1	Decreasing subseasonal temperature variability in the northern extratropics
2	attributed to human influence
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Changes in subseasonal temperature variability are linked with the altered probability of 8 weather extremes and have important impacts on society and ecological systems. Earlier 9 studies based on observations up to 2014 have shown a general decrease in subseasonal 10 temperature variability over Northern Hemisphere extratropical land. However, these 11 12 changes have been confined to specific regions and seasons, have limited statistical significance, and human influence is yet to be determined. Here we show using up-to-date 13 observations and climate model simulations that a human fingerprint, or pattern, of change 14 in subseasonal variability has recently emerged over the Northern Hemisphere extratropics. 15 The fingerprint features decreased near-surface air temperature variability over land in the 16 high-northern latitudes in autumn, further extending into mid-latitudes in winter. Using 17 large ensembles of single-forcing model experiments, we attribute the pattern of reduced 18 19 temperature variability primarily to increased anthropogenic greenhouse gas 20 concentrations, with anthropogenic aerosols playing a secondary role. Our results reveal that human influence is now detectable in hemispheric-wide day-to-day temperature variability 21

and motivates research into the impacts of reduced temperature volatility on societal and
 ecological systems.

24 Main

While the human influence on time averaged temperature changes is unequivocal<sup>1</sup>, the influence 25 on day-to-day (or week-to-week, and so on) temperature variability is less certain. Observations 26 over recent decades show trends in subseasonal temperature variability that vary by region and 27 season. Reductions in temperature variability have been observed over the northern high-latitudes 28 in autumn<sup>2</sup> and over North America in both winter<sup>3</sup> and summer<sup>4</sup>, while increased variability has 29 been observed over Eurasia during summer<sup>4</sup>. Human influence on temperature variability has yet 30 to be determined, but is expected to depend on various processes including changes in thermal 31 advection linked to altered background temperature gradients<sup>2,5–7</sup>, atmospheric circulation 32 variability<sup>8</sup>, snow cover extent<sup>9</sup>, and soil moisture-temperature feedbacks<sup>10</sup>. 33

Climate models forced with projected increases in greenhouse gas concentrations show a 34 robust decrease in subseasonal temperature variability over the mid-to-high latitudes, during all 35 seasons except summer<sup>2</sup>. This decrease in variability has been attributed in some studies to the 36 reduction of temperature gradients associated with the faster warming of Arctic compared with 37 lower latitudes<sup>2,5,7,11</sup>, also known as Arctic amplification. In contrast, other studies have argued 38 that temperature variability in the mid-latitudes could increase due to Arctic-forced changes in 39 atmospheric circulation variability<sup>12–14</sup>. It has also been suggested that models may not accurately 40 capture these potential circulation responses to Arctic warming<sup>15</sup>, calling into question the 41 projected decrease in temperature variability. 42

In this study, we apply tried-and-tested formal detection methods to identify the human
fingerprint of subseasonal temperature variability in observations over the 1979-2020 period. The

fingerprint is determined from the spatial pattern of the subseasonal temperature variability response to external forcing in large ensembles of simulations from two climate models. We find that the fingerprint has recently become detectable above internal variability in observations during autumn and winter. Using simulations that isolate individual climate drivers, we then attribute the observed reduction in temperature variability primarily to anthropogenic greenhouse gas emissions.

#### 51 **Trends in subseasonal temperature variability**

52 We start by examining spatial maps of linear trends in subseasonal near-surface temperature 53 variability over land from 1979-2020 in two reanalysis datasets (ERA5 and NCEP-DOE reanalysis 2). The temperature variability is determined by calculating the standard deviation of daily 54 temperature anomalies for each season, resulting in a 41-year time series at each grid point (see 55 Methods). During autumn (September-October-November; SON), both reanalyses show 56 statistically significant decreases over Northern Canada and Northern Eurasia (Fig 1a,c). In winter 57 (December-January-February; DJF), the decreases in temperature variability extend further into to 58 59 the mid-latitudes, covering most of North America and Eurasia (Fig 1b,d). The two reanalyses depict highly similar spatial patterns of trends, but the magnitude of the variability reduction is 60 61 greater in the NCEP reanalysis.

To check whether the trends seen in the reanalyses accurately reflect observed trends, we also examined trends from the Berkeley Earth daily gridded temperature data (Extended Data Fig. 1). Over the common time period (data were only available up to winter 2018 for Berkeley Earth data), there is strong agreement in trends between gridded observations and reanalyses, particularly

with ERA5. This gives high confidence that the trends in the reanalyses faithfully capture observedchanges.

To test whether the trends could be externally forced (e.g. from increasing greenhouse gas 68 concentrations in the atmosphere), we examine analogous trends in subseasonal temperature 69 variability from model simulations with observed historical forcings over 1979-2005 and then 70 projected forcings from 2006-2020. We use large ensembles of simulations from two Earth System 71 72 Models (CanESM2 and CESM1) to test the robustness of the projected changes. Both models show ensemble-mean reductions in temperature variability over mid- and high-latitudes, in agreement 73 with the reanalyses (Fig 1e-h). The modelled trends show a smoother spatial structure compared 74 75 to reanalysis, which is expected because internal variability is averaged out when taking an ensemble mean. Both models show similar seasonality of trends compared to the reanalyses, with 76 reduced variability largely confined to the high latitudes in autumn, but extending into the mid-77 latitudes in winter. The similarity of the simulated trends (which consist of only the forced 78 79 response) to observed trends (which consist of both a forced component and internal variability), suggests that the observed trends may reflect a forced response. In the next section, we will 80 determine whether this forced trend seen in the models has emerged above the background internal 81 82 variability in observations.

In Extended Data Fig. 2, we show the subseasonal temperature variability trends, in reanalyses and models, in spring (March-April-May; MAM) and summer (June-July-August; JJA). Both simulated and observed trends are weaker in these seasons, but the models show reduced temperature variability over high latitudes during spring, and weakly increased temperature variability over the high latitudes in summer. Compared to autumn and winter, there is less consistency in the spatial pattern of the trends in temperature variability across the reanalyses, and

between the reanalyses and models. The weaker signals in spring and summer likely results from
the different physical processes that drive changes temperature variability during these seasons.
Because of these weaker trends in spring and summer, we will focus on autumn and winter for the
rest of the analysis.

## 93 Fingerprint analysis

The strong similarity in the observed and simulated trend patterns motivates a fingerprint analysis 94 to determine whether human influence is detectable in observations. We use standard 95 96 fingerprinting methods that have been used to detect human influence on many aspects of the climate system (see Methods)<sup>16–18</sup>. First, we define the fingerprints to be the leading Empirical 97 Orthogonal Function (EOF) from the ensemble-mean temperature variability anomalies (relative 98 99 to climatology) of each of the two models. These are shown in the Extended Data Fig. 3, but are 100 nearly identical in pattern to the linear trends shown Fig. 1e-h. Next, we calculate the signal time-101 series, by projecting the temperature variability anomalies from the reanalyses and observations onto the model fingerprints. Figure 2 shows these signal time-series in autumn and winter, using 102 the fingerprints defined from each model. In both autumn and winter, all signal time-series 103 104 generally increase over the past 41 years, indicating growing similarity of observed anomalies to 105 the model fingerprints. The different reanalyses and observations are in good agreement on the trend and interannual variability of the signal time-series, again providing confidence in robustness 106 107 of the result.

We next examine whether the signals of the fingerprint (from Fig. 2) have emerged above the background noise. We calculate signal-to-noise ratios by comparing the magnitude of trend in the signal time-series to the standard deviation of the magnitude of trends of the same length from time series containing only noise (see Methods). Figure 3 shows the signal-to-noise ratios, for 112 increasing trend length starting in 1979, calculated separately using the fingerprint from each model. In autumn, the observed signal-to-noise ratios became statistically significant (at the 5% 113 level) between 2005 and 2007, depending on the observational/reanalysis product and model used 114 to determine the fingerprint. Over the 1979 to 2019 period, the fingerprint is highly detectable in 115 both reanalyses with signal-to-noise ratios ranging from 4.7 to 5.3. In winter, the time of detection 116 117 has a stronger dependence on the observations and model used, with time of detection ranging from 2005 to 2018. However, regardless of the observations and model used, the signal-to-noise 118 ratio is statistically significant for trends ending in 2020, with signal-to-noise ratios ranging from 119 120 2.5 to 3.7. Also shown in Figure 3 are the signal-to-noise ratios from the individual realizations of the models. The observations and reanalyses fit within the ensemble range, indicating that the 121 122 magnitudes of their changes are within the model spread.

So far, our analysis has only examined the temperature variability over land because the reanalysis data over land is likely better constrained by observations, and because the daily observational data from Berkeley Earth only includes land stations. Keeping the uncertainty of the reanalysis data in mind, we found that signal-to-noise ratios are even higher if ocean regions are included in the analysis (Extended Data Fig. 4), because both the reanalyses and models show strong reductions in subseasonal temperature variability over high-latitude ocean regions near the sea ice edge (Extended Data Fig 5).

## 130 Attribution to human influence

We now use single-forcing large-ensemble experiments to attribute the externally forced response of subseasonal temperature variability to individual climate drivers. In response to only increased anthropogenic greenhouse gas concentrations, both models show reduced temperature variability (Fig 4a-d) similar to that in the historical experiment (pattern correlations range from 0.71 to 0.93, 135 depending on model and season), indicating that the anthropogenic greenhouse gases are the primary driver of the forced response. In autumn, both models also show reduced temperature 136 variability in high latitudes in response to anthropogenic aerosol reductions over recent decades. 137 The spatial pattern of the autumn response to aerosol forcing is similar to that to greenhouse gas 138 forcing (pattern correlations of 0.50 and 0.75 in CanESM2 and CESM1 respectively), but with 139 140 weaker magnitude. In winter, there is less consistency in the responses to aerosols between the two models, which partly explain the minor differences in the winter responses to historical forcing in 141 the two models (Fig. 1f, h). 142

Previous work has attributed reduced temperature variability in model projections to the 143 decreased meridional temperature gradients associated with Arctic amplification<sup>2,5</sup>. This is 144 supported by targeted modelling experiments forced with imposed sea ice loss and Arctic 145 amplification in isolation<sup>7,11,19–22</sup>. Consistent with this interpretation, our simulations display 146 seasonal-mean warming (Extended Data Fig. 6) and reduced meridional temperature gradients 147 (Extended Data Fig. 7). In response to greenhouse gas forcing only, there is clear Arctic 148 amplification and associated reduced meridional temperature gradients, which have similar spatial 149 patterns to the reduced temperature variability. Furthermore, aerosol forcing has also caused 150 151 amplified warming in the Arctic, leading to reduced meridional temperature gradients. Other processes such decreases in land-sea temperature gradients<sup>6</sup> and decreases in snow cover extent<sup>9</sup> 152 may also contribute to the decrease in temperature variability. 153

154 Implications for future change

There has been considerable disagreement, between model-based and observation-based studies, on the response of midlatitude circulation and winter temperature extremes to Arctic amplification<sup>15</sup>. However, we find a convergence between modelled and observed trends in subseasonal winter temperature variability, increasing confidence in model projections. Consistent with previous investigations into the mechanisms<sup>2,5,7,11,19–22</sup>, we have shown that temperature variability is linked to changes in the large-scale meridional temperature gradients<sup>5</sup>, which are highly similar in observations and models. The human influence on atmospheric circulation variability, which is undetectable in observations<sup>23</sup> and inconsistent between models<sup>24</sup>, likely plays a smaller role.

Our results demonstrate that human influence on climate is now detectable in subseasonal 164 temperature variability, not only in shifts in climate averages. The narrowing of the distribution of 165 daily temperatures that we find implies that cold extremes are becoming less frequent and less 166 167 severe at a faster rate than predicted by the shift in mean temperature; and conversely, hot extremes are becoming more frequent and more severe at a slower rate than predicted by a shift in the mean. 168 However, changes in higher order moments of the temperature distribution<sup>8,25</sup>, such as skewness 169 170 and kurtosis that were not examined here, could also contribute to frequency and severity of extremes. The reduced temperature variability we have detected over recent decades is expected 171 to continue with ongoing greenhouse gas emissions<sup>2</sup>. Our work encourages investigation into the 172 impacts of human-induced changes in temperature variability on society and ecosystems<sup>26–29</sup>. 173

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## 241 Methods

#### 242 Reanalysis and observations

We use daily-mean near-surface (2-meter) temperature data from two reanalysis products: ERA5 reanalysis<sup>30</sup> and NCEP-DOE reanalysis 2<sup>31</sup>. Data is used from March 1979 to February 2020, resulting in 41 years for each season. We also use gridded observations of daily near-surface temperature from Berkeley Earth<sup>32</sup> from March 1979 to February 2018. The observations from Berkeley Earth were only available up to 2018, but we chose to use reanalysis data up to 2020 to include the latest data.

#### 249 Modelling experiments

We use output from large-ensemble climate model experiments from two models: the Canadian 250 Earth System Model version 2 (CanESM2)<sup>33,34</sup> and the Community Earth System Model version 251 1 (CESM1)<sup>35,36</sup>. For CanESM2, four experiments with different forcings were performed: all-252 forcings (ALL), only anthropogenic aerosols (AER), only Natural forcings (NAT), and only 253 254 stratospheric ozone forcings (OZ). Each experiment consists of 50 ensemble members with 255 historical forcing from 1950 to 2005 and RCP8.5 forcing from 2006 to 2100. Historical emissions over 2006-2020 closely match the emissions from RCP8.5<sup>37</sup>, justifying the comparison. Each of 256 257 the 50 ensemble members was initiated from a different initial condition, so differences between ensemble members arise only due to internal variability. The response to anthropogenic 258 259 greenhouse forcing is determined as a residual from subtracting the ensemble-mean trends of the 260 AER, NAT and OZ from the ALL simulations. This approach means that land use changes are included in the response to anthropogenic greenhouse forcing for the CanESM2 simulations. 261

The CESM1 simulations consist of a 40-member all-forcing experiment, with historical forcing from 1920 2005 and RCP8.5 forcing from 2006 to 2100. The single forcing experiments from CESM1 were performed differently than those with CanESM2. These experiments were identical to the CESM1 all forcing ensemble, except one forcing at a time was held fixed at 1920
levels. Three such experiments were performed: a 20-member ensemble with fixed greenhouse
gases, a 20-member ensemble with fixed industrial aerosols, and a 15-member ensemble with fixed
biomass burning aerosols. The responses to individual forcings (Fig 4) were determined by
subtracting the ensemble-mean of each of fixed forcing experiments from the ensemble mean of
the all-forcing experiment.

### 271 Temperature variability calculation

272 All daily 2-meter temperatures from reanalysis, observations and models were first interpolated to 273 a common 2°x 2° grid. We have confirmed that the results are insensitive to reasonable changes in the resolution that the analysis is performed on. Daily temperature anomalies were calculated 274 by first subtracting the climatological (1979-2020) average temperature for each day and at each 275 276 grid point. For model data, the climatological temperature was calculated separately for each 277 ensemble member (not the ensemble mean) to be consistent with reanalysis and observations. Next, for each day and grid point, the linear trends over 1979-2020 period from the ensemble mean 278 were removed. For reanalysis and observations, the mean of the two model's ensemble means were 279 removed. Nearly identical results were obtained using other methods, including removing the 280 281 linear trends from the ensemble-mean from each individual model, and removing the linear trend 282 from each individual ensemble member, or from observations, or from reanalyses. For the model data, nearly identical results were also obtained if the daily ensemble-means were removed. For 283 284 each year and season, the standard deviations of the daily temperature anomalies were calculated. 285 This results in 41-year time-series of the subseasonal temperature variability at each grid point and 286 each season.

#### 287 Fingerprint and signal to noise calculations

We use standard fingerprint methods<sup>16-18</sup> to determine whether the signal is detectable in 288 observations. First, we calculate the ensemble average of subseasonal temperature variability for 289 each of the all-forcing experiments. Next, we calculate anomalies of temperature variability by 290 291 subtracting the climatological mean at each grid-point. We then define the fingerprint to be the leading empirical orthogonal function (EOF) of the anomalies over all land grid points over the 292 30°-90°N latitude region for the 1979-2020 period. The results are not sensitive to the exact region 293 chosen. The signal time-series are calculated by projecting the observed anomalies in temperature 294 variability onto the fingerprint: 295

296 
$$S(t) = \sum_{x=1}^{N_x} O(x,t) \cdot F(x)$$

Where S(t) is the signal time series, O(x,t) is the observed anomalies, F(x) is fingerprint, t represents the year, x represents the grid-point and N<sub>x</sub> the number of grid-points. Fields are appropriately area-weighted prior to calculations by multiplying each grid-point by the square root of the cosine of the latitude.

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The noise time-series are calculated by first subtracting the appropriate ensemble-mean of the temperature variability time-series from individual ensemble members of the ensemble (including all single forcing experiments). Next, each of these time-series are then projected onto the fingerprint, as was done for observed anomalies in the equation above. This results in 200 41year noise time-series for CanESM2 and 95 41-year noise time-series for CESM1. These noise time-series consist only of unforced internal variability.

To calculate whether the observed trends are statistically significant, we calculate the signal-to-noise ratios. Here, the signal is defined as the magnitude of the linear trend of the observed signal time-series. The noise is defined as the standard deviation of the distribution of linear trend magnitudes of the noise time-series of the same length as the signal. Signal-to-noise ratios are calculated as a function of increasing trend length for trends starting in 1979, with a minimum trend length of 15 years. We define statistical significance to be a signal-to-noise ratio of 1.645, which is the 5% significance threshold using a one-sided Students t-test. The time of detection is determined by the first year where the signal-to-noise ratio exceeds and subsequently remains above this level.

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### **335 Data Availability**

- 336 CanESM2 data is available at: https://open.canada.ca/data/en/dataset/aa7b6823-fd1e-49ff-a6fb-
- 337 68076a4a477c. CESM1 data is available at:
- 338 https://www.earthsystemgrid.org/dataset/ucar.cgd.ccsm4.CESM\_CAM5\_BGC\_LE.html. ERA5
- reanalysis data is available at: https://cds.climate.copernicus.eu/cdsapp#!/dataset/reanalysis-era5-
- 340 single-levels?tab=overview. NCEP-DOE reanalysis 2 data is available at:
- 341 https://psl.noaa.gov/data/gridded/data.ncep.reanalysis2.html. Berkeley Earth observations are
- 342 available at: http://berkeleyearth.org/archive/data/.

# 343 Code availability

- Code is available form the corresponding author upon reasonable request.
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- 351 manuscript. J.C.F and J.A.S. discussed the results and made suggestions and edits to the
- 352 manuscript.
- **353 Competing Interests:** The authors declare no competing interests.
- 354 **Correspondence and request for materials** should be addressed to R.B.

# 356 Figures







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Figure 2: Signal time-series of the subseasonal temperature variability fingerprint. Signal time-series of the subseasonal temperature variability fingerprint in the ERA5 reanalysis (blue), NCEP reanalysis (red) and Berkeley Earth observations (orange). Time series are shown for autumn (SON; a, c), winter (DJF; b, d) and for the fingerprint calculated from the CanESM2 simulations (a, b) and CESM1 simulations (c, d).





Figure 3: Signal-to-noise ratios for increasing trend length. Signal-to-noise ratios as a function of trend length for trends starting in 1979/80, for the ERA5 reanalysis (blue), NCEP reanalysis (red) and Berkeley Earth observations (orange), and individual model realizations (grey). Signalto-noise ratios are shown for autumn (SON; a, c), winter (DJF; b, d) and for the fingerprint calculated from the CanESM2 simulations (a, b) and CESM1 simulations (c, d). The purple line indicates statistical significance at the 5% level. Note the different vertical axes in the different panels.



Figure 4: Drivers of subseasonal temperature variability trends. Trends in subseasonal nearsurface air temperature variability (°C/decade) in autumn (SON; a, c, e, g) and winter (DJF; b, d,
f, h) over the 1979-2020 period from the single-forcing experiments. Trends are shown for
anthropogenic greenhouse gas forcing (GHG; a-d), anthropogenic aerosol forcing (AER; e-h), in
CanESM2 simulations (a, b, e, f) and CESM1 simulations (c, d, g, h).

# 383 Extended Data Figures



Extended Data Figure 1: Subseasonal near-surface temperature variability trends in reanalysis and observations. Trends in subseasonal near-surface air temperature variability (°C/decade) in autumn (SON; a, c, e) and winter (DJF; b, d, f) over the 1979-2018 period. Trends are shown for ERA5 reanalysis (a, b) NCEP-DOE-reanalysis 2 (c, d), and Berkeley Earth observations (e, f). The stippling indicates trends that statistically significant at the 5% level using a two-sided student's t-test.

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393 Extended Data Figure 2: Subseasonal near-surface temperature variability trends in spring

and summer. As in Fig 1, but for spring (MAM) and summer (JJA).

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Extended Data Figure 3: Fingerprints of subseasonal temperature variability. The
fingerprints of subseasonal temperature variability from CanESM2 (a, b) and CESM1 (c, d) for
autumn (SON; a, c) and winter (DJF; b, d).



Extended Data Figure 4: Signal-to-noise ratios for increasing trend length. As in Fig 3, but
with grid-points over ocean included in the fingerprint.



## 407 Extended Data Figure 5: Subseasonal near-surface temperature variability trends. As in Fig

<sup>408 1,</sup> but with grid points over ocean included.



412 Extended Data Figure 6: Seasonal-mean near-surface temperature trends. As in Fig 4, but

<sup>413</sup> for seasonal-mean temperature trends (°C/decade).



417 Extended Data Figure 7: Meridional temperature gradient trends. As in Fig 4, but for

418 meridional temperature gradient trends (°C 1000 km<sup>-1</sup> decade<sup>-1</sup>).