of interventions
See: Summary of findings for the main comparison Antispasmodics compared to placebo for recurrent abdominal pain

Comparison 1. Tricyclic antidepressants compared to placebo
Two studies, Bahar 2008 and Saps 2009, evaluated amitriptyline compared to placebo in 123 children with functional gastrointestinal disorders as defined by the Rome II criteria (Rasquin-Weber 1999).

Bahar 2008 assessed multiple outcomes measuring pain (using a pain-rating scale and VAS), quality of life, and IBS symptoms for up to 13 weeks. The quality of life tool had been validated in adults (Patrick 1997). The authors reported no significant difference in pain between the two groups, but provided no data to support this. The mean quality of life scores were 109.4 and 127.5 at baseline for the intervention and control groups, respectively. At 13 weeks' follow-up, they were 126.2 and 129.8 for the intervention and control groups, respectively. The authors suggested that a higher proportion of participants in the intervention group showed a 15% increase in quality of life at 13 weeks compared to the control group (P = 0.002). The clinical significance of this is unclear. Of note, the two groups had a 14% difference of mean quality of life scores at baseline. This may have been a post-hoc analysis, as it is not mentioned in the study. The study authors did not provide the standard deviations for these data. We wrote to the study authors to request further information but received no response (Abbott 2014b [pers comm]).

Saps 2009 used author-defined outcomes to evaluate improvement in pain from amitriptyline compared to placebo at four weeks' follow-up. They found no difference between the intervention and control groups when assessing the self-reported outcomes of “how well did the medication relieve your pain” and “overall how do you feel your problem is”. Out of the 46 children in the amitriptyline group, 23 experienced an improvement in pain, compared to 22 out of 44 in the control group (OR 1.05, 95% CI 0.45 to 2.45).

We were unable to perform a meta-analysis due to the use of different outcome measures and lack of numerical data. The GRADE quality rating for this outcome was very low, due to the small number of studies with methodological flaws.

Comparison 2. Antibiotics compared to placebo
Two studies, Collins 2011 and Heyland 2012, evaluated antibiotics compared to placebo in 112 children with Rome II criteria for IBS, functional dyspepsia, functional abdominal pain, or abdominal migraine (Rasquin-Weber 1999), and RAP (Apley 1958).

Collins 2011 assessed the effectiveness of rifaximin compared to placebo. The authors used 10 self reported outcomes on a VAS and a pain questionnaire to evaluate improvement in pain. The 10 self reported outcomes included: incomplete evacuation, abdominal pain, diarrhoea, constipation, urgency to pass stool, passage of mucus, straining, and faecal soiling. The authors stated that there were no differences between the intervention and control groups for any of the outcomes. The authors did not present their data to support this conclusion. We contacted the authors but received no response (Abbott 2014a [pers comm]).

Heyland 2012 also reported the use of antibiotics to treat RAP. They compared co-trimoxazole to placebo on the mean pain index. This was a 10-point VAS measuring pain severity for each group. The mean pain index changed from 6.9 pre-treatment to 4.1 post-treatment in the intervention group, a mean change of −2.9. In the placebo group, the mean pain index changed from 7.4 pre-treatment to 3.0 post-treatment, a mean change of −4.4. The authors found no difference in scores on the mean change of pain index between the two groups. No raw data were given. We
contacted the authors but received no response (Martin 2016c [pers comm]).

We were unable to perform a meta-analysis due to lack of numerical data. These papers offer no evidence of the effectiveness of antibiotics to treat functional gastrointestinal disorders. The GRADE quality rating for this outcome was very low, due to lack of reliable outcome data to assess precision of treatment effect, and small and single studies for any drug intervention.

**Comparison 3. Antimuscarinic drugs compared to usual care**

Karabulut 2013 compared trimetobutine to usual care in 78 children with IBS as defined by Rome III criteria (Rasquin 2006). The authors reported a self defined outcome of pain improvement, “adequate relief”, as assessed by the parents. There was an improvement in 37 out of 39 children treated with trimetobutine and in eight out of 39 children treated with usual care (P < 0.0001; OR 71.7, 95% CI 14.2 to 362.7). However, it is important to note that the intervention was not blinded and that the outcome measure was parental assessment, therefore performance and detection bias alone could explain the results. Consequently, the GRADE quality rating was very low, and the results should be interpreted with caution.

**Comparison 4. 5-HT4 agonists compared to usual care**

Khoshoo 2006 compared tegaserod with usual care in 48 children with constipation-predominant IBS. The authors reported 14 out of 21 children with good pain reduction in the treatment group compared with 5 out of 27 children in the control group (P < 0.05 (exact P value not reported in paper); OR 8.8, 95% CI 2.3 to 33.2). This suggests that tegaserod may be effective in relieving abdominal pain in constipation-predominant IBS. Due to the small number of children participating in the study and the risk of performance and detection bias, this result should be interpreted with caution. Importantly, it is not clear if this was a post-hoc analysis. The GRADE quality rating for this comparison was therefore very low.

**Comparison 5. Antispasmodics compared to placebo**

Four trials compared antispasmodics to placebo in 377 children with functional abdominal pain (Asgarshirazi 2015; Kline 2001; Narang 2015; Pourmoghaddas 2014). For a summary of antispasmodics, see Summary of findings for the main comparison.

Asgarshirazi 2015 compared peppermint oil, placebo and a symbiotic Lactol (containing a probiotic and fructo-oligosaccharide) in a three-arm RCT. One hundred and twenty children with functional gastrointestinal disorders based on Rome III criteria were included in the study, but those with abdominal migraine were excluded. Changes in severity, duration, and frequency of pain and any adverse effects were reported at four weeks. The results for the symbiotic intervention are reported in an accompanying dietary review (Newlove-Delgado in press). Thirty-four children completed the study and were analysed in the peppermint oil intervention group and 25 in the placebo group. The authors reported the intervention group compared to placebo: mean difference (MD) in pain duration -25.4 minutes/day (95% CI -35.5 to -15.3), MD in frequency of pain -1.4 episodes/week (95% CI -2.0 to -0.8), and MD in severity of pain -1.1 on a numerical rating scale (95% CI -1.8 to -0.4). Notably, this trial was at risk of bias from incomplete outcome data and differential loss of children between groups (38% drop out in the placebo group and 15% in the intervention group). The placebo was different in preparation and dose timing compared to the intervention drug.

Kline 2001 also compared peppermint oil capsules with placebo in 42 children with IBS as defined by the Rome criteria (Rasquin-Weber 1999). The authors reported that, at two weeks, 15 out of 21 children in the peppermint group had a clinician-judged improvement in pain compared to 9 out of 21 in the placebo group (OR 3.33, 95% CI 0.93 to 12.01). The authors reported the 15-item Gastrointestinal Symptom Rating Scale (Svedulend 1988), which measures frequency, duration, and impact on daily life, as showing no difference between groups, but do not provide the data in the study. We contacted the trial authors but received no response (Martin 2016f [pers comm]). The authors reported that daily diaries completed by children showed significantly lower mean pain severity in the peppermint oil group. The authors provided neither data nor explanation of how the analysis of the daily diaries was carried out, although the P value for this comparison was reported to be less than 0.03. No side effects were reported in either group. This study therefore provides insufficient evidence to support the use of peppermint oil in the treatment of RAP.

Narang 2015 compared drotaverine to placebo in a RCT with 132 children with RAP as defined by Apley (Apley 1958). They assessed children’s pain severity and frequency, school attendance, and parental judgement of well-being (Likert scale) after four weeks. The authors provided no results for pain severity, only pain frequency, reporting the mean number of episodes of pain in four weeks and number of pain-free days in four weeks. They found a mean of 10.3 (SD = 14) episodes of pain in the 64 children receiving drotaverine and a mean of 21.6 episodes (SD = 32.4) in the 60 children receiving placebo. The MD between the groups in episodes of pain was 11.3 (95% CI 2.4 to 20.1). The mean number of pain-free days was 17.4 (SD = 8.2) in the intervention group and 15.6 (SD = 8.7) in the placebo group; MD 1.8, 95% CI -1.2 to 4.8. The authors reported the mean number of school days missed as 0.24 (SD = 0.85) in the drotaverine group and 0.71 (SD = 1.59) in the placebo group. The MD between these two groups was 0.46 (95% CI 0.01 to 0.91).

Pourmoghaddas 2014 randomised 115 children to mebeverine versus placebo, and found no effectiveness of mebeverine in treating functional abdominal pain in children. Self reported treatment response rates in the mebeverine and placebo groups were 40.6% and 30.3%, respectively at four weeks' postintervention (P = 0.469; OR 1.6, 95% CI 0.7 to 3.4) and 54.2% and 41.0%, respectively at 12 weeks' postintervention (P = 0.416; OR 1.7, 95% CI 0.8 to 3.6). This was an intention-to-treat analysis; the authors used last observed carried forward to substitute for missing data. The authors found no difference between the groups in scores on the physician-rated change of the Clinical Global Improvement-Severity scale, at four or 12 weeks' postintervention (NIMH 1985).

A meta-analysis of studies evaluating antispasmodic drugs was not possible due to the heterogeneity of the interventions and variation in outcome measures. The overall GRADE quality rating for evidence evaluating this comparison was very low.

**Comparison 6. Selective serotonin re-uptake inhibitors (SSRIs) compared to placebo**

Rooheza 2014 compared citalopram and placebo in 115 children with functional abdominal pain as defined by Rome III criteria.
Pharmacological interventions for recurrent abdominal pain in childhood (Review)

Comparison 7. Antihistamines compared to placebo
Sadeghian 2008 randomised 29 children with functional abdominal pain as defined by Rome II criteria to cyproheptadine or placebo (Rasquin-Weber 1999). They reassessed the outcome measures of pain intensity, pain frequency, and global self judgement of improvement in symptoms at two weeks. Pain intensity and pain frequency improved in 9 out of 15 children in the treatment group and 2 out of 14 children in the placebo group (OR 9.0, 95% CI 1.46 to 55.48). While this result suggests effectiveness, it should be interpreted with caution due to the very low GRADE quality rating of the evidence. The study was at risk of bias from the small sample size (imprecision of treatment effect), used non-validated measurement tools, and the findings have not been reproduced in other studies.

Comparison 8. H2 receptor antagonists compared to placebo
One study, See 2001, compared famotidine versus placebo in a randomised cross-over trial of 25 children with RAP, as defined by Apley 1958. The authors provided results from the first period of the trial during which 12 children received famotidine and 13 children received placebo. They reported that 8 out of 12 children receiving famotidine in the first period improved compared to 2 out of 13 children receiving placebo in the first period (OR 11.0, 95% CI 1.6 to 75.5). They also reported scores on an "abdominal pain score" (APS), which included three components: pain frequency, pain severity, and a peptic index score, which evaluates a number of symptoms, including nausea, vomiting, and nocturnal wakening. This appears not to be a validated tool used in other studies. The authors reported no statistically significant difference in scores between famotidine and placebo, considering all the data from both periods. They reported finding an improvement in APS on famotidine (mean 3.37 (SD ± 3.53)) and on placebo (mean 1.66 (SD ± 2.7)). The MD in improvement on APS is 1.71. We were not able to provide confidence intervals around this difference due to insufficient data from the study authors. This is a cross-over trial, and we did not impute values for the correlation coefficient. The authors of the previous review were unsuccessful in contacting the trial authors (Huertas-Ceballos 2008a). The GRADE quality rating was very low due to the lack of primary data to confirm findings, no washout period between cross-over of interventions, and the use of an invalidated tool.

Comparison 9. Serotonin antagonists compared to placebo
Symon 1995 reported the effects of pizotifen versus placebo in a restricted subgroup of 16 children with RAP. Although the children satisfied the standard definition of RAP (Apley 1958), they also had to report associated facial pallor, and either one first-degree relative or two second-degree relatives with a history of migraine or throbbing headache to be included in the study. This was a cross-over study comparing the MD in the number of days on which children reported abdominal pain while taking pizotifen and placebo. The children reported a MD of 8.21 (95% CI 2.93 to 13.48) fewer days of pain while taking the intervention drug. The authors also reported the MD on the "Index of Severity", which was -16.21 (95% CI -26.51 to -5.90), and the MD on the "Index of Misery", which was -56.07 (95% CI -94.07 to -18.07). These appear not to be validated tools, but were judged by the trial authors to measure improvement. The study authors reported P values of 0.005 and 0.007 for these two comparisons, respectively. The authors reported that the trial was stopped early as a result of an interim analysis conducted when their initial supplies of drug preparations reached their expiry dates. They did not provide details of the size of the sample they had originally planned to include in the study. These findings have not been replicated in other studies. The GRADE rating was very low for this comparison.

Comparison 10. Dopamine receptor antagonist compared to placebo
Karunayake 2015 compared domperidone to placebo in 89 children aged five to 12 years with abdominal pain-predominant functional gastrointestinal disorders as defined by Rome III criteria (Rasquin 2006). Two primary outcomes were specified: cure and improvement. Cure was abdominal pain less than 25 mm on the VAS and no impact on daily activity. Improvement was pain relief and sense of improvement recorded on the Global Assessment Scale. No further information on how these primary outcomes were defined, when they were measured, or who assessed them was provided. The study is available as a conference abstract, and despite having written to the authors twice (Martin 2016a [pers comm]; Martin 2016d [pers comm]), we have no further published or unpublished information. The authors state that there was no difference in "cure" between the two groups, but provided no data to support this. They state that 37 out of 47 (78.7%) of the children treated with domperidone had "significant improvement" and 25 out of 42 (59.5%) in the placebo group had "significant improvement". This gives an OR for improvement of 2.52 (95% CI 0.99 to 6.39). Due to the lack of information about this outcome measurement and the limited possible 'Risk of bias' assessment, this result should be interpreted with caution. Regarding secondary outcomes, the domperidone group reported a significant reduction in abdominal pain severity (70.84% versus 48.18%) and improvement on the motility index (29.3% versus 8.6%) after intervention. No such difference was seen in improvement of quality of life and family impact. It was also unclear how these secondary outcomes were defined, measured, or who assessed them. The GRADE rating was very low for this comparison.

Comparison 11. Hormone treatment compared to placebo
Zbybach 2016 reported the therapeutic effect of melatonin compared to placebo in 12 children aged 11 to 16 years with functional abdominal pain as defined by Rome III criteria (Rasquin 2006). This was a cross-over study with no washout period. No power calculation was performed, but due to the low number of
children, the study may have been underpowered. The authors found no difference in pain response reported by those treated with melatonin compared to placebo: OR 0.71 (95% CI 0.14 to 3.58). The authors also reported no change in mean sleep duration: melatonin group 9.9 (SD ± 3.53) hours; placebo group 9.41 (SD ± 2.7) hours. The GRADE rating for this comparison was very low.

**D I S C U S S I O N**

**Summary of main results**

Recurrent abdominal pain is an extremely common condition in childhood, and survey data in the USA suggest that many paediatricians use drug treatment for RAP (Edwards 1994). The most striking result of this review is the paucity of good-quality, placebo-controlled trials for all of the drugs that have been recommended for use in children with RAP.

In 2006, the paediatric Rome III criteria were devised to classify paediatric functional gastrointestinal disorders (Rasquin 2006). Diagnoses included within this classification comprised five categories defined on the basis of symptom profiles: IBS, functional dyspepsia, functional abdominal pain, abdominal migraine, and functional abdominal pain syndrome. The consistent symptom in each of these profiles is unexplained abdominal pain, which, prior to the development of this classification, was the complaint used to identify patients. It remains unclear the extent to which separating children into these subgroups defines patients who have different psychological or pathophysiological mechanisms underlying their symptoms or predicting their treatment response.

The studies with positive results are either small, single studies that have not been replicated or are larger studies with methodological flaws. Therefore, there is insufficient evidence to recommend any drug treatment.

There was no evidence from Bahar 2008 or Saps 2009 to suggest that amitriptyline is effective in treating RAP. Similarly, there was no evidence that antibiotics have a role in the treatment of RAP (Collins 2011; Heyland 2012). A single study reported that trimethubine (an antimuscarinic drug) was extremely effective in treating RAP (Karabulut 2013), but this was based on reports of symptoms by parents who were aware of whether their children were either receiving active treatment or not. A single, small study with a high risk of bias evaluated tegaserod (a 5-HT4 agonist) (Khoshoo 2006). The findings of this study suggest that tegaserod may be effective, but further evidence is required before it can be recommended. We found two trials of peppermint oil: one trial showed no efficacy, but the small numbers may mean it was underpowered (Kline 2001), and the second trial had key methodological flaws requiring that it be interpreted with caution (Asgarshirazi 2015). There is no current evidence to recommend peppermint oil. Narang 2015 suggested benefit from drobameter in some of the outcomes measured; others were either unreported or the children received no benefit from the intervention. Pourmoghaddas 2014 found no effectiveness of mebeverine in treating functional abdominal pain in children. Similarly, Roohafza 2014 compared citalopram and placebo and found no difference in treatment response rate between the groups. Sadeghian 2008 and See 2001 suggested that antihistamines and H2 receptor antagonists respectively may be effective in treating RAP. However, these results should be interpreted with caution due to risk of bias in the studies, small sample numbers, and therefore imprecision of estimates.

In addition, the results of these single studies have not been reproduced. Symon 1995 reported the effects of pizotifen versus placebo in a subgroup of children with RAP fulfilling the definition of abdominal migraine. In the 14 children studied, the mean number of days of pain was less in the pizotifen group. The results of this small study, which was stopped early as this interim analysis was conducted when the drug supply had expired, has not been replicated in the last 20 years. Karunanayake 2015 published in abstract form a study of domperidone versus placebo; there was insufficient information on the outcomes measured and the quality of the study to conclude if domperidone may be effective in treating RAP. Zybach 2016 investigated melatonin compared to placebo in a small number of children in a cross-over trial and found no efficacy. For a summary of these results, please see **Summary of findings for the main comparison**.

**Overall completeness and applicability of evidence**

This review highlights some issues concerning the overall completeness and applicability of the evidence of benefits and harms of pharmacological interventions for children and adolescents with RAP: the lack of trials conducted in specific subgroups of RAP as defined by the Rome III criteria (Rasquin 2006); the lack of trials assessing the same class of drug; and the lack of sustained intervention and follow-up beyond the period of intervention.

The majority of studies included children within the broad diagnosis of RAP, which meant that children could be presenting with a variety of RAP classifications such as IBS, functional abdominal pain, or functional dyspepsia. This meant that it was not possible to investigate whether particular classes of drugs benefited particular subgroups of RAP more than other subgroups.

Lastly, most of the interventions were relatively short in duration (two to six weeks), and very few had medium- or long-term follow-up, which limits the ability to assess whether any benefits are sustained in the long term.

**Quality of the evidence**

The overall quality of this evidence was low.

**Potential biases in the review process**

The present systematic review has many strengths. We developed a protocol for this review according to guidance provided in the Cochrane Handbook for Systematic Reviews of Interventions (Higgins 2011b). We published our protocol before we embarked on the review itself (Martin 2014a). We conducted extensive searches of relevant databases and checked forward and backward citations of all included studies. We also contacted authors of included studies for additional data when the presented data were insufficient or data were missing, to maximise our ability to pool data on comparable outcomes within comparable intervention types. Two review authors, working independently, selected trials for inclusion and extracted data. Disagreements were resolved by discussion between team members. We assessed the risk of bias in all trials according to the recommendations provided in the Cochrane Handbook for Systematic Reviews of Interventions (Higgins 2011a).

We did not include studies that had a mix of children, adolescents, and young adults when it was not possible to separate the data for children younger than 18 years of age. Likewise, we did not include
studies that did not specify recruiting children or adolescents, and which presented mean ages of the population sample exceeding 20 years of age. Both these issues raise the possibility of bias in our review process, as we did not write to these authors asking whether or not they collected data for children younger than 18 years of age. However, we believe this potential bias is not likely to have changed our conclusions.

Agreements and disagreements with other studies or reviews
The previous version of this review, Huertas-Ceballos 2008a, included only three studies (Kline 2001; See 2001; Symon 1995). This update includes 16 studies and reached a similar conclusion: there is no evidence to support the use of drugs to treat RAP or functional gastrointestinal disorders in children. Another Cochrane review evaluating the effectiveness of antidepressants in pain-related functional abdominal disorders in children also agrees with this conclusion (Kaminski 2009).

AUTHORS' CONCLUSIONS
Implications for practice
Overall, this review provides only extremely weak evidence for the efficacy of some pharmacological agents in children with RAP. The lack of clear evidence of effectiveness for any drug suggests that there is little reason for their use outside of well-conducted clinical trials. Clinicians may choose to prescribe drugs to children whose symptoms are severe and who have not responded to simple management. However, when using drugs as a ‘therapeutic trial’, clinicians need to be aware that RAP is a fluctuating condition and any ’response’ may reflect the natural history of the condition or a placebo effect, rather than drug efficacy.

Implications for research
The pathogenesis of RAP in children remains unclear (Hyams 1998). There is an obvious need for further studies to be conducted to elucidate this aetiology. It may be that the complaint of pain is a unifying manifestation for a wide variety of causal pathways and triggers relating to psychological and physical processes. It is unlikely that RAP is a single disease entity. Further trials are therefore needed not only to guide the management of children with RAP, but also to validate the usefulness of suggested classifications (Rasquin 2006).

ACKNOWLEDGEMENTS
We acknowledge the work of the original review authors: Angela Huertas-Ceballos, Stuart Logan, Cathy Bennett, Sarah See, Colin Macarthur, and Morris Zwi. We would also like to acknowledge the work of the Cochrane Developmental, Psychosocial and Learning Problems Review Group, the editors, statistician, and peer reviewers who commented on early drafts of this review.
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Kline 2001 {published data only}

Pourmoghaddas 2014 {published data only}

Roohafza 2014 {published data only}

Sadeghian 2008 {published data only}

Saps 2009 {published data only}

Symon 1995 {published data only}

Zybach 2016 {published data only}

References to studies excluded from this review

Christensen 1995 {published data only}
Christensen MF. Letter: Double blind placebo controlled trial of pizotifen syrup in the treatment of abdominal migraine.

REFERENCES
Archives of Disease in Childhood 1995;73(2):183. [DOI: 10.1136/adc.73.2.183-a]

Cucchiara 1992 [published data only]

Dehghani 2011 [published data only]

Di Nardo 2011 [published data only]

Everitt 2010 [published data only]

Grillage 1990 [published data only]

Kaminski 2009 [published data only]

Lloyd-Still 1990 [published data only]

Van Outryve 2005 [published data only]

Xiao 2013 [published data only]

Yadav 1989 [published data only]

Additional references
Abbott 2014a [pers comm]
Abbott R. Study information [personal communication]. Email to: B Collins 14 March 2014.

Abbott 2014b [pers comm]
Abbott R. Study information [personal communication]. Email to: R Bahar 25 April 2014.

Abbott 2017

Abu-Arafeh 1995

Anonymous 1984

Apley 1958

Apley 1975

Asgari-Shirazi 2016 [pers comm]
Asgari-Shirazi M. Re: study information [personal communication]. Email to: A Martin 1 November 2016.

Bain 1974
Pharmacological interventions for recurrent abdominal pain in childhood (Review)

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Pharmacological interventions for recurrent abdominal pain in childhood (Review)

Copyright © 2017 The Cochrane Collaboration. Published by John Wiley & Sons, Ltd.

**Huertas-Ceballos 2009**

**Hyams 1995**

**Hyams 1998**

**Kovacs 1985**

**Martin 2014b [pers comm]**
Martin A. Tegasterod in IBS [personal communication]. Email to: V Khoshoosh 16 August 2014.

**Martin 2014c [pers comm]**
Martin A. RAP paper in Gastroenterology [personal communication]. Email to: M Saps 6 August 2014.

**Martin 2016a [pers comm]**
Martin A. Domperidone in RAP [personal communication]. Email to: A de Silva 6 July 2016.

**Martin 2016b [pers comm]**
Martin A. Study information [personal communication]. Email to: C Friesen 29 June 2016.

**Martin 2016c [pers comm]**
Martin A. Antibiotics and RAP in children [personal communication]. Email to: C Braegger 7 December 2016.

**Martin 2016d [pers comm]**
Martin A. Domperidone in RAP [personal communication]. Email to: A de Silva 27 November 2016.

**Martin 2016e [pers comm]**
Martin A. Trimebutine + IBS [personal communication]. Email to: O Beser 7 December 2016.

**Martin 2016f [pers comm]**
Martin A. Peppermint oil paper for IBS [personal communication]. Email to: R Kline 7 December 2016.

**Martin 2016g [pers comm]**
Martin A. Study information [personal communication]. Email to: M Asgarshirazi 29 September 2016.

**Mayer 2002**

**Milla 2001**

**Newlove-Delgado in press**

**NIMH 1985**

**Painter 1964**

**Patrick 1997**

**Peltonen 1970**

**Pittler 1998**

**Poley 1973**

**Primelles 1990**

**Rasquin 2006**

**Rasquin-Weber 1999**
Rasquin-Weber A, Hyman PE, Cucchiara S, Fleisher DR, Hyams JS, Milla PJ, et al. Childhood functional gastrointestinal...
Zhang 1991

Øster 1972

References to other published versions of this review
Huertas-Ceballos 2002

Huertas-Ceballos 2008a

Martin 2014a

CHARACTERISTICS OF STUDIES

Characteristics of included studies [ordered by study ID]

Asgarshirazi 2015

Methods
Design: 3-arm randomised placebo-controlled trial
Duration: 4 weeks

Participants
Country: Iran
Setting: paediatric gastroenterology outpatient clinic
Sample size: 120 children (40 synbiotic, 40 peppermint oil, 40 control)
Gender: peppermint oil group: 15 girls and 19 boys; control group: 17 girls and 8 boys
Mean age (SD): 7.06 (+/-2.38) years peppermint oil group; 7.42(+/-2.49) years control group
Dropouts/withdrawals: 32 withdrawals, number analysed = 88
Diagnosis: functional gastrointestinal disorders using Rome III (Rasquin 2006)

Interventions
Intervention: Colpermin (peppermint oil 187 mg capsule), 1 capsule if < 45 kg and 2 capsules if > 45 kg), 3 x day, 30 minutes before each meal
Control: folic acid tablet (1 mg), 1 tablet, once daily, 30 minutes before breakfast or lunch
3rd arm: synbiotic Lactol tablet (containing a probiotic and fructo-oligosaccharide) (see dietary review Newlove-Delgado in press for this comparison)

Outcomes
• Pain severity (using rating scale 0 to 10)
• Pain duration: minutes per day
• Pain frequency: episodes per week

Notes
Study dates: September 2012 to August 2014
Funding: not stated
Conflicts of interest: not reported

Risk of bias

Bias
Authors’ judgement
Support for judgement
| **Method** | **Design:** double-blinded, placebo-controlled, randomised trial  
**Duration:** 13 weeks  
**Participants** |  
**Country:** USA  
**Setting:** private paediatric gastroenterology outpatient clinic  
**Sample size:** 33 children (intervention 16, control 16)  
**Gender:** 9 boys, 24 girls  
**Mean age:** intervention 15.3 years; control 14.2 years. Overall range 12 to 18 years  
**Dropouts/withdrawals:** 0  
**Diagnosis:** meet the Rome II criteria for IBS [Rasquin-Weber 1999]  
**Interventions** |  
**Intervention:** 7-week course of oral amitriptyline (10 mg if 30 to 50 kg, 20 mg if 50 to 80 kg, 30 mg if > 80 kg), taken at night  
**Control:** placebo  
**Outcomes** |  
- Quality of life  
- Pain intensity and frequency  

**Asgarshirazi 2015 (Continued)**

<table>
<thead>
<tr>
<th><strong>Bias</strong></th>
<th><strong>Risk</strong></th>
<th><strong>Description</strong></th>
</tr>
</thead>
<tbody>
<tr>
<td>Random sequence generation (selection bias)</td>
<td>Unclear risk</td>
<td>The authors state that participants were randomised in blocks, but give no details of the method of random sequence generation.</td>
</tr>
<tr>
<td>Allocation concealment (selection bias)</td>
<td>Unclear risk</td>
<td>The authors do not acknowledge this concept in the paper. In personal correspondence, the authors reported that it was not important [Asgarshirazi 2016 [pers comm]]. Therefore, on balance, we judged this to be at unclear of risk of bias.</td>
</tr>
<tr>
<td>Blinding of participants and personnel (performance bias) All outcomes</td>
<td>Unclear risk</td>
<td>The study authors do not specifically state that participants were blinded, although they do say that the nurse that carried out the questionnaire was blinded.</td>
</tr>
<tr>
<td>Blinding of outcome assessment (detection bias) All outcomes</td>
<td>Unclear risk</td>
<td>As above, note also that the placebo regimen of folic acid tablets was once daily and the intervention regimen of peppermint oil capsules was 3 times daily. These different preparations and timings could have introduced bias, as the outcome was self reported and therefore influenced by the participants experience.</td>
</tr>
<tr>
<td>Incomplete outcome data (attrition bias) All outcomes</td>
<td>High risk</td>
<td>32 participants were excluded or withdrawn after randomisation. Attrition differed between the groups, 15% of the intervention group and 38% of the placebo group.</td>
</tr>
<tr>
<td>Selective reporting (reporting bias)</td>
<td>Low risk</td>
<td>All outcomes in the methods appear to be reported.</td>
</tr>
<tr>
<td>Other bias</td>
<td>High risk</td>
<td>No power calculation. The intervention and placebo groups differed at baseline: duration of pain in the intervention group was 67.05 mean minutes/day (SD +/- 36.97) and the placebo group was 53.4 mean minutes/day (SD +/- 16.81).</td>
</tr>
</tbody>
</table>
Bahar 2008 (Continued)

- IBS symptom checklist - validated in adults (Patrick 1997)
- Pain rating scale

**Timing of outcome assessment:** measured at 6, 10, and 13 weeks

**Notes**

- **Study dates:** 2002 to 2005
- **Funding:** James and Diane Brooks Medical Research Foundation and AstraZeneca
- **Conflicts of interest:** not mentioned

**Risk of bias**

<table>
<thead>
<tr>
<th>Bias</th>
<th>Authors' judgement</th>
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<td>The authors state that participants were randomised, but give no details of the method of randomisation.</td>
</tr>
<tr>
<td>Allocation concealment (selection bias)</td>
<td>High risk</td>
<td>The authors do not mention this concept in the paper.</td>
</tr>
<tr>
<td>Blinding of participants and personnel (performance bias) All outcomes</td>
<td>Unclear risk</td>
<td>This is mentioned in the title but no details are provided.</td>
</tr>
<tr>
<td>Blinding of outcome assessment (detection bias) All outcomes</td>
<td>Unclear risk</td>
<td>The outcome measurement is self assessed, therefore if the participant is blinded so too is the outcome assessment. No details of the blinding are provided.</td>
</tr>
<tr>
<td>Incomplete outcome data (attrition bias) All outcomes</td>
<td>Low risk</td>
<td>The authors state that no participants dropped out after starting treatment.</td>
</tr>
<tr>
<td>Selective reporting (reporting bias)</td>
<td>High risk</td>
<td>The authors state in their methods that improvement in overall quality of life score is the primary outcome but provide no details of a threshold regarded as improvement. At baseline, the quality of life score for the intervention group was 14% lower than that of the placebo group, 109.4 and 127.5 respectively. In the analysis, at 10 and 13 weeks' follow-up, the quality of life score was 128.0 and 126.2 respectively for the intervention group and 129.4 and 129.8 respectively for the placebo group. The authors conclude that the number of participants with a greater than 15% improvement is significantly higher in the intervention group; this could be wholly explained by the difference at baseline. This may also be a post-hoc analysis, as the authors do not mention this in the methods. The authors also selected some symptoms from the functional abdominal pain score and reported these as improved, e.g. periumbilical and right lower quadrant pain, but abdominal pain in other areas showed no improvement with the intervention.</td>
</tr>
<tr>
<td>Other bias</td>
<td>High risk</td>
<td>There is no power calculation. Baseline data of the participants are not provided. Compliance with the intervention is mentioned. There is no raw data to verify the statistics quoted. We wrote to the authors to request these data but received no response. The validity of the outcome assessment tool or clinical relevance of the reported outcomes are not mentioned. For example, is a 15% improvement in the quality of life scale at 13 weeks, but not at 10 weeks, clinically important? The number of participants is small and the follow-up is short.</td>
</tr>
</tbody>
</table>
Collins 2011

Methods

Design: double-blinded, placebo-controlled, randomised trial

Duration: 2 weeks

Participants

Country: USA

Setting: paediatric gastroenterology clinic at a children’s hospital and an outreach clinic

Sample size: 75 children (intervention 49, control 26)

Gender: intervention 15 boys, 34 girls; control 6 boys, 20 girls

Mean age: intervention 12.5 years (SD 3.01); control 13.16 years (SD 2.97)

Dropouts/withdrawals: 5: intervention 3, control 2

Diagnosis: Rome II criteria for IBS, functional dyspepsia, functional abdominal pain, or abdominal migraine (Rasquin-Weber 1999)

Interventions

Intervention: 550 mg rifaximin 3 times daily for 10 days

Control: placebo

Outcomes

- 10 symptoms rated on a visual analogue scale (0 to 10)

Timing of outcome assessment: 2 weeks

Notes

10 symptoms = bloating, excess gas, incomplete evacuation, abdominal pain, diarrhoea, constipation, urgency, passage of mucus, straining, faecal soiling

Study dates: 2010

Funding: the study was funded by Salix Pharmaceuticals, Saban Research Institute, and Children's Hospital of Los Angeles

Conflicts of interest: none reported

Risk of bias

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<td>The authors state that participants were randomised, but provide no details of the method of randomisation.</td>
</tr>
<tr>
<td>Allocation concealment (selection bias)</td>
<td>Low risk</td>
<td>The authors state that randomisation was performed by personnel not associated with the study and that the study personnel were blinded. This implies allocation concealment.</td>
</tr>
<tr>
<td>Blinding of participants and personnel (performance bias)</td>
<td>Low risk</td>
<td>The authors state that the participants and personnel were blinded and that a matching placebo was used.</td>
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<tr>
<td>Blinding of outcome assessment (detection bias)</td>
<td>Low risk</td>
<td>The outcomes were measured by self report in a questionnaire and the participants were blinded.</td>
</tr>
<tr>
<td>Incomplete outcome data (attrition bias)</td>
<td>Low risk</td>
<td>The authors state that outcome assessment is missing for 5 participants, 3 from the intervention group and 2 from the placebo group. This is a low num-</td>
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</tbody>
</table>
Collins 2011 (Continued)

<table>
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<td>The authors state that participants were randomised, but provide no details of the method of randomisation.</td>
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</tbody>
</table>

The number of participants is small and the follow-up is short.

Selective reporting (reporting bias) Unclear risk
The authors do not present the raw data, i.e. from the 10-point visual analogue scale of GI symptoms or the overall improvement score. They provide only a statement that none of the differences were significant. We wrote to the authors to provide the raw data to verify this but received no response. We are therefore unable to comment on the completeness of the outcome data or the degree of selective reporting.

Other bias High risk
The power calculation is not based on a primary outcome but a laboratory test - the detection of small intestinal bacteria overgrowth - and is therefore not valid. The outcome is measured by self report for 10 symptoms, but the validity of this method is not mentioned. The number of participants is small and the follow-up is short.

Heyland 2012

Methods
Design: double-blinded, placebo-controlled, randomised trial
Duration: 3 weeks

Participants
Country: Switzerland
Setting: patients referred to the Department of Paediatric Gastroenterology and Nutrition at the University Children’s Hospital, Zurich
Sample size: 37 children (intervention 20, control 17)
Gender: intervention 9 boys, 11 girls; control 7 boys, 10 girls
Mean age: intervention 10.6 years (SD not reported; range not reported); control 11.4 years (SD not reported; range not reported)
Dropouts/withdrawals: 3, not known from which group
Diagnosis: Apley’s criteria for RAP (Apley 1958)

Interventions
Intervention: 6 mg/kg/day trimethoprim and 30 mg/kg/day sulfamethoxazole in 2 divided doses for 7 days
Control: placebo twice daily for 7 days

Outcomes
• Pain Index measured on a visual analogue scale (0 to 10)
Timing of outcome assessment: 2 weeks

Notes
Study dates: 2004 to 2008
Funding: not stated in the paper
Conflicts of interest: none reported

Risk of bias

<table>
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<td>Unclear risk</td>
<td>The authors state that participants were randomised, but provide no details of the method of randomisation.</td>
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<tr>
<td><strong>Heyland 2012</strong> (Continued)</td>
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<tr>
<td><strong>Allocation concealment</strong> (selection bias)</td>
<td>Unclear risk</td>
<td>The authors state that randomisation occurred after obtaining written, informed consent, but no additional information is provided. It is therefore unclear if there was allocation concealment.</td>
</tr>
<tr>
<td><strong>Blinding of participants and personnel</strong> (performance bias)</td>
<td>Low risk</td>
<td>The authors state that the intervention was blinded and a placebo was used.</td>
</tr>
<tr>
<td>All outcomes</td>
<td></td>
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</tr>
<tr>
<td><strong>Blinding of outcome assessment</strong> (detection bias)</td>
<td>Low risk</td>
<td>Outcome was self reported and the participants were blinded.</td>
</tr>
<tr>
<td>All outcomes</td>
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</tr>
<tr>
<td><strong>Incomplete outcome data</strong> (attrition bias)</td>
<td>Low risk</td>
<td>The authors state that 3 participants withdrew from the study after inclusion, but this is unlikely to have significantly altered the results.</td>
</tr>
<tr>
<td>All outcomes</td>
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<tr>
<td><strong>Selective reporting</strong> (reporting bias)</td>
<td>Low risk</td>
<td>The primary outcome, Pain Index, was reported for all participants.</td>
</tr>
<tr>
<td>All outcomes</td>
<td></td>
<td></td>
</tr>
<tr>
<td><strong>Other bias</strong></td>
<td>Unclear risk</td>
<td>The power calculation is inadequate for this outcome measurement, as it is based on the breath test findings and not the primary outcome. The baseline data were presented and showed no imbalances. The number of participants is small and the follow-up is short. The funding is not stated. The authors report no conflicts of interest.</td>
</tr>
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</table>

**Karabulut 2013**

**Methods**

**Design:** randomised controlled trial comparing trimetubine with usual care  
**Duration:** 3 weeks

**Participants**

**Country:** Turkey  
**Setting:** patients attending the general paediatric clinic at the Istanbul University Hospital, Department of Paediatric Gastroenterology  
**Sample size:** 78 children (intervention 39, control 39)  
**Gender:** 31 boys, 47 girls. Distribution between groups not given.  
**Mean age:** 9.79 years (SD 3.45). Distribution between groups not given.  
**Dropouts/withdrawals:** 0  
**Diagnosis:** IBS according to the Rome III criteria (Rasquin 2006)

**Interventions**

**Intervention:** 3 mg/kg/day trimetubine maleate, in 3 divided doses for 3 weeks  
**Control:** usual care

**Outcomes**

- Treatment responders, measured by parental reporting of pain relief  
- Parents were asked: "Did your child have adequate relief?", responding "yes or no"  
**Timing of outcome assessment:** 3 weeks

**Notes**

**Study dates:** 2007 to 2008
Karabulut 2013 (Continued)

**Funding:** none

**Conflicts of interest:** none reported

### Risk of bias

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<td>The authors state that participants were randomised, but give no details of the method of randomisation.</td>
</tr>
<tr>
<td>Allocation concealment (selection bias)</td>
<td>High risk</td>
<td>This concept is not mentioned in the paper.</td>
</tr>
<tr>
<td>Blinding of participants and personnel (performance bias)</td>
<td>High risk</td>
<td>The participants and study personnel could not have been blinded as there was no placebo for the control group, who received standard care instead.</td>
</tr>
<tr>
<td>Blinding of outcome assessment (detection bias)</td>
<td>High risk</td>
<td>The parents were asked to report the outcome, and they would have known whether the child was taking trimetinone or not.</td>
</tr>
<tr>
<td>Incomplete outcome data (attrition bias)</td>
<td>Unclear risk</td>
<td>The authors do not mention this; the raw data is not reported, only the percentage of children who improved. We therefore cannot be sure about the loss to follow-up or how the missing data were handled.</td>
</tr>
<tr>
<td>Selective reporting (reporting bias)</td>
<td>Low risk</td>
<td>The authors state that parental assessment of pain relief was their only outcome.</td>
</tr>
<tr>
<td>Other bias</td>
<td>High risk</td>
<td>No power calculation was performed and no baseline data are given. Compliance with the intervention is not mentioned. The outcome measurement is not validated. The number of participants is small and the follow-up is short. We wrote to the authors to request details regarding the above issues, but received no response.</td>
</tr>
</tbody>
</table>

Karunanayake 2015

**Methods**

**Design:** randomised, placebo-controlled trial

**Duration:** insufficient detail. The intervention was given for 8 weeks but when the outcomes were measured is not stated.

**Participants**

**Country:** Sri Lanka

**Setting:** outpatient clinic of the university paediatric unit

**Sample size:** 100 children

**Gender:** no details about the children at recruitment. In the 89 children that completed the study, 33/47 in the intervention group were girls, and 22/42 in the placebo group were girls.

**Mean age:** no details. The authors state that the study was designed to recruit children aged 5 to 12 years.

**Dropouts/withdrawals:** 11
Karunanayake 2015 (Continued) **Diagnosis:** abdominal pain-predominant functional gastrointestinal diseases fulfilling Rome III criteria (Rasquin 2006)

**Interventions**

- **Intervention:** domperidone 10 mg, 3 times per day, before meals for 8 weeks
- **Control:** placebo

**Outcomes**

- **Primary outcomes:**
  - cure (abdominal pain less than 25 mm on the visual analogue scale and no impact on daily activities);
  - improvement (pain relief and sense of improvement recorded on Global Assessment Scale).

- **Secondary outcomes were significant improvement in:**
  - symptoms, severity was recorded on a validated 100-millimetre visual analogue scale;
  - gastric motility, assessed using a validated ultrasound method;
  - quality of life, validated PedsQL Generic Core Scale, version 4.0;
  - family impact, used PedsQL Family Impact Module.

**Timing of outcome assessment:** no details

**Notes**

- **Study dates:** no details
- **Funding:** no details
- **Conflicts of interest:** no details

This study is published as a conference abstract. We wrote to the authors and received no further information (Martin 2016a [pers comm]; Martin 2016d [pers comm]).

### Risk of bias

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<tbody>
<tr>
<td>Random sequence generation (selection bias)</td>
<td>Low risk</td>
<td>The authors state that computer-generated random numbers were used.</td>
</tr>
<tr>
<td>Allocation concealment (selection bias)</td>
<td>Unclear risk</td>
<td>The authors do not mention this.</td>
</tr>
<tr>
<td>Blinding of participants and personnel (performance bias) All outcomes</td>
<td>Low risk</td>
<td>The authors state that participants were blinded.</td>
</tr>
<tr>
<td>Blinding of outcome assessment (detection bias) All outcomes</td>
<td>Low risk</td>
<td>The authors state that participants and investigators who assessed outcome measures were blinded.</td>
</tr>
<tr>
<td>Incomplete outcome data (attrition bias) All outcomes</td>
<td>Unclear risk</td>
<td>There is insufficient detail in the published information to assess this.</td>
</tr>
<tr>
<td>Selective reporting (reporting bias)</td>
<td>Low risk</td>
<td>All stated outcomes appear to be reported, although with very limited data.</td>
</tr>
<tr>
<td>Other bias</td>
<td>Unclear risk</td>
<td>There is insufficient detail in the published information to assess this.</td>
</tr>
</tbody>
</table>
### Khoshoo 2006

**Methods**

**Design:** randomised controlled trial comparing tegaserod plus laxative with laxative alone

**Duration:** 4 weeks

**Participants**

**Country:** USA

**Setting:** medical centre in North America (no further detail given by authors)

**Sample size:** 48 children (intervention 21, control 27)

**Gender:** intervention 8 boys, 13 girls; control 11 boys, 16 girls

**Mean age:** intervention 15.1 years (SD not reported); control 15.37 years (SD not reported)

**Dropouts/withdrawals:** 0

**Diagnosis:** constipation-predominant IBS according to Rome II criteria (Rasquin-Weber 1999)

**Interventions**

**Intervention:** 6 mg tegaserod once daily, oral, on an empty stomach

**Control:** usual care

**Outcomes**

- Pain measured by visual analogue scale (0 to 10)
- Compared mean pain scores and number of children with a good pain reduction, defined by reduction in score by at least 3 points

**Timing of outcome assessment:** 4 weeks

**Notes**

**Study dates:** 2006

**Funding:** authors state no external financial support was received

**Conflicts of interest:** none reported

### Risk of bias

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<td>Random sequence generation (selection bias)</td>
<td>Low risk</td>
<td>Authors state that they used a simple random numbers table to allocate the treatment groups.</td>
</tr>
<tr>
<td>Allocation concealment (selection bias)</td>
<td>Unclear risk</td>
<td>This concept is not mentioned in the paper.</td>
</tr>
<tr>
<td>Blinding of participants and personnel (performance bias)</td>
<td>High risk</td>
<td>The control group was not given a placebo, therefore if informed consent was obtained before inclusion, participants would have been aware of the implications of their group allocation. The authors state in the discussion that the participants were unaware of the different groups. This is concerning. We asked the authors about this but have not received a reply.</td>
</tr>
<tr>
<td>Blinding of outcome assessment (detection bias)</td>
<td>High risk</td>
<td>Self reported outcome and therefore not blinded, as explained in performance bias section, see above</td>
</tr>
<tr>
<td>Incomplete outcome data (attrition bias)</td>
<td>Low risk</td>
<td>There appears to be no loss to follow-up.</td>
</tr>
</tbody>
</table>
Khoshoo 2006 (Continued)

Selective reporting (reporting bias) Unclear risk The authors report all the outcomes they stated in the methods, but it is unclear from the methods if the definition of a good pain reduction was a prespecified outcome; it may be post-hoc analysis.

Other bias Unclear risk The 2 groups are similar at baseline. A power calculation is performed. The eligibility criteria are stated. It is unclear if the pain-rating tool is validated as the authors state "standard pain rating scale" but give no further details.

Kline 2001

Methods Design: double-blinded, placebo-controlled, randomised trial
Duration: 2 weeks

Participants Country: USA
Setting: 3 specialist paediatric gastroenterology departments
Sample size: 42 children (intervention 21, control 21)
Gender: 17 boys, 25 girls. Distribution between groups not given.
Mean age: 12 years (range 8 to 17 years). Distribution between groups not given.
Dropouts/withdrawals: 8 (4 from intervention, 4 from control)
Diagnosis: RAP, defined by Apley (Apley 1958)

Interventions Intervention: peppermint oil capsule taken 3 times daily
Control: placebo (arachis oil) capsule

Outcomes • Severity of pain, scale: good to bad (1 to 5)
• Change in pain experience, scale: better to worse (1 to 5)
• 15-item gastrointestinal symptom rating scale
Timing of outcome assessment: 2 weeks

Notes Study dates: 1999
Funding: Tillotts Pharma AG Switzerland
Conflicts of interest: none declared

Risk of bias

<table>
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<td>The authors state this was a randomised controlled trial, but give no details of method of randomisation.</td>
</tr>
<tr>
<td>Allocation concealment (selection bias)</td>
<td>Unclear risk</td>
<td>The authors did not mention this concept.</td>
</tr>
<tr>
<td>Blinding of participants and personnel (performance bias) All outcomes</td>
<td>Unclear risk</td>
<td>The authors state that the trial was double-blinded and describe that the same company produced identical intervention and placebo tablets. However, it is unclear who, if any, of the study personnel were blinded.</td>
</tr>
</tbody>
</table>
Kline 2001 (Continued)

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<td>Blinding of outcome assessment (detection bias)</td>
<td>Unclear risk</td>
<td>There is insufficient information in the paper to judge this because the outcome was based on clinician judgement, parent report, and child report, and the blinding of all these outcome assessors was not explained.</td>
</tr>
<tr>
<td>Incomplete outcome data (attrition bias)</td>
<td>Low risk</td>
<td>The authors state that 8 participants withdrew from the study. Reasons for withdrawal (4 from each group) are clearly stated in the paper and appear not to be related to treatment effects (2 because of long travel distances to the clinic; 2 because they were prescribed antibiotics for other reasons; 4 because they were unable to swallow pills).</td>
</tr>
<tr>
<td>Selective reporting (reporting bias)</td>
<td>High risk</td>
<td>The authors do not report all 5 outcomes mentioned in the methods section.</td>
</tr>
<tr>
<td>Other bias</td>
<td>Unclear risk</td>
<td>There was no power calculation. The eligibility criteria were clearly stated. The authors state that the groups were similar at baseline, but the details are not reported. Compliance with the intervention is unclear. The follow-up is short.</td>
</tr>
</tbody>
</table>

Narang 2015

Methods

- **Design:** double-blinded, placebo-controlled, randomised trial
- **Duration:** 4 weeks

Participants

- **Country:** India
- **Setting:** tertiary paediatric gastroenterology clinic
- **Sample size:** 132 children (intervention 66, control 66)
- **Gender:** intervention group: 33 boys, 33 girls; control group: 39 boys, 27 girls
- **Mean age:** intervention group: 7.1 (+/- 2.1); control group: 7.4 (+/- 2.6)
- **Dropouts/withdrawals:** 8 (4 from intervention, 4 from control)
- **Diagnosis:** RAP, defined by Apley (Apley 1958)

Interventions

- **Intervention:** drotaverine 20 mg (10 mL) liquid 3 x daily for children 4 to 6 years, 40 mg tablet 3 x daily for children >6 years
- **Control:** placebo

Outcomes

- Number of episodes of pain in 4 weeks
- Number of pain-free days in 4 weeks
- Pain severity measured by visual analogue scale and FACES Pain Scale
- Number of school days missed in 4 weeks
- Parental reporting of their perception of child’s mood, activity, alertness, oral intake, and comfort, quantified by Likert scale

**Timing of outcome assessment:** 4 weeks

Notes

- **Study dates:** September 2012 to September 2013
- **Funding:** Walter Bushnell Pvt. Ltd. The company supplied a grant and medication for the study. They had no role in trial design, data collection, or manuscript preparation.
Narang 2015 (Continued)

Conflicts of interest: none declared

Risk of bias

<table>
<thead>
<tr>
<th>Bias</th>
<th>Authors' judgement</th>
<th>Support for judgement</th>
</tr>
</thead>
<tbody>
<tr>
<td>Random sequence generation (selection bias)</td>
<td>Low risk</td>
<td>Authors state that they used computer-generated block randomisation with variable block sizes. Stratified for age (4 to 6 and &gt; 6 years).</td>
</tr>
<tr>
<td>Allocation concealment (selection bias)</td>
<td>Low risk</td>
<td>Clearly explained by authors</td>
</tr>
<tr>
<td>Blinding of participants and personnel (performance bias) All outcomes</td>
<td>Low risk</td>
<td>Clearly explained by authors</td>
</tr>
<tr>
<td>Blinding of outcome assessment (detection bias) All outcomes</td>
<td>Low risk</td>
<td>The outcomes were self and parental reported, therefore participant blinding meant the outcome assessment was also blinded.</td>
</tr>
<tr>
<td>Incomplete outcome data (attrition bias) All outcomes</td>
<td>Low risk</td>
<td>Low loss to follow-up: 6 from placebo group and 2 from intervention group</td>
</tr>
<tr>
<td>Selective reporting (reporting bias)</td>
<td>High risk</td>
<td>The outcomes explained in the methods were not all reported, e.g. results for pain severity are not given. In addition, the outcome 'number of school days missed in 4 weeks' reported in the results was not mentioned in the methods.</td>
</tr>
<tr>
<td>Other bias</td>
<td>High risk</td>
<td>We were concerned that the method of giving further drug doses if pain is reported could alter the trial outcome reporting, as if the drug works the child may ask for it more frequently. Therefore, administration of the intervention and outcome assessment are not independent.</td>
</tr>
</tbody>
</table>

Pourmoghaddas 2014

Methods

Design: double-blinded, placebo-controlled, randomised trial of mebeverine versus placebo

Duration: 12 weeks

Participants

Country: Iran

Setting: paediatric gastroenterology outpatient clinic

Sample size: 115 children (intervention 59, control 56)

Gender: 39 boys, 48 girls (children who completed follow-up)

Mean age: 8.5 years (SD 2.1 years) for all children who completed follow-up. Distribution between treatment groups not given.

Dropouts/withdrawals: intervention 19, control 17

Diagnosis: functional abdominal pain, defined by Rome III criteria (Rasquin 2006)

Interventions

Intervention: mebeverine tablets, 135 mg twice daily, for a duration of 4 weeks

Control: placebo
Outcomes

Primary outcome: 2-point reduction in FACES Pain Scale (scale 1 to 6) or report of "no pain"

Secondary outcome: physician-rated global severity and improvement using the Clinical Global Impression Severity and Improvement Scales (scale 1 to 7)

Timing of outcome assessment: 12 weeks

Notes

Study dates: 2013

Funding: The Isfahan University of Medical Sciences

Conflicts of interest: none declared

Risk of bias

<table>
<thead>
<tr>
<th>Bias</th>
<th>Authors' judgement</th>
<th>Support for judgement</th>
</tr>
</thead>
<tbody>
<tr>
<td>Random sequence generation (selection bias)</td>
<td>Low risk</td>
<td>Randomisation by computer-generated random numbers in 4 blocks</td>
</tr>
<tr>
<td>Allocation concealment (selection bias)</td>
<td>Low risk</td>
<td>Authors state allocation was concealed.</td>
</tr>
<tr>
<td>Blinding of participants and personnel (performance bias)</td>
<td>Low risk</td>
<td>Authors state that the doctors, study participants, and study personnel assessing the outcomes were blinded. The drug codes for labelling the bottles were known only by the pharmacist.</td>
</tr>
<tr>
<td>Blinding of outcome assessment (detection bias)</td>
<td>Low risk</td>
<td>The primary outcome was self reported and therefore blinded.</td>
</tr>
<tr>
<td>Incomplete outcome data (attrition bias)</td>
<td>Low risk</td>
<td>The loss to follow-up was equal between the 2 groups and was 28/115 (24%) of randomised participants, shown in flow diagram.</td>
</tr>
<tr>
<td>Selective reporting (reporting bias)</td>
<td>Low risk</td>
<td>All outcomes stated in the methods are reported.</td>
</tr>
<tr>
<td>Other bias</td>
<td>Low risk</td>
<td>2 groups similar at baseline. Power calculation performed. No conflict of interest. Funded by an academic institution</td>
</tr>
</tbody>
</table>

Methods

Design: double-blinded, placebo-controlled, randomised trial of citalopram versus placebo

Duration: 12 weeks

Participants

Country: Iran

Setting: paediatric gastroenterology outpatient clinic

Sample size: 115 children (intervention 59, control 56)

Gender: intervention group: 11 males, 32 females; control group: 19 males, 24 females (children who completed follow-up)
**Mean age:** intervention 10.4 years (SD 1.9 years); control 8.5 years (SD 2.2 years) for children who completed follow-up

**Dropouts/withdrawals:** 29 children (intervention 16, control 13)

**Diagnosis:** functional abdominal pain, defined by Rome III criteria (Rasquin 2006)

### Interventions

**Intervention:** citalopram tablets 10 mg/day for the 1st week and then 20 mg/day for a total duration of 4 weeks

**Control:** placebo

### Outcomes

**Primary outcome:** 2-point reduction in FACES Pain Scale (scale 1 to 6) or report of "no pain"

**Secondary outcome:** self rated depression, anxiety, and somatization scores

**Timing of outcome assessment:** 12 weeks

### Notes

**Study dates:** 2013

**Funding:** The Isfahan University of Medical Sciences

**Conflicts of interest:** none declared

### Risk of bias

<table>
<thead>
<tr>
<th>Bias</th>
<th>Authors' judgement</th>
<th>Support for judgement</th>
</tr>
</thead>
<tbody>
<tr>
<td>Random sequence generation (selection bias)</td>
<td>Low risk</td>
<td>Randomisation by computer-generated random numbers in 4 blocks</td>
</tr>
<tr>
<td>Allocation concealment (selection bias)</td>
<td>Low risk</td>
<td>Authors state that allocation was concealed.</td>
</tr>
<tr>
<td>Blinding of participants and personnel (performance bias)</td>
<td>Low risk</td>
<td>Authors state that the doctors, study participants, and study personnel assessing the outcomes were blinded. The drug codes for labelling the bottles were known only by the pharmacist.</td>
</tr>
<tr>
<td>Blinding of outcome assessment (detection bias)</td>
<td>Low risk</td>
<td>The primary outcome was self reported and therefore blinded.</td>
</tr>
<tr>
<td>Incomplete outcome data (attrition bias)</td>
<td>Low risk</td>
<td>The loss to follow-up was equal between the 2 groups and was 29/115 (25%) of randomised participants, shown in flow diagram.</td>
</tr>
<tr>
<td>Selective reporting (reporting bias)</td>
<td>Low risk</td>
<td>All outcomes stated in the methods are reported.</td>
</tr>
<tr>
<td>Other bias</td>
<td>Low risk</td>
<td>2 groups similar at baseline, except for slight difference in age. Power calculation performed. No conflict of interest. Funded by an academic institution</td>
</tr>
</tbody>
</table>

**Sadeghian 2008**

### Methods

**Design:** double-blinded, placebo-controlled, randomised trial

**Duration:** 2 weeks
Sadeghian 2008 (Continued)

Participants

- **Country**: Iran
- **Setting**: paediatric gastroenterology clinic in Tehran
- **Sample size**: 36 children (included after dropouts 29 children: intervention 15, control 14)
- **Gender**: intervention: 7 boys, 8 girls; control: 5 boys, 9 girls
- **Mean age**: intervention 7.2 years (range 4.5 to 11 years); control 7.7 years (range 5 to 12 years)
- **Dropouts/withdrawals**: 7 (unclear from which group)
- **Diagnosis**: functional abdominal pain, defined by Rome II criteria (Rasquin-Weber 1999)

Interventions

- **Intervention**: 0.25 to 0.5 mg/kg/day cyproheptadine in 2 divided doses for 2 weeks
- **Control**: placebo

Outcomes

- **Pain**: * self reported change in abdominal pain frequency (scale 1 to 6); * self reported change in abdominal pain intensity (scale 1 to 6); * self reported impression of improvement (scale 1 to 4); * parental reporting of improvement (satisfactory or not).
- **Timing of outcome assessment**: 2 weeks

Notes

- **Study dates**: 2006 to 2007
- **Funding**: none declared
- **Conflicts of interest**: none declared

Risk of bias

<table>
<thead>
<tr>
<th>Bias</th>
<th>Authors’ judgement</th>
<th>Support for judgement</th>
</tr>
</thead>
<tbody>
<tr>
<td>Random sequence generation (selection bias)</td>
<td>Low risk</td>
<td>The authors state that the random allocation was performed using a randomised, quadruple order of A and B.</td>
</tr>
<tr>
<td>Allocation concealment (selection bias)</td>
<td>Low risk</td>
<td>The authors explain that the study personnel (research nurse and study physician) were blinded. This implies allocation concealment, as they did not know the group to which participants were allocated.</td>
</tr>
<tr>
<td>Blinding of participants and personnel (performance bias) All outcomes</td>
<td>Low risk</td>
<td>The authors state that the placebo was similar.</td>
</tr>
<tr>
<td>Blinding of outcome assessment (detection bias) All outcomes</td>
<td>Low risk</td>
<td>The participants were blinded and the outcome was self reported.</td>
</tr>
<tr>
<td>Incomplete outcome data (attrition bias) All outcomes</td>
<td>Low risk</td>
<td>36 participants were randomised. The authors state that 7 participants were excluded from the final analysis (5 for poor compliance and 2 for incomplete outcome data). We assume that they came equally from each group, as the included numbers of participants were 15 and 14 for the intervention and control groups, respectively (total 29).</td>
</tr>
<tr>
<td>Selective reporting (reporting bias)</td>
<td>Low risk</td>
<td>The primary outcome was reported.</td>
</tr>
</tbody>
</table>
Sadeghian 2008 (Continued)

Other bias

High risk

No power calculation was performed. The outcome was not measured by a validated tool. The sample size was small and the outcome was measured over a short time duration. There was no intention-to-treat analysis as the non-compliant participants were excluded.

Saps 2009

Methods

Design: double-blinded, placebo-controlled, randomised trial

Duration: 4 weeks

Participants

Country: USA

Setting: 6 paediatric tertiary care hospitals, gastroenterology clinics

Sample size: 90 children (intervention 46, control 44)

Gender: 18 boys, 72 girls. Distribution between groups not given.

Mean age: 12.7 years (range 8 to 17 years)

Dropouts/withdrawals: 7 (3 from intervention group, 4 from control group)

Diagnosis: functional abdominal pain, functional dyspepsia, or IBS, defined by Rome II criteria (Rasquin-Weber 1999)

Interventions

Intervention: 10 mg/day amitriptyline if weight < 35 kg, 20 mg/day amitriptyline if weight > 35 kg, for 4 weeks

Control: placebo

Outcomes

• Self reporting of symptoms; children were asked:
  * "How do you feel your problem is?" Responses were ordinal: better, same, or worse.
  * "How well did the medication relieve your pain?" Responses were ordinal: excellent, good, fair, poor, or failed.

• These outcomes were then dichotomised into “improved” and “not improved”.

Timing of outcome assessment: 4 weeks

Notes

Study dates: 2003 to 2006

Funding: grant from American College of Gastroenterology and National Center for Research Resources, National Institutes of Health

Conflicts of interest: none declared

Risk of bias

<table>
<thead>
<tr>
<th>Bias</th>
<th>Authors' judgement</th>
<th>Support for judgement</th>
</tr>
</thead>
<tbody>
<tr>
<td>Random sequence generation (selection bias)</td>
<td>Unclear risk</td>
<td>The authors state that participants were randomised, but give no details of the method of randomisation. The 2 groups look similar at baseline, which suggests the method of randomisation was adequate.</td>
</tr>
<tr>
<td>Allocation concealment (selection bias)</td>
<td>Unclear risk</td>
<td>This concept is not mentioned in the paper.</td>
</tr>
</tbody>
</table>
Methods

Design: double-blinded, placebo-controlled, randomised, cross-over trial

Duration: 6 weeks

Participants

Country: USA

Setting: tertiary paediatric hospital clinic at Mount Sinai Medical Center

Sample size: 25 children (cross-over trial so all participants received both interventions)

Gender: 13 boys, 12 girls

Mean age: 10.5 years (range 5.5 to 16.77 years)

Dropouts/withdrawals: 1 participant, excluded after enrolment (as Giardia detected in stool)

Diagnosis: RAP, defined by Apley’s criteria (Apley 1958)

Interventions

Intervention: 1 mg/kg/day famotidine in 2 divided doses for 3 weeks

Control: placebo

Outcomes

- Abdominal pain score (= pain frequency + pain severity + peptic index), self reported
- Global assessment = self reporting of response to intervention, “Have you felt better, not better or worse?”

Timing of outcome assessment: 3 weeks

Notes

Study dates: 1998

Funding: none declared

Conflicts of interest: none declared

Risk of bias

<table>
<thead>
<tr>
<th>Bias</th>
<th>Authors’ judgement</th>
<th>Support for judgement</th>
</tr>
</thead>
</table>

Cochrane Database of Systematic Reviews

Pharmacological interventions for recurrent abdominal pain in childhood (Review)

Copyright © 2017 The Cochrane Collaboration. Published by John Wiley & Sons, Ltd.
Symon 1995

Methods

Design: double-blinded, placebo-controlled, randomised trial
Duration: 4 months

Participants

Country: UK
Setting: hospital general paediatric clinic
Sample size: 14 children
Gender: not stated by authors
Mean age: not reported (range 5 to 13 years)
Dropouts/withdrawals: 2
Diagnosis: abdominal migraine = RAP (Apley 1958), with episodes associated with facial pallor, first-degree relative with headaches

Interventions

Intervention: 0.5 mg pizotifen syrup (0.25 mg/5 mL) in 2 divided doses per day. After 1 month the dose was increased to 0.75 mg/day if there had been no improvement.
Control: placebo
### Symon 1995 (Continued)

**Outcomes**
- Days of abdominal pain
- Index of severity (mild = 1, moderate = 2, severe = 3 attack; totaled for whole period)
- Index of misery (= pain severity multiplied by duration)

**Timing of outcome assessment:** 4 months

**Notes**
- **Study dates:** 1995
- **Funding:** none declared
- **Conflicts of interest:** none declared

### Risk of bias

<table>
<thead>
<tr>
<th>Bias</th>
<th>Authors’ judgement</th>
<th>Support for judgement</th>
</tr>
</thead>
<tbody>
<tr>
<td>Random sequence generation (selection bias)</td>
<td>Low risk</td>
<td>From correspondence with the authors, it was established that the order for treatment allocation was both blinded and randomised. This was done in the hospital pharmacy, therefore the study personnel were also blinded.</td>
</tr>
<tr>
<td>Allocation concealment (selection bias)</td>
<td>Low risk</td>
<td>From correspondence with the authors, it was established that the randomisation was performed by the hospital pharmacy, and therefore the study personnel were also blinded. This implies there was allocation concealment.</td>
</tr>
<tr>
<td>Blinding of participants and personnel (performance bias) All outcomes</td>
<td>Low risk</td>
<td>As explained above</td>
</tr>
<tr>
<td>Blinding of outcome assessment (detection bias) All outcomes</td>
<td>Low risk</td>
<td>The outcomes were reported by parent and child, who were blinded.</td>
</tr>
<tr>
<td>Incomplete outcome data (attrition bias) All outcomes</td>
<td>Low risk</td>
<td>2 participants (14%) were excluded, as they had no outcome data.</td>
</tr>
<tr>
<td>Selective reporting (reporting bias)</td>
<td>Low risk</td>
<td>All outcomes mentioned in the methods were reported.</td>
</tr>
<tr>
<td>Other bias</td>
<td>High risk</td>
<td>No baseline data, no power calculation, no washout period for a cross-over trial. The intervention was changed according to response; the authors stated that the medication was doubled if there was no response. The trial stopped early due to lack of drug supply.</td>
</tr>
</tbody>
</table>

### Zybach 2016

**Methods**
- **Design:** cross-over, placebo-controlled, randomised trial
- **Duration:** 5 weeks

**Participants**
- **Country:** USA
- **Setting:** hospital general paediatric clinic
- **Sample size:** 14 children
Zybach 2016

**Gender:** intervention group (girls 58%; boys 5, girls 7), placebo group (girls 58%; boys 5, girls 7)

**Mean age:** intervention group: 13.8 (+/- 1.6 years), placebo group: 13.8 (+/- 1.6 years)

**Dropouts/withdrawals:** 2

**Diagnosis:** functional abdominal pain, according to Rome III criteria

### Interventions

**Intervention:** melatonin 5 mg (20 mL) at night, once daily for 14 days

**Control:** placebo

### Outcomes

- Abdominal pain. Global clinical score: participants were interviewed to measure change in abdominal pain, graded 1 to 5. Grade 3 or above = responders, grade 1 to 2 = non-responders
- Sleep. Actigraphy (validated tool for measuring sleep), included time in bed, sleep latency, sleep duration from movements measured by accelerometer

**Timing of outcome assessment:** 2 weeks

### Notes

**Study dates:** not stated by authors, published in 2016

**Funding:** none declared

**Conflicts of interest:** none declared

### Risk of bias

<table>
<thead>
<tr>
<th>Bias</th>
<th>Authors' judgement</th>
<th>Support for judgement</th>
</tr>
</thead>
<tbody>
<tr>
<td>Random sequence generation (selection bias)</td>
<td>Unclear risk</td>
<td>The authors state that participants were randomised, but give no details of the method of randomisation.</td>
</tr>
<tr>
<td>Allocation concealment (selection bias)</td>
<td>Low risk</td>
<td>Authors clearly state that parents were consented for the trial, and participant’s assent was obtained prior to study procedures, which implies that allocation concealment was achieved.</td>
</tr>
<tr>
<td>Blinding of participants and personnel (performance bias) All outcomes</td>
<td>Low risk</td>
<td>Blinding explained, complete for participant and study team, the drug and placebo were labelled in pharmacy.</td>
</tr>
<tr>
<td>Blinding of outcome assessment (detection bias) All outcomes</td>
<td>Low risk</td>
<td>The outcome assessment was self reported pain by the child, who was interviewed by the study team. Both the child and study team were blinded.</td>
</tr>
<tr>
<td>Incomplete outcome data (attrition bias) All outcomes</td>
<td>Low risk</td>
<td>14 enrolled, 12 completed the study; 14% dropout rate</td>
</tr>
<tr>
<td>Selective reporting (reporting bias)</td>
<td>Low risk</td>
<td>No evidence from the paper of outcomes that were not reported.</td>
</tr>
<tr>
<td>Other bias</td>
<td>High risk</td>
<td>Small study with no power calculation. This is a cross-over study with no washout period.</td>
</tr>
</tbody>
</table>

GI: gastrointestinal
IBS: irritable bowel syndrome
PedsQL: Pediatric Quality of Life Inventory
RAP: recurrent abdominal pain
SD: standard deviation
### Characteristics of excluded studies [ordered by study ID]

<table>
<thead>
<tr>
<th>Study</th>
<th>Reason for exclusion</th>
</tr>
</thead>
<tbody>
<tr>
<td>Christensen 1995</td>
<td>Ineligible study design</td>
</tr>
<tr>
<td>Cucchiara 1992</td>
<td>Ineligible study design</td>
</tr>
<tr>
<td>Dehghani 2011</td>
<td>Ineligible study design</td>
</tr>
<tr>
<td>Di Nardo 2011</td>
<td>Ineligible population</td>
</tr>
<tr>
<td>Everitt 2010</td>
<td>Ineligible population</td>
</tr>
<tr>
<td>Grillage 1990</td>
<td>Ineligible comparator</td>
</tr>
<tr>
<td>Kaminski 2009</td>
<td>Ineligible study design: literature review</td>
</tr>
<tr>
<td>Lloyd-Still 1990</td>
<td>Ineligible population</td>
</tr>
<tr>
<td>Van Outryve 2005</td>
<td>Ineligible population</td>
</tr>
<tr>
<td>Xiao 2013</td>
<td>Ineligible comparator</td>
</tr>
<tr>
<td>Yadav 1989</td>
<td>Ineligible population</td>
</tr>
</tbody>
</table>

### ADDITIONAL TABLES

**Table 1. Assessment of risk of bias in included studies** (Continued)

<table>
<thead>
<tr>
<th>Domain</th>
<th>'Risk of bias' judgement</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Low</td>
</tr>
<tr>
<td><strong>Selection bias</strong></td>
<td></td>
</tr>
<tr>
<td>Random sequence generation</td>
<td>If the study details any of the following methods: (1) simple randomisation (such as coin-tossing, throwing dice, or dealing previously shuffled cards, a list of random numbers, or computer-generated random numbers); or (2) restricted randomisation such as blocked, ideally with varying block sizes or stratified groups, provided that within-group randomisation is not affected</td>
</tr>
<tr>
<td>Allocation concealment</td>
<td>If the study details concealed allocation sequence in sufficient detail to determine that allocations could not have been foreseen in advance of, or during, enrolment</td>
</tr>
</tbody>
</table>

**Perfor-**

...
Table 1. Assessment of risk of bias in included studies (Continued)

<table>
<thead>
<tr>
<th>Blinding of participants and personnel</th>
<th>Detection bias</th>
<th>Attrition bias</th>
</tr>
</thead>
<tbody>
<tr>
<td>If the study details a method of blinding participants and personnel. Detail would need to be sufficient to show that participants and personnel were unable to identify the therapeutic intervention from the control intervention.</td>
<td>If the study details a blinded outcome assessment. This may only be possible for outcomes that are externally assessed.</td>
<td>If the study reports attrition and exclusions, including the numbers in each intervention group (compared with total randomised participants), reasons for attrition or exclusions, and any re-inclusions; the impact of missing data is not believed to have altered the conclusions; and reasons for the missing data are acceptable.</td>
</tr>
<tr>
<td>If the methods detail that the participants or study personnel were not blinded to the study medication or placebo.</td>
<td>If the outcome assessment is not blinded. We expect this may be unavoidable for self rated outcomes of unblinded interventions.</td>
<td>We may judge the risk of attrition bias to be high due to the amount, nature, or handling (such as per-protocol analysis) of incomplete outcome data.</td>
</tr>
<tr>
<td>If there is insufficient detail to judge the risk of bias.</td>
<td>If there is insufficient detail to judge the risk of bias.</td>
<td>If there is insufficient detail to judge the risk of bias.</td>
</tr>
</tbody>
</table>

<table>
<thead>
<tr>
<th>Reporting bias</th>
</tr>
</thead>
<tbody>
<tr>
<td>Selective reporting</td>
</tr>
<tr>
<td>If there is insufficient detail to judge the risk of bias.</td>
</tr>
</tbody>
</table>

<table>
<thead>
<tr>
<th>Other sources of bias</th>
</tr>
</thead>
<tbody>
<tr>
<td>Other bias</td>
</tr>
<tr>
<td>If there is insufficient detail to judge the risk of bias.</td>
</tr>
</tbody>
</table>

APPENDICES

Appendix 1. Search strategies

Cochrane Central Register of Controlled Studies (CENTRAL; 2016, Issue 5) in the Cochrane Library

Search dates: 19 April 2013 (990 records); 11 April 2014 (1271 records); 26 March 2015 (49 records); 10 June 2016 (81 records).

#1 Pain*:ti,ab
#2 Acne*:ti,ab
Pharmacological interventions for recurrent abdominal pain in childhood (Review)

Copyright © 2017 The Cochrane Collaboration. Published by John Wiley & Sons, Ltd.
#65 Adolec*nt*,ti,ab
#66 Boy*,ti,ab
#67 Gir*l*,ti,ab
#68 teen*,ti,ab
#69 Schoolchild*,ti,ab
#70 Young adult*,ti,ab
#71 Youth*,ti,ab
#72 Pediatric*,ti,ab
#73 Paediatric*,ti,ab
#74 Student*,ti,ab
#75 Pupil*,ti,ab
#76 Juvenile*,ti,ab
#77 Young person*,ti,ab
#78 MeSH descriptor: [Child] explode all trees
#79 MeSH descriptor: [Adolescent] explode all trees
#80 MeSH descriptor: [Young Adult] explode all trees
#81 MeSH descriptor: [Students] explode all trees
#82 #64 or #65 or #66 or #67 or #68 or #69 or #70 or #71 or #72 or #73 or #74 or #75 or #76 or #77 or #78 or #80 or #81
#83 #63 and #82

Ovid MEDLINE In-Process & Other Non-Indexed Citations and Ovid Medline (1946 to present)

Search dates: 11 April 2013 (6238 records); 11 April 2014 (5957 records); 25 March 2015 (223 records); 9 June 2016 (300 records).

1 stomach*.tw.
2 abdomen*.tw.
3 intestin*.tw.
4 viscera*.tw.
5 tummy.tw.
6 bowel*.tw.
7 belly.tw.
8 gastrointestinal.tw.
9 gi.tw.
10 gastric.tw.
11 1 or 2 or 3 or 4 or 5 or 6 or 7 or 8 or 9 or 10
12 pain*.tw.
13 Ache*.tw.
14 Sore*.tw.
15 Discomfort*.tw.
16 Distress*.tw.
17 Cramp*.tw.
18 Disorder$1.tw.
19 Symptom$1.tw.
20 Migraine$1.tw.
21 Epilep*.tw.
22 syndrome$1.tw.
23 colic*.tw.
24 12 or 13 or 14 or 15 or 16 or 17 or 18 or 19 or 20 or 21 or 22 or 23
25 irritable bowel$.tw.
26 lbs.tw.
27 functional dyspepsia.tw.
28 25 or 26 or 27
29 ((stomach* or abdomen* or intestin* or viscera* or tummy or bowel* or belly or gastrointestinal or gi or gastric) adj3 (pain* or Ache* or Sore* or Discomfort* or Distress* or Cramp* or Disorder$1 or Symptom$1 or Migraine$1 or Epilep* or syndrome$1 or colic*)).tw.
30 exp Irritable Bowel Syndrome/
31 exp Colonic Diseases/
32 exp Abdominal Pain/
33 exp Dyspepsia/
34 exp Colic/
35 exp Abdomen, Acute/
36 30 or 31 or 32 or 33 or 34 or 35
37 28 or 29 or 36
38 Recurr*.tw.

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39 Chronic*.tw.
40 Intermittent*.tw.
41 Bout$.tw.
42 spasm*.tw.
43 Transitory.tw.
44 Transient.tw.
45 Functional.tw.
46 Continuous.tw.
47 Paroxysmal.tw.
48 Persistent.tw.
49 Idiopathic.tw.
50 unspecific*.tw.
51 Non specific*.tw.
52 nonspecific*.tw.
53 motility.tw.
54 episod*.tw.
55 38 or 39 or 40 or 41 or 42 or 43 or 44 or 45 or 46 or 47 or 48 or 49 or 50 or 51 or 52 or 53 or 54
56 exp Recurrence/
57 55 or 56
58 37 and 57
59 irritable bowel syndrome*.tw.
60 58 or 59
61 randomized controlled trial.pt.
62 controlled clinical trial.pt.
63 randomi#ed.ab.
64 placebo$.ab.
65 random.ly.ab.
66 trial.ab.
67 groups.ab.
68 exp animals/ not humans.sh.
69 or/61-67
70 69 not 68
71 60 and 70
72 exp Child/
73 exp Adolescent/
74 exp Young Adult/
75 exp Students/
76 Child*.tw.
77 Adolescent*.tw.
78 Young person*.tw.
79 Boy*.tw.
80 Girl*.tw.
81 teen*.tw.
82 Schoolchild*.tw.
83 Young adult*.tw.
84 Youth*.tw.
85 P*ediatric*.tw.
86 Student*.tw.
87 Pupil*.tw.
88 Juvenile*.tw.
89 72 or 73 or 74 or 75 or 76 or 77 or 78 or 79 or 80 or 81 or 82 or 83 or 84 or 85 or 86 or 87 or 88
90 71 and 89

**Embase Ovid (1974 to present)**

Search dates: 11 April 2013 (2272 records); 11 April 2014 (2523 records); 25 March 2015 (250 records); 9 June 2016 (345 records).

1 recurr*.tw.
2 chronic*.tw.
3 intermittent*.tw.
4 bout$.tw.
5 spasm*.tw.
6 transitory.tw.
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7 transient.tw.
8 functional.tw.
9 continu*.tw.
10 paroxysmal.tw.
11 persistent.tw.
12 idiopathic.tw.
13 unspecifi*.tw.
14 non specifi*.tw.
15 nonspecifi*.tw.
16 motility.tw.
17 episod*.tw.
18 or/1-17
19 exp recurrent disease/
20 or/1-19
21 stomach*.tw.
22 abdom*.tw.
23 intestin*.tw.
24 visceral*.tw.
25 tummy.tw.
26 bowel*.tw.
27 belly.tw.
28 gastrointestinal.tw.
29 gi.tw.
30 gastric.tw.
31 or/1-30
31 21 or 22 or 23 or 24 or 25 or 26 or 27 or 28 or 29 or 30
32 pain*.tw.
33 Ache*.tw.
34 Sore*.tw.
35 Distress*.tw.
36 Discomfort*.tw.
37 Cramp*.tw.
38 Disorder$.tw.
39 Symptom$.tw.
40 Migraine$.tw.
41 Epilep*.tw.
42 syndrome$.tw.
43 colic*.tw.
44 or/1-43
44 32 or 33 or 34 or 35 or 36 or 37 or 38 or 39 or 40 or 41 or 42 or 43
45 irritable bowel$.tw.
46 lbs.tw.
47 functional dyspepsia.tw.
48 32 or 33 or 34 or 35 or 36 or 37 or 38 or 39 or 40 or 41 or 42 or 43
49 (stomach* or abdom* or intestin* or visceral or tummy or bowel* or belly or gastrointestinal or gi or gastric) adj3 (pain* or Ache* or Sore* or Discomfort* or Distress* or Cramp* or Disorder$ or Symptom$ or Migraine$ or Epilep* or syndrome$ or colic*).tw.
50 or/1-49
51 exp colic/
52 exp irritable colon/
53 exp abdominal pain/
54 exp dyspepsia/
55 colon disease/
56 or/1-55
56 50 or 51 or 52 or 53 or 54 or 55
57 20 and 56
58 irritable bowel syndrome*.tw.
59 or/1-58
59 57 or 58
60 Clinical trial/
61 Randomized controlled trial/
62 Randomization/
63 Single blind procedure/
64 Double blind procedure/
65 Crossover procedure/
66 Placebo/
67 Random?ed controlled trial$.tw.
Cochrane Database of Systematic Reviews

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CINAHL Healthcare Databases Advanced Search (Cumulative Index to Nursing and Allied Health Literature; 1981 to present)

Search dates: 18 April 2013 (175 records); 11 April 2014 (195 records); 26 March 2015 (21 records); 9 June 2016 (11 records).

1. CINAHL; recurr*.ti,ab;
2. CINAHL; chronic*.ti,ab;
3. CINAHL; intermittent*.ti,ab;
4. CINAHL; (bout OR bouts).ti,ab;
5. CINAHL; spasm*.ti,ab;
6. CINAHL; transitory.ti,ab;
7. CINAHL; transient.ti,ab;
8. CINAHL; functional.ti,ab;
9. CINAHL; continu*.ti,ab;
10. CINAHL; paroxysmal.ti,ab;
11. CINAHL; persistent.ti,ab;
12. CINAHL; idiopathic.ti,ab;
13. CINAHL; unspecific*ti,ab;
14. CINAHL; "non specific"*:ti,ab;
15. CINAHL; nonspecific*.ti,ab;
16. CINAHL; motility.ti,ab;
17. CINAHL; episod*.ti,ab;
18. CINAHL; exp RECURRENCE/;
19. CINAHL; 1 OR 2 OR 3 OR 4 OR 5 OR 6 OR 7 OR 8 OR 9 OR 10 OR 11 OR 12 OR 13 OR 14 OR 15 OR 16 OR 17 OR 18;
20. CINAHL; stomach*.ti,ab;
21. CINAHL; abdom*.ti,ab;
22. CINAHL; intestin*.ti,ab;
23. CINAHL; visceral*.ti,ab;
24. CINAHL; tummy.ti,ab;
25. CINAHL; bowel*.ti,ab;
26. CINAHL; belly.ti,ab;
27. CINAHL; gastrointestinal.ti,ab;
28. CINAHL; gi.ti,ab;
29. CINAHL; gastric.ti,ab;
30. CINAHL; 20 OR 21 OR 22 OR 23 OR 24 OR 25 OR 26 OR 27 OR 28 OR 29;
31. CINAHL; pain*.ti,ab;
32. CINAHL; Ache*.ti,ab;
33. CINAHL; Sore*.ti,ab;
34. CINAHL; Discomfort*.ti,ab;
35. CINAHL; Distress*.ti,ab;
36. CINAHL; Cramp*.ti,ab;
37. CINAHL; (Disorder OR Disorders).ti,ab;
38. CINAHL; (Symptom OR Symptoms).ti,ab;
39. CINAHL; (Migraine OR Migraines).ti,ab;
40. CINAHL; Epilep*.ti,ab;
41. CINAHL; (syndrome OR syndromes).ti,ab;
42. CINAHL; colic*.ti,ab;
43. CINAHL; 31 OR 32 OR 33 OR 34 OR 35 OR 36 OR 37 OR 38 OR 39 OR 40 OR 41 OR 42;
44. CINAHL; 30 AND 43;
45. CINAHL; "irritable bowel"*.ti,ab;
46. CINAHL; ibs.ti,ab;
47. CINAHL; "functional dyspepsia".ti,ab;
48. CINAHL; exp COLIC/;
49. CINAHL; exp IRRITABLE BOWEL SYNDROME/;
50. CINAHL; exp COLONIC DISEASES, FUNCTIONAL/;
51. CINAHL; exp ABDOMINAL PAIN/;
52. CINAHL; exp DYSPESIA/;
53. CINAHL; 45 OR 46 OR 47 OR 48 OR 49 OR 50 OR 51 OR 52;
54. CINAHL; 44 OR 53;
55. CINAHL; 19 AND 54;
56. CINAHL; (irritable AND bowel AND syndrome*) .ti,ab;
57. CINAHL; 55 OR 56;
58. CINAHL; Child*.ti,ab;
59. CINAHL; Adolescence*.ti,ab;
60. CINAHL; "Young person"*.ti,ab;
61. CINAHL; Boy*.ti,ab;
62. CINAHL; Girl*.ti,ab;
63. CINAHL; teen*.ti,ab;
64. CINAHL; Schoolchild*.ti,ab;
65. CINAHL; "Young adult"*.ti,ab;
66. CINAHL; Youth*.ti,ab;
67. CINAHL; Student*.ti,ab;
68. CINAHL; Pupil*.ti,ab;
69. CINAHL; Juvenile*.ti,ab;
70. CINAHL; exp CHILD/;
71. CINAHL; exp STUDENTS/;
72. CINAHL; 70 OR 71;
73. CINAHL; Pediatric*.ti,ab;
74. CINAHL; Paediatric*.ti,ab;
75. CINAHL; 67 OR 68 OR 69 OR 70 OR 71 OR 72 OR 73 OR 74;
76. CINAHL; 63 OR 64 OR 65 OR 66;
77. CINAHL; 58 OR 59 OR 60 OR 61 OR 62;
78. CINAHL; 70 OR 73 OR 74 OR 75;
79. CINAHL; 57 AND 78;
80. CINAHL; exp RANDOMIZED CONTROLLED TRIALS/;
81. CINAHL; random*.ti,ab;
82. CINAHL; "clin" trial"*.ti,ab;
83. CINAHL; (singl* OR doubl* OR tripl* OR trebl*).ti,ab;
84. CINAHL; (mask* OR blind*).ti,ab;
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85. CINAHL; 83 AND 84;
86. CINAHL; "random allocate*"ti,ab;
87. CINAHL; "random assign*"ti,ab;
88. CINAHL; exp RANDOM ASSIGNMENT/;
89. CINAHL; exp CLINICAL TRIALS/;
90. CINAHL; exp META ANALYSIS/;
91. CINAHL; 88 OR 89 OR 90;
92. CINAHL; 80 OR 81 OR 82 OR 85 OR 86 OR 87;
93. CINAHL; 91 OR 92;
94. CINAHL; 79 AND 93;

PsycINFO Ovid (1806 to present)

Search dates: 18 April 2013 (238 records); 11 April 2014 (757 records); 25 March 2015 (47 records); 9 June 2016 (87 records).

1 stomach*.tw.
2 abdom*.tw.
3 intestin*.tw.
4 viscera*.tw.
5 tummy.tw.
6 bowel*.tw.
7 belly.tw.
8 gastrointestinal.tw.
9 gi.tw.
10 gastric.tw.
11 1 or 2 or 3 or 4 or 5 or 6 or 7 or 8 or 9 or 10
12 pain*.tw.
13 Ache*.tw.
14 Sore*.tw.
15 Discomfort*.tw.
16 Distress*.tw.
17 Cramp*.tw.
18 Disorder$.1.tw.
19 Symptom$.1.tw.
20 Migraine$.1.tw.
21 Epilep*.tw.
22 syndrome$.1.tw.
23 colic*.tw.
24 12 or 13 or 14 or 15 or 16 or 17 or 18 or 19 or 20 or 21 or 22 or 23
25 irritable bowel$.tw.
26 lbs.tw.
27 functional dyspepsia.tw.
28 25 or 26 or 27
29 ((stomach* or abdom* or intestin* or viscera* or tummy or bowel* or belly or gastrointestinal or gi or gastric) adj3 (pain* or Ache* or Sore* or Discomfort* or Distress* or Cramp* or Disorder$1 or Symptom$1 or Migraine$1 or Epilep* or syndrome$1 or colic*)).tw.
30 exp Irritable Bowel Syndrome/
31 exp Dyspepsia/
32 recurr*.tw.
33 chronic*.tw.
34 intermittent*.tw.
35 bout$.1.tw.
36 spasm*.tw.
37 transitory.tw.
38 transient.tw.
39 functional.tw.
40 continu*.tw.
41 paroxysmal.tw.
42 persistent.tw.
43 idiopathic.tw.
44 unspecific*.tw.
45 non specific*.tw.
46 nonspecific*.tw.
47 motility.tw.
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(ab,ti(Stomach*) OR ab,ti(Abdom*) OR ab,ti(Intestin*) OR ab,ti(Viscera*) OR ab,ti(Tummy) OR ab,ti(Bowel*) OR ab,ti(Belly) OR ab,ti(Gastrointestinal) OR ab,ti(GI) OR ab,ti(gastric)

AND

(ab,ti(Pain*) OR ab,ti(Ache*) OR ab,ti(Sore*) OR ab,ti(Discomfort*) OR ab,ti(Bowel*) OR ab,ti(Distress*) OR ab,ti(Disorder) OR ab,ti(Disorders) OR ab,ti(Symptom*) OR ab,ti(Symptoms) OR ab,ti(Migraine*) OR ab,ti(Epilep*) OR ab,ti(Syndrome) OR ab,ti(Syndromes) OR ab,ti(colic*)

AND

(ab,ti(Recurr*) OR ab,ti(Chronic*) OR ab,ti(Intermittent*) OR ab,ti(Episode*) OR ab,ti(Bout) OR ab,ti(bouts) OR ab,ti(Spasm*) OR ab,ti(Transitory) OR ab,ti(Transient) OR ab,ti(Functional) OR ab,ti(Continu*) OR ab,ti(Paroxysmal) OR ab,ti(Persistent) OR ab,ti(Idiopathic) OR ab,ti(Unspecifi*) OR ab,ti(Non specific*) OR ab,ti(motility))

OR

(ab,ti(IRRITABLE BOWEL) OR ab,ti(ibs) OR ab,ti(functional dyspepsia))

Allied and Complementary Medicine Healthcare Databases Advanced Search (AMED; 1985 to present)

Search dates: 18 April 2013 (63 records); 11 April 2014 (74 records); 25 March 2015 (1 record); 9 June 2016 (1 record).

1. AMED; Recurr*.ti,ab;
2. AMED; Chronic*.ti,ab;
3. AMED; Intermittent*.ti,ab;
4. AMED; Episode*.ti,ab;
5. AMED; (Bout OR Bouts).ti,ab;
6. AMED; Spasm*.ti,ab;
7. AMED; Transitory.ti,ab;
8. AMED; Transient.ti,ab;
9. AMED; Functional.ti,ab;
10. AMED; Continu*.ti,ab;
11. AMED; Paroxysmal.ti,ab;
12. AMED; Persistent.ti,ab;
13. AMED; Idiopathic.ti,ab;
14. AMED; Unspecific*.ti,ab;
15. AMED; "Non specific**”.ti,ab;
16. AMED; Nonspecific*.ti,ab;
17. AMED; Motility.ti,ab;
18. AMED; exp RECURRENT/;
19. AMED; 1 OR 2 OR 3 OR 4 OR 5 OR 6 OR 7 OR 8 OR 9 OR 10 OR 11 OR 12 OR 13 OR 14 OR 15 OR 16 OR 17 OR 18;
20. AMED; Pain*.ti,ab;
21. AMED; Ache*.ti,ab;
22. AMED; Sore*.ti,ab;
23. AMED; Discomfort*.ti,ab;
24. AMED; Distress*.ti,ab;
25. AMED; Cramp*.ti,ab;
26. AMED; Disorder OR Disorders).ti,ab;
27. AMED; (Symptom OR Symptoms).ti,ab;
28. AMED; (Migraine OR Migraines).ti,ab;
29. AMED; Epilep*.ti,ab;
30. AMED; Colic*.ti,ab;
31. AMED; (Syndrome OR Syndromes).ti,ab;
32. AMED; 20 OR 21 OR 22 OR 23 OR 24 OR 25 OR 26 OR 27 OR 28 OR 29 OR 30 OR 31;
33. AMED; Stomach*.ti,ab;
34. AMED; Abdom*.ti,ab;
35. AMED; Intestin*.ti,ab;
36. AMED; Viscera*.ti,ab;
37. AMED; Tummy.ti,ab;
38. AMED; Bowel*.ti,ab;
39. AMED; Belly.ti,ab;
40. AMED; Gastrointestinal.ti,ab;
41. AMED; GI.ti,ab;
42. AMED; Gastric.ti,ab;

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43. AMED; 33 OR 34 OR 35 OR 36 OR 37 OR 38 OR 39 OR 40 OR 41 OR 42;
44. AMED; 32 AND 43;
45. AMED; "Colonic disease".ti,ab;
46. AMED; "Irritable bowel".ti,ab;
47. AMED; IBS.ti,ab; 86
48. AMED; "Functional dyspepsia".ti,ab;
49. AMED; exp IRRITABLE BOWEL SYNDROME/;
50. AMED; exp COLONIC DISEASE/;
51. AMED; exp ABDOMINAL PAIN/;
52. AMED; exp DYSPESPIA/;
53. AMED; 45 OR 46 OR 47 OR 48 OR 49 OR 50 OR 51 OR 52;
54. AMED; 44 OR 53;
55. AMED; 19 AND 54;
56. AMED; (irritable AND bowel AND syndrome*).ti,ab;
57. AMED; Child*.ti,ab;
58. AMED; Adolescent*.ti,ab;
59. AMED; Boy*.ti,ab;
60. AMED; Girl*.ti,ab;
61. AMED; teen*.ti,ab;
62. AMED; Schoolchild*.ti,ab;
63. AMED; "Young adult".ti,ab;
64. AMED; Youth*.ti,ab; 767 results.
65. AMED; (Pediatric* OR Paediatric*).ti,ab;
66. AMED; Student*.ti,ab;
67. AMED; Pupil*.ti,ab;
68. AMED; Juvenile*.ti,ab;
69. AMED; "Young person".ti,ab;
70. AMED; exp CHILD/;
71. AMED; exp ADOLESCENT/;
72. AMED; exp STUDENTS/;
73. AMED; 57 OR 58 OR 59 OR 60 OR 61 OR 62 OR 63 OR 64 OR 65 OR 66 OR 67 OR 68 OR 69 OR 70 OR 71 OR 72;
74. AMED; 55 OR 56;
75. AMED; 74 AND 73;

LILACS (Latin American and Caribbean Health Science Information database; lilacs.bvsalud.org/en; all available years)
Search dates: 19 April 2013 (11 records); 11 April 2014 (13 records); 26 March 2015 (0 records); 9 June 2016 (0 records).

((Pt randomized controlled trial OR Pt controlled clinical trial OR Mt randomized controlled trials OR Mt random allocation OR Mt double-blind method OR Mt single-blind method) AND NOT (Ct animal AND NOT (Ct human and Ct animal)) OR (Pt clinical trial OR Ex E05.318.760.535$ OR (Tw clin$ AND (Tw tria$l OR Tw ensa$l OR Tw estud$ OR Tw experim$ OR Tw inves$tiga$s)) OR ((Tw sing$l OR Tw simple$ OR Tw doubl$s OR Tw doble$s OR Tw duplo$s OR Tw trebl$s OR Tw tri$p)s) (tria$l$ OR estud$ OR experim$ OR inves$tiga$s OR singl$s OR simple$s OR Tw doubl$s OR Tw doble$s OR Tw duplo$s OR Tw trebl$s OR Tw tri$p)s AND (Tw blind$ OR Tw cego$s OR Tw ciego$s OR TW mask$ OR Tw mascar$s)) OR Mt placebos OR Tw placebo$s OR (Tw random$ OR Tw randon$ OR Tw casual$ OR TW acaso$s OR TW azar OR TW aleator$ OR Mt research design) AND NOT (Ct animal AND NOT (Ct human and Ct animal)) OR (Ct comparative study OR Ex E05.337$ OR Mt follow-up studies OR Mt prospective studies OR Mt control$s OR Tw prospectiv$s OR Tw volunt$s OR Tw volunter$s) AND NOT (Ct animal AND NOT (Ct human and Ct animal))) [Palavras]

and ([(recurs$ or chronic$ or intermittent$ or bout or bouts or spasms$ or transitory or transient or functional or continu$ or Paroxysmal or Persistent or idiopathic or unspecific$ or Non specific$ or nonspecific$s or motility or episode$s) [Palavras] and (pain$ or ache$ or sore$ or discomfort$ or distress$ or cramp$s or colic$s or disorder or disorders or symptom or symptoms or Migraine$ or Epilep*$ or syndrome$s) and (stomach$ or abdomen$ or intestine$s or viscera$ or tummy$s or bowel$s or belly or gastrointestinal or gi or gastric)) [Palavras]

OpenGrey (www.opengrey.eu; 1980 to present)
Search dates : 19 April 2013 (1 record); 11 April 2014 (1 record); 26 March 2015 (0 records); 9 June 2016 (0 records).

Irritable bowel syndrome*
Lbs
functional dyspepsia
Chronic$ AND (abdom$ OR stomach$ OR intestine$s OR viscera$ OR tummy OR bowel$ OR belly or gastrointestinal OR gi OR gastric))
Recur$ AND (abdom$ OR stomach$ OR intestine$s OR viscera$ OR tummy OR bowel$ OR belly or gastrointestinal OR gi OR gastric))
Intermittent$ AND (abdom$ OR stomach$ OR intestine$s OR viscera$ OR tummy OR bowel$ OR belly or gastrointestinal OR gi OR gastric))
Bout$ AND (abdom$ OR stomach$ OR intestine$s OR viscera$ OR tummy OR bowel$ OR belly or gastrointestinal OR gi OR gastric))
spasm* AND (abdom* OR stomach* OR intestin* OR viscera* OR tummy OR bowel* OR belly or gastrointestinal OR gi OR gastric))
Transitory AND (abdom* OR stomach* OR intestin* OR viscera* OR tummy OR bowel* OR belly or gastrointestinal OR gi OR gastric))
Transient AND (abdom* OR stomach* OR intestin* OR viscera* OR tummy OR bowel* OR belly or gastrointestinal OR gi OR gastric))
Functional AND (abdom* OR stomach* OR intestin* OR viscera* OR tummy OR bowel* OR belly or gastrointestinal OR gi OR gastric))
Continu* AND (abdom* OR stomach* OR intestin* OR viscera* OR tummy OR bowel* OR belly or gastrointestinal OR gi OR gastric))
Paroxysmal AND (abdom* OR stomach* OR intestin* OR viscera* OR tummy OR bowel* OR belly or gastrointestinal OR gi OR gastric))
Persistent AND (abdom* OR stomach* OR intestin* OR viscera* OR tummy OR bowel* OR belly or gastrointestinal OR gi OR gastric))
Idiopathic AND (abdom* OR stomach* OR intestin* OR viscera* OR tummy OR bowel* OR belly or gastrointestinal OR gi OR gastric))
Unspecific* AND (abdom* OR stomach* OR intestin* OR viscera* OR tummy OR bowel* OR belly or gastrointestinal OR gi OR gastric))
Moality AND (abdom* OR stomach* OR intestin* OR viscera* OR tummy OR bowel* OR belly or gastrointestinal OR gi OR gastric))
Episod* AND (abdom* OR stomach* OR intestin* OR viscera* OR tummy OR bowel* OR belly or gastrointestinal OR gi OR gastric))

ClinicalTrials.gov (clinicaltrials.gov; 2007 to present)
Search dates: 11 April 2014 (69 records); 26 March 2015 (35 records); 9 June 2016 (62 records).

“irritable bowel” OR “abdominal pain” in the condition field. Limited to children.

World Health Organization International Clinical Trials Registry Platform (WHO ICTRP; apps.who.int/trialsearch; 1999 to present)
Search dates: 11 April 2014 (106 records); 26 March 2015 (4 records); 9 June 2016 (32 records).

“irritable bowel” OR “abdominal pain” in the condition field. Limited to children and interventional studies.

Appendix 2. Additional methods

We detailed methods for each of the topics below in our protocol (Martin 2014a), but did not use them in the review because we had insufficient data to perform a meta-analysis.

- Types of outcome measures.
- Unit of analysis issues.
- Dealing with missing data.
- Assessment of heterogeneity.
- Assessment of reporting bias.
- Data synthesis.
- Subgroup analysis and investigation of heterogeneity.
- Sensitivity analysis.

We describe these methods, which have been archived for use in future updates of this review, in the table below.

### Types of outcome measures

We expect studies to vary in their duration of postintervention follow-up. We will therefore group studies according to duration of follow-up: immediate outcome measurement, short term (less than three months), medium term (three to 12 months), and long term (greater than 12 months).

### Unit of analysis issues

If we find the following three types of trials, we will consider their results as per guidance in the Cochrane Handbook for Systematic Reviews of Interventions (Higgins 2011c, section 9.3).

#### Cross-over trials

For cross-over trials with random allocation to period and an appropriate washout period, we will include the relevant effect estimate within the meta-analysis, using the generic inverse variance method in Review Manager 5 (Review Manager 2014). An appropriate washout period may vary with the interventions (including drug pharmacokinetics) and outcome measurements. Considering that RAP can be a stable and chronic condition, a washout period of several weeks may be sufficient.

#### Cluster-RCTs
Cluster-randomised trials randomise groups of people rather than individuals. For each cluster-randomised trial, we will first determine whether or not the data incorporate sufficient controls for clustering (such as robust standard errors or hierarchical linear models). If data do not have proper controls, then we will attempt to obtain an appropriate estimate of the data’s intracluster correlation coefficient. If we cannot find an estimate in the trial report, then we will request an estimate from the trial report authors. If the authors do not provide an estimate, if possible, we will obtain one from a similar study and conduct a sensitivity analysis to determine if the results are robust when different values are imputed. We will do this according to procedures described in Higgins 2011c (section 9.3). This will prevent the meta-analysis from being based on clustered data that has not been properly controlled.

**Trials with multiple intervention groups**

This is a common scenario. To avoid any unit of analysis errors in the meta-analysis, we will use the following approach for a study that could contribute multiple comparisons:

- We will only analyse the interventions together if they are clinically similar, that is testing the same class of drug. In this situation we will not split the control group, but instead will combine the intervention groups to enable a single pair-wise comparison for the meta-analysis. If the interventions are similar enough to be in a single meta-analysis but cannot be combined, then we will split the control group.

- If the interventions are not similar, we will perform separate meta-analyses. Consequently, a single study could contribute data to different meta-analyses (e.g. if the interventions involve eliminating different classes of drugs); this does not require the control group to be split.

We will not perform a multiple-treatment meta-analysis, as the clinical heterogeneity would mean the results have little clinical meaning.

**Dealing with missing data**

We may carry out a sensitivity analysis to establish if the inclusion of studies with high levels of missing data significantly alters the findings of the review.

**Assessment of reporting bias**

**Publication bias**

If we identify sufficient trials (at least 10), we will use the outcome data to produce a funnel plot to investigate the likelihood of overt publication bias (Sutton 2000). Any asymmetry of the funnel plot may indicate possible publication bias. We will explore other reasons for asymmetry such as poor methodological quality or heterogeneity. We will look for publication bias by comparing the results of the published and unpublished data.

**Assessment of heterogeneity**

We will describe statistical heterogeneity (observed variability in study results that is greater than that expected to occur by chance) by calculating $I^2$ (Higgins 2003). $I^2$ describes approximately the proportion of variation in point estimates due to heterogeneity rather than sampling error. An $I^2$ more than 50% may indicate significant heterogeneity.

We will use a Chi² test to further assess the role of heterogeneity on the strength of the evidence. We will regard any result with a $P$ value lower than 0.10 as indicative of heterogeneity. We will interpret this cautiously and use it to help quantify the impact of heterogeneity on the results of the meta-analysis (Higgins 2003).

**Data synthesis**

We will use Review Manager 5 for statistical analysis (Review Manager 2014). Two review authors (RAA, AB, JTC, TVND, or AEM) will independently enter data into Review Manager 5. We will report summary statistics for continuous data as mean differences or standardised mean differences using a random-effects model. For dichotomous data, we will report odds ratios using a random-effects model. We intend to use a a random-effects model as we anticipate significant statistical and clinical heterogeneity.

**Subgroup analysis and investigation for heterogeneity**

If sufficient trials are available, we will examine the following subgroups to explore clinical heterogeneity:

- type of RAP (as defined by the Rome III criteria) (Rasquin 2006);
- age; and
Subgroup analysis can be misleading because the studies may not be designed and powered to show difference within subgroups. We will therefore undertake subgroup analyses with caution.

Sensitivity analysis

We will conduct primary analyses based on available data on the outcomes of interest. Following this, we will use a sensitivity analysis to assess the robustness of conclusions in relation to two aspects of study design:

1. the effect of inadequate allocation concealment, by the removal of studies judged to be at high or unclear risk of bias for that domain; and
2. the effect of inadequate blinding to treatment allocation by outcome assessors, by the removal of studies judged to be at high or unclear risk of bias for that domain.

We will also conduct a sensitivity analysis to establish the effect of missing data on the estimate of treatment effect, by performing the analysis with and without the studies with significant missing data to determine if this alters the conclusions.

Footnotes

RAP: recurrent abdominal pain
RCT: randomised controlled trial

WHAT'S NEW

<table>
<thead>
<tr>
<th>Date</th>
<th>Event</th>
<th>Description</th>
</tr>
</thead>
<tbody>
<tr>
<td>10 June 2016</td>
<td>New citation required and conclusions have changed</td>
<td>Following a new search in April 2013 and updated searches in April 2014, March 2015, and June 2016, we added 13 new studies.</td>
</tr>
<tr>
<td>31 January 2015</td>
<td>New search has been performed</td>
<td>The review supersedes the previous version, following a new protocol and new search up to June 2016 (see Published notes).</td>
</tr>
</tbody>
</table>

CONTRIBUTIONS OF AUTHORS

- Review design: AEM, SL
- Review co-ordination: AEM
- Data collection:
  * Search strategy design: AEM, AB
  * Searches undertaken: AEM, AB
  * Search results screened: AEM, TVND, RAA, AB, JTC, RW
  * Retrieval of papers: AEM, AB
  * Paper screening and appraisal, and extraction of data: AEM, TVND, RAA, AB, JTC, RW
  * Writing to authors for additional information: AEM, AB, RAA, TVND
  * Entering the data into Review Manager 5: AEM, TVND, RAA, AB, JTC
- Analysis of the data: AEM, TVND, RAA, AB, JTC, SL
- Interpretation of the data:
  * Methodological perspective: AEM, TVND, RAA, AB, JTC
  * Clinical perspective: AEM, TVND, SL
DECLARATIONS OF INTEREST

The work of the evidence synthesis team is funded by the National Institute for Health Research Collaboration for Leadership in Applied Health Research and Care South West Peninsula (PenCLAHRC). The Funder has had no role in the review itself.

Alice E Martin: none known.
Tamsin V Newlove-Delgado: none known.
Rebecca A Abbott: none known.
Alison Bethel: none known.
Joanna Thompson-Coon: none known.
Rebecca Whear: none known
Stuart Logan: none known.

The authors who practice clinical paediatrics are Alice E Martin and Stuart Logan. Alice is a Paediatric Trainee and works under the guidance of various consultant paediatricians. Stuart is a Consultant Paediatrician and treats children according to current best evidence, in light of their preference. There are therefore no conflicts of interest with this review.

SOURCES OF SUPPORT

Internal sources

• National Institute for Health Research Collaboration for Leadership in Applied Health Research and Care South West Peninsula (PenCLAHRC), UK.

Funds the work of the evidence synthesis team.

External sources

• None, Other.

DIFFERENCES BETWEEN PROTOCOL AND REVIEW

The differences between the protocol and the review are detailed in the Additional Methods Table in Appendix 2.

Rebecca Whear joined the team following the publication of the protocol.

NOTES

This is a new review, which supersedes a previously published review (Huertas-Ceballos 2008a).

INDEX TERMS

Medical Subject Headings (MeSH)

Abdominal Pain [*drug therapy]; Randomized Controlled Trials as Topic; Recurrence; Treatment Outcome

MeSH check words

Adolescent; Child; Child, Preschool; Humans