1	Food for Thought
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3	South Africa's experimental fisheries closures and recovery of the
4	endangered African penguin
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23 Abstract. In a scientifically-transformative project, South Africa implemented a decade-long 24 field experiment to understand how fisheries may be affecting its most iconic seabird, the African penguin Spheniscus demersus. This unique effort prohibits the take of anchovy and 25 26 sardine within relatively small areas around four African penguin breeding colonies, two in the 27 Benguela upwelling ecosystem and two in the adjacent Agulhas region. For the Benguela, 28 fisheries closures within the birds' primary foraging range increased their breeding productivity 29 and perhaps reduced parental foraging efforts, indicating that the fisheries are competing with the 30 birds for food. Results were less clear for foraging behaviour in the Agulhas, but no data on 31 breeding success was collected there. The African penguin is endangered and its population continues to decline, and fisheries closures have been demonstrated improved demographic traits 32 33 that clearly contribute to population growth. Therefore, given the critical conservation status of 34 the species, fisheries closures should be maintained, at least at Dassen Island where the 35 population has great capacity to expand and support other nearby colonies. Continuing or 36 implementing corresponding fisheries closures in the Agulhas region, as well as creating and 37 testing the value of pelagic closed areas during the non-breeding season when the penguins 38 disperse widely across these ecosystems, also is warranted. These management actions would 39 increase penguin food supplies and help to meet societal goals of halting the decline of the 40 penguin population, as well as maintaining the economic and cultural services provided by 41 fisheries and ecotourism.

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43 <u>Keywords</u>: forage fish, anchovy, sardine, seabirds, conservation, societal conflicts, field
 44 experiment

46 <u>Main Text</u>

A long-term and growing issue in marine ecosystem conservation and management concerns the 47 simultaneous use in time and space of small pelagic fish and euphausiid crustaceans by industrial 48 49 fisheries and marine wildlife, which may cause local- to regional-scale competition, and reduce 50 prev availability to predators (Croll and Tershy, 1998; Smith et al., 2011; Pikitch et al., 2012, 51 2014; Sydeman et al., 2017, Gremillet et al., 2018). Fisheries' reductions in available prey can affect foraging behaviour and effort, rates of energy acquisition, and eventually population 52 53 parameters such as breeding success or recruitment. This problem is particularly severe for 54 seabirds during reproduction when they are constrained to forage relatively close to colonies by offspring provisioning duties (Furness, 1982; Duffy and Schneider, 1994; Tasker et al., 2000; 55 56 Sydeman et al., 2017; Free et al., 2021). The daily metabolic requirements of seabirds during 57 breeding when parents are feeding themselves and developing offspring are amongst the largest 58 (proportional to body mass) estimated for marine vertebrates.

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#### 60 African penguins and food supplies

61 The most significant ongoing example of this issue concerns the critically-imperiled African 62 penguin Spheniscus demersus, which has suffered population fragmentation and a >95% decline over the past century (Sherley et al., 2020). It is now classified as endangered by the 63 64 International Union for Conservation of Nature (Birdlife International, 2020) and is one of three 65 most vulnerable penguin species (Boersma et al., 2020). A combination of climate change and fisheries removals has reduced the availability of the penguins' primary food resources, 66 67 including sardines Sardinops sagax and anchovies Engraulis encrasicolus in Namibia and South 68 Africa (Coetzee et al., 2008; Roux et al., 2013; Crawford et al., 2015; Sherley et al., 2017;

Crawford *et al.*, 2019). This decline in prey abundance has reduced the survival rates of adult and juvenile penguins, most acutely since 2000 (Sherley *et al.*, 2014, 2017; Robinson *et al.*, 2015). Recent adult survival estimates for African penguin (Sherley *et al.*, 2014; Robinson *et al.*, 2015) are extremely low in comparison other seabirds in general (Hamer *et al.*, 2001; Appendix 2 in Schreiber and Burger, 2001; Table S1 of Bird *et al.*, 2020), and food stress-related changes in survival and recruitment is believed to be the primary mechanism underpinning the population collapse.

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#### 77 The fisheries experiment

In response to the steep decline of its most iconic seabird, government agencies, conservationists, 78 79 and academics with expertise in seabird ecology and fisheries science collaborated to implement 80 a carefully-planned and unprecedented ocean field experiment along the Western and Eastern 81 capes of South Africa. This experiment was designed to i) determine whether the fisheries were 82 affecting the penguins through food resource limitation, and ii) investigate if chick survival and 83 other proxies for penguin breeding productivity could be enhanced using fisheries time-area closures, a standard fisheries management technique (e.g., Dunn et al., 2011). Time-area 84 85 closures remove specific locations from fishing pressure during selected times. They may 86 thereby minimize potential fisheries competition with marine wildlife during periods of 87 biological significance, such as the chick-rearing period for seabirds when parents are 88 provisioning themselves as well as developing offspring, and are spatially-constrained to centralplace foraging where natural competition for prey is generally high (Burke and Montevecchi, 89 90 2009; Weber et al., 2021).

91 The experiment focused on four penguin colonies, two along South Africa's Western 92 Cape at Dassen and Robben islands, and two along the Eastern Cape at Bird and St. Croix 93 islands (Figure 1). Fisheries targeting anchovies and/or sardines were alternately closed within 94 the typical penguin foraging radius (20 km) at each island. For the Western Cape, fisheries at 95 Dassen Island were closed from 2008–2009 (inclusive) and again from 2014–2016; closures at 96 Robben Island were implemented from 2011–2013 and again from 2017–2019. This unusual 97 Before and After Control Impact (BACI) design, with replication and reference colonies, has 98 provided informative and compelling results.

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#### 100 Chick survival

101 Survival of penguin offspring was studied within the central portion of the highly productive, yet 102 extremely variable Benguela upwelling system (Fig. 1). On Dassen and Robben islands, the 103 African penguin once numbered over 300,000 breeding pairs, but today fewer than 3,500 remain, 104 which currently represents  $\sim 20\%$  of the world's population (Sherley *et al.*, 2020). While forage 105 fish abundance off the Western Cape varied substantially among the years of experimental time-106 area closures, survival of penguin chicks to the age of fledging, the key population parameter 107 under scrutiny, increased during years when fishing was excluded from the foraging areas of 108 breeding penguins (Sherley et al., 2018; Sherley, 2020, Supplementary Materials). While the 109 exact effect sizes ranged from  $\sim 2\%$  to  $\sim 11\%$  depending on model structure (Supplemental 110 Materials), population models suggest that this increase in chick survival is sufficient to improve 111 the penguin population growth rate by ~0.5–1.2% per annum (Sherley et al., 2018; OLSPS 112 Marine, 2021; Ross-Gillespie and Butterworth, 2021). While other interventions also are 113 required to increase the penguin's population growth rate, these analyses clearly demonstrate that

114 1) the forage fish fisheries operating in proximity to breeding colonies are competing with the 115 penguins for food and 2) permanent fisheries closures could offset  $\sim 20\%$  of the penguin 116 population decline which has averaged  $\sim 5\%$  per annum (Sherley *et al.*, 2020). Given the high 117 temporal and spatial variability in forage fish abundance (DFFE, 2021), and recent poor 118 accessibility of some prey, notably anchovies (e.g., Crawford et al., 2019), the finding of robust 119 local-scale changes in offspring survival related to fisheries closures is remarkable. This result 120 suggests that even quite small (<1100 km<sup>2</sup>) fisheries time-area closures near breeding colonies 121 are likely to positively affect breeding success in seabirds (Free *et al.*, 2021), and suggests that 122 larger permanent fisheries exclusion zones may confer even greater conservation benefits for 123 ecosystems and marine wildlife.

124 This critical result, however, has been the subject of heated debate, with no less than ten 125 recent non-peer reviewed reports prepared for the South African fisheries agencies arguing 126 technicalities of data analysis and modelling approaches (Supplementary Materials). Topics of 127 debate, for example, have included whether to use individual-level observations or aggregate 128 data to annual means (see Supplementary Materials for details). Nonetheless, despite detailed re-129 analyses by international panels reviewing experimental results and other interested scientists, 130 there have been no major changes in the original key findings of Sherley et al., (2018) showing 131 that chick survival and offspring condition are higher without fisheries in the vicinity of colonies. 132 These exchanges serve to highlight the critical importance of the penguin demographic data, 133 especially chick survival, which was not measured across all four colonies, and suggest these 134 parameters should be weighted most heavily when considering the effects of forage fish 135 removals in penguin foraging habitat. The implication is clearly that industrial fisheries 136 removing anchovy and sardine near breeding colonies compete with African penguins for food.

Arguably, this is the first time fisheries competition with seabirds for food has been adequately demonstrated through the use of a well-designed long-term field experiment (see preliminary comments in Sydeman *et al.*, 2017). The South African experiment thereby shows what is needed to assess the effects of forage fish fisheries at the local scale to seabirds breeding and foraging from colonies.

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# 143 Foraging behaviour

144 While experimental fisheries closures have provided noteworthy results on breeding success at 145 Robben and Dassen islands, the effects of fishing closures on penguin foraging behaviours, 146 designed to provide mechanistic explanations for changes in chick survival, were not definitive. 147 Foraging parameters measured include foraging trip duration and maximum foraging distance, 148 behaviours thought to respond to the local availability of food resources (Campbell *et al.*, 2019; 149 McInness *et al.*, 2019). The estimates of the impacts on foraging behaviour are generally much 150 more variable than those on chick survival or condition (Pichegru *et al.*, 2012; Ross-Gillespie 151 and Butterworth, 2021), particularly for colonies on the Eastern Cape where chick survival data 152 are unavailable for context (Bird and St. Croix islands; Fig. 1). The lack of robust results 153 regarding penguin foraging behaviour undoubtedly relates to the high variability in foraging 154 measurements among individuals that may be due to day-to-day changes in prey availability (cf. 155 Bertrand et al., 2012). More robust studies of individual movements and foraging behaviour 156 would likely require great effort and expense (Soanes et al., 2013). The rapid pace of the decline 157 of the African penguin (~5% per annum; Sherley et al., 2020) exacerbates these concerns. 158 Indeed, while the cost of devices to track seabirds has declined since the fisheries closures began 159 in 2008, the difficulty of finding enough birds to study of the right age and stage of the breeding

160 season has greatly increased. Most importantly to us, though, is the fact that measurements of 161 foraging behaviour remain difficult to link to changes in survival of chicks, immature 162 individuals, or adults, which are key elements in the population ecology of seabirds in general 163 and specifically to the recovery of the African penguin. Unfortunately, the logistical difficulties 164 of working on St. Croix and Bird islands mean that chick survival was measured only for the 165 colonies on the Western Cape, making full multivariate comparisons with those on the Eastern 166 Cape impossible.

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#### 168 Horizon scan

169 There are remarkably few BACI ocean experiments from which to draw comparisons about the efficacy of fisheries exclusions on fish availability to predators and wildlife population dynamics 170 171 (Sydeman et al., 2017). Following the widespread decline of Steller sea lions Eumetopias 172 *jubatus* in the Alaskan North Pacific, 20–40-km trawling exclusion zones were established along 173 the Aleutian Archipelago during the late 1990s. Regrettably, proposed replication of these time-174 area closures and control sites were not instituted as was done in South Africa, resulting in 175 ambiguous findings (McDermott et al., 2016). Additionally, in the early 2000s a quasi-176 experimental closure off the North Sea coast of the UK strongly suggested a negative impact of 177 the sandeel Ammodytes marinus fishery on black-legged kittiwake Rissa tridactyla breeding 178 productivity on the Isle of May, but not that of four diving species also breeding on the island 179 (Frederiksen et al., 2008). Elsewhere, data from Magellanic penguins Spheniscus magellanicus, 180 a species found in Argentina that is closely related to the African penguin, shows that penguins 181 that forage closer to their breeding colony have higher reproductive success, further supporting 182 the importance of protected areas around penguin breeding colonies (Boersma and Rebstock,

183 2009; Boersma et al., 2015). Looking forward, the Convention for the Conservation of Antarctic 184 Marine Living Resources (CCAMLR), of which South Africa is part, is considering the 185 implementation of time-area closures of krill fisheries in Antarctic Peninsula ecosystems to 186 protect penguin and seal breeding productivity (Trathan and Godø, 2018). Norway, South 187 Korea, Japan, Ukraine, and other nations have economic or conservation interests in Antarctic 188 krill fisheries, and the South African experiment is demonstrating the value of area-based 189 fisheries closures for marine predators. Finally, relative to this discussion, fisheries policies 190 designed to eliminate the potential for localized prey depletion and effects on marine predators 191 appear to work very well. In the Alaskan North Pacific, the U.S. North Pacific Fishery 192 Management Council prohibited fishing for most forage fish (e.g., Pacific sand lance Ammodytes 193 hexapterus and capelin Mallotus villosus) and euphausiid crustaceans in 1997 (Witherell et al., 194 2000). Hence, in this region, forage fish fisheries competition with marine wildlife has never 195 developed into a management issue, and its most important groundfish fisheries are considered 196 healthy as a result of this basic policy to protect mid trophic level food resources for the benefit 197 of all upper trophic level species.

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#### 199 Implications for penguin conservation

The African Penguin Biodiversity Management Plan guides conservation actions to facilitate penguin recovery in southern Africa. The original plan, which discussed a variety of ongoing threats to the penguin, called for provision of adequate prey resources in both the breeding and non-breeding seasons, with fisheries considered a key factor in determining year-round food resource availability. To date, the fisheries closure experiment has only addressed the potential effects on penguin breeding productivity. Given the key role of juvenile and adult survival to the

206 decline in the penguin population (Sherley et al., 2014, 2017; Robinson et al., 2015), data 207 collection could be broadened to investigate the importance of fishing during the non-breeding 208 season by implementing pelagic time-area closures. Moreover, South Africa could link its 209 approach with efforts to restore the marine ecosystem and fisheries of its Benguela Current 210 Commission partner, Namibia (Utne-Palm et al., 2010). In Namibia, the sardine fishery was 211 closed in 2018 after the stock dropped below the crash reference point of 50,000 t of spawning 212 biomass (Kainge et al., 2020; Erasmus et al., 2021) following over-fishing decades ago (Roux et 213 al., 2013). The Namibian Islands Marine Protected Area (NIMPA; Fig. 1) was implemented in 214 2009 (Ludynia et al., 2012) and is designed to promote both seabird and fisheries recovery. 215 However, because the forage fish community was so severely depleted, Namibian African 216 penguin populations collapsed and have yet to recover to any appreciable degree (Crawford, 217 2007; Sherley et al., 2020). That this ecological breakdown persists after four years of a 218 complete moratorium on sardine fishing in Namibia could be viewed as a worst-case scenario for 219 the situation in South Africa if the fisheries and penguin competition is not carefully managed. 220 The evidence suggests the importance of reducing competition with fisheries immediately to 221 stem the decline of the African penguin.

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To that end, we welcome the recent news that the Minister of Forestry, Fisheries and the Environment has established a Governance Forum which is, in turn, seeking to find a balance between minimizing the costs to the small-pelagics fishing industry while maximizing the protection of foraging areas for breeding African penguins. Through a series of meetings, this group has developed a proposal to close areas around three penguin colonies to purse-seine fishing: Dassen Island, Dyer Island, and St. Croix Island (Coetzee *et al.*, 2021). These three 229 colonies have all hosted populations of  $\sim 20,000$  pairs at some point since 1979. Dver Island 230 occupies an important strategic location between the formerly large colonies on the Western 231 Cape (Dassen and Robben islands) and the Eastern Cape colonies. St. Croix Island was, until 232 recently, the only colony still home to >3,000 breeding pairs, but numbers there have declined 233 sharply from  $\sim$ 3,600 pairs 2019 to  $\sim$ 1,500 pairs in 2021 (Sherley *et al.*, 2020, 2021). Dassen 234 Island was once the location of the largest African penguin colony in the meta-population and is 235 the nearest colony of >150 pairs to Namibia (Sherley *et al.*, 2020). Continued increasing 236 breeding productivity on Dassen Island, as has happened in the past decade with closures, could 237 lead to spill-over effects. For example, there may be increased dispersal and recruitment of 238 young penguins into Namibia as well as to Robben Island near Cape Town (Sherley et al., 2017). 239 The public-relations opportunity of increasing the Robben Island population is enormous. At 240 one colony in South Africa where the penguin population is relatively small but stable, Simon's 241 Town (also known as "Boulders"; Sherley et al., 2020; Fig. 1), penguin ecotourism is valued at >10 million USD/year (van Zyl, 2014; van Zyl and Kinghorn, 2018). The Robben Island 242 243 population, once numbering about 8,500 pairs (Sherley et al., 2020), is visited by approximately 244 325,000 people/year (2017-2019; https://nationalgovernment.co.za/entity annual/1936/2019-245 robben-island-museum-annual-report.pdf). Thus, Robben Island penguins allow South Africa to 246 engage and educate hundreds of thousands of people as they visit a recovering penguin colony. 247 This opportunity to herald the fisheries closures could enhance appreciation for marine wildlife 248 and demonstrate the commitment of South Africa to manage fisheries using a multi-species, 249 ecosystem-wide perspective.

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251 Conclusion

252 The scientific, conservation, and economic value of South Africa's long-term fisheries closures 253 with forage fish and penguins is considerable. The Benguela Current ecosystem is highly 254 dynamic and both fish and penguins are mobile. On one hand, we know that many seabird 255 species, including the African penguin, exhibit a non-linear threshold response in breeding 256 success to fluctuations in local forage stock abundance (Cury et al., 2011) and to the distance 257 they must travel to feed chicks (Boersma and Rebstock, 2015; Weber et al., 2021). On the other 258 hand, Free *et al.* (2021) suggest that mobile marine top predators and forage fish may not 259 respond to fisheries protected areas as well as more immobile organisms, though this depends 260 upon the persistence of forage patches within the birds foraging range (e.g., Davoren et al., 261 2003). There is recognition, however, that restricting fishing near predator breeding locations 262 would likely produce greater benefits (Free et al., 2021). This is also supported by key results 263 from the Western Cape, which show the importance of fisheries time-area closures to African penguins as centrally-foraging seabirds. Indeed, the closure experiment has demonstrated that 264 265 protecting primary foraging areas to safeguard predators' access to marine food resources 266 reduces competition with forage fish fisheries.

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The African penguin shares life history traits with other seabird species, and its fisheries case history resembles other ongoing and developing fisheries-wildlife conflicts both in South Africa and elsewhere in the world (e.g., Bertrand *et al.*, 2012; Frederiksen *et al.*, 2008; Watters *et al.*, 2020). Because most seabirds breed on land in conspicuous groups, their lifestyle provides opportunities to measure and evaluate important traits such as offspring production and survival of juveniles and adults, thereby providing clear results of experimental forage fish fisheries manipulations. South Africa fisheries management may be considered a model system for

275	resolving complex natural resources management issues in marine ecosystems globally. If
276	fisheries exclusions can be continued and perhaps increased in number and area, the recovery
277	and health of the African penguin throughout its range and the scientific basis for solutions to
278	minimize the impacts of forage fish fisheries on marine wildlife could be significantly enhanced.
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447

448

#### 449 **Data availability**

- 450 No data were used in the production of the main text of this article, however, there were data
- 451 utilized in the analyses described in the Supplementary Materials. Those data and their
- 452 associated code are available on GitHub (<u>https://github.com/rbsherley/IJMS\_AP\_IC</u>). Please see

453 the Supplementary material for more information.

# 455 Supplementary material

456 Supplementary material detailing "Best practices to analyse data relative to forage fish fisheries457 closures at penguin colonies" is available at ICESJMS online.

458

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and does not imply endorsement by the U.S. Government.

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African penguin colonies and fisheries closures adjacent to breeding islands, Cape Town, 472

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False Bay, and the Namibian Islands' Marine Protected Area (NIMPA).
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# **Supplementary Materials**

## Best Practices to Analyse Data Relative to Forage Fish Fisheries Closures at Penguin Colonies

# 5 On the use of aggregated versus disaggregated data

6 One of the primary concerns whether the penguin data should be analysed at the level of the 7 annual mean (e.g., Table S1), referred to as the "aggregated data" approach (e.g., Ross-Gillespie 8 and Butterworth, 2021a,b), or at the level of the original observations (i.e., each penguin nest, 9 foraging trip or chick), referred to as the disaggregated data approach (e.g., Sherley *et al.*, 2018). 10 In both cases, mixed models and random effects may be applied to account for the non-11 independence of samples made within each year. For the aggregated approach, 2 annual means 12 or estimates, one from each island, are produced per year with associated error terms. For the 13 disaggregated approach, 10s (adult foraging behaviour) to 100s (chick condition and survival) of 14 observations made within each island in a given year (e.g., 417 chicks monitored for survival at 15 Robben Island in 2008; Table S1) would be used. These approaches may result in differences in 16 apparent fisheries closure effects on the birds. 17 18 Critics of the disaggregated approach argue that it produces estimates of closure effects that are 19 too precise (i.e., negatively biased standard errors) relative to the aggregated approach 20 (Butterworth, 2020a), unless "island" is included in both the fixed and random components of 21 the models. Critics of the aggregated data argue that this approach will suffers from a risk of 22 over-parameterisation of coefficient estimates as the sample size would fall well short of the 10-23 15 data points per predictor variable recommended by simulation studies to obtain stable 24 parameter estimates (Babyak, 2004). Consequently, the concern is that the aggregated approach 25 lacks statistical power to detect all but very large effect sizes (Green, 1991). For example, 26 models using the aggregated approach to examine the effect of fisheries on chick survival (e.g., 27 model R3 in Table B2 of Ross-Gillespie and Butterworth, 2021a) typically use 22 data points 28 (see Table S1) to estimate 6 parameters (the intercept, a fixed island main effect, a fixed closure 29 main effect, a fixed interaction between island and the closure effect, the SE for year as a random 30 effect, and the residual SE). If models using aggregated data lack statistical power, there is a 31 greater risk of making a Type II error, but if models applied to the disaggregated data give overly 32 precise estimates, there is a risk of making a Type I error, i.e., if the random effect structure is 33 not carefully chosen (Arnqvist, 2020; Silk et al., 2020).

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Simulation studies have demonstrated that linear mixed-effects models (LMMs) can be used in
 these circumstances, when random effects are chosen based on the known sampling structure in

the data (Silk *et al.*, 2020), and particularly when model selection methods are used to choose the random effect structure (Matuschek *et al.*, 2017).

39

40 Matuschek *et al.* (2017) state: 41

42 *"Our simulations have shown that determining a parsimonious model with a standard model"* 

43 selection criterion is a defensible choice to find this middle ground between Type I error rate

- 44 *and power*".
- 45

This is (presumably) why a 2019 international review panel that assessed the various analysesfrom the fisheries closures said,

48

49 "Given the nature of the experiment, use of individual data is to be preferred. However, this is
50 only the case if an appropriate random effects structure is chosen. Model selection methods
51 should be applied to select an appropriate random effects structure" (Die et al., 2019).

52

In re-assessing the aggregated and disaggregated approaches again in 2020, an equivalent
 international panel made additional recommendations, noting that:

55

56 "Individual-based models are vulnerable to underestimating the standard errors of important

57 model outputs if the model is incorrectly specified; aggregated models are vulnerable to

58 assigning inappropriate weights by stratum. Both types of error can lead to biased estimates of 59 closure effects and the standard errors of the estimates of these effects" (Haddon et al., 2020).

60

61 For the aggregated approach, they recommended taking into account the variance associated

62 with, and number of data points used to calculate, the annual averages. For the disaggregated

approach they suggested including a model with a hierarchical random effects structure with

64 Island nested within Year (i.e., specified as Year + Year:Island, or Year/Island in R) in the

- 65 candidate model set.
- 66

67 Below is a summary of the latest analyses that have been carried out towards implementing the

2019 and 2020 panel recommendations, focusing on the chick condition and chick survival
 datasets at Robben and Dassen islands. This focus was chosen because these are metrics where

biologically magningful effects have been detected at the nonveltion level

<sup>70</sup> biologically meaningful effects have been detected at the population level.

71

# 72 Chick condition

Ross-Gillespie and Butterworth (2021a,b) implemented recommendations for the aggregated
 approach and compared outputs from these models to models using the disaggregated approach

75 with a Year/Island hierarchical random effects structure. These comparisons on like-for-like

76 datasets yield similarly precise estimates of the closure effect (see Figure 1 of Ross-Gillespie and 77 Butterworth 2021a). For example, Figure S1 shows the estimated effect of fishing (and

77 Butterworth 2021a). For example, Figure S1 shows the estimated effect of fishing (and

approximate 95% confidence intervals) on chick condition at Dassen and Robben islands from
 four models. Effect sizes have been converted to % change in the population growth rate using

the conversion factor of 0.108 from Appendix A of Ross-Gillespie and Butterworth (2021a). A

negative effect indicates a negative effect of fishing on the penguin population growth rate.

Model 1 (RGB A Y) and Model 2 (RGB D Y) of Figure S1 shows the effect sizes from what

Ross-Gillespie and Butterworth (2021a) call the "best" implementation of the aggregated

84 approach, and the disaggregated approach respectively (values are taken from Table 1 in Ross-

61 approach, and the abaggregated approach respectively (values are taken from racio r in resp. 85 Gillespie and Butterworth, 2021a). Note that the SE (printed above the x-axis) are nearly

- 86 identical.
- 87

88 Ross-Gillespie and Butterworth (2021a,b) did not implement the 2019 or 2020 panel

89 recommendations for the disaggregated approach, which would include "...best practices for

- 90 fitting mixed effects models (e.g., Zuur et al., 2009),... and that "model selection methods
- 91 should be applied to select an appropriate random effects structure" (Die et al., 2019). In

- 92 particular, the 2020 panel recommended the addition of Year/Island to the candidate set, not that
- 93 Year/Island necessarily reflects the best choice of the random effect. Leaving Month out of the
- 94 random effect ignores a source of hierarchical structure in the data that is known to be important,
- 95 and it is far from clear that Month can reasonably be nested in Island (Haddon et al., 2020).
- 96 Having Island in both the fixed and random components of the model also may lead to parameter
- 97 estimation issues (Sherley, 2020a), and is also not in keeping with best practices for fitting mixed 98 effects models; as Zuur et al. (2009) state,
- 99
- 100 "Because we have as many explanatory variables as possible in the fixed component, the random 101 component (hopefully) does not contain any information that we would like to have in the fixed 102 component".
- 103

104 We followed the 2019 and 2020 panel recommendations and applied information-criterion model

- 105 selection to objectively compare the relative support for a model with a Year/Island random
- 106 effect versus a Year/Month random effect (as proposed by Sherley et al., 2018). This approach
- 107 yields a  $\triangle$ AICc of 685.7 in favour of the model with the Year/Month random effect structure.
- 108 Model 3 and 4 are fit to the disaggregated data using the 'lmer' function in the *lme4* library for R
- 109 (Bates et al., 2015), with an island main effect, a closure main effect, and a closure×Island
- 110 interaction as the fixed effects. Although Model 3 does yield more precise estimates of the
- 111 closure effect than Model 4 (with the Year/Island random effect structure; compare Model 3.
- D YM and Model 4. D YI in Figure S1), it unequivocally provides a better fit to the data based 112
- 113 on the objective use of information criterion-based model selection. The code and data to
- 114 reproduce Figure S1 and to make the AICc-based model comparison are available on GitHub
- 115 (https://github.com/rbsherley/IJMS AP IC).
- 116



 $\begin{array}{c} 117\\118\end{array}$ 

Figure S1: Fisheries effects point estimates (circles) and approximate 95% confidence intervals (mean  $\pm$ 119  $\times$  2 standard errors (SE) for 1 and 2, and mean  $\pm \times$  1.96 SE for 3 and 4) for models examining the 120 impact of fishing with 20 km of Dassen Island (orange) and Robben Island (purple) on African penguin chick condition from 2008 to 2018. The effect sizes have been converted into the % change in the 121 122 population growth rate using the conversion factor of 0.108 from Appendix A of Ross-Gillespie and 123 Butterworth (2021a); a negative effect means that fishing being allowed around that island is predicted 124 to have a negative effect on population growth rate. The raw means calculated directly from the data are 125 shown as plus signs (+) for reference. The SE for each effect is printed above the x-axis below the

- 126 corresponding effect. RGB = point estimates and approximate 95% confidence intervals taken directly
- from Table 1 of Ross-Gillespie and Butterworth (2021a); A = uses the aggregated data approach; D =
- 128 uses the disaggregated data approach; Y = Year is used in the random effect structure; YM = 120
- Year/Month (month of sampling nested in the year of sampling) is used in the random effect structure.
   YI = Year/Island (island of sampling, nested in the year of sampling) is used in the random effect
- 131 structure. Model 3 gives a more parsimonious fit to the data than Model 4 ( $\Delta AICc = 685.7$ ).
- 132
- 133 Use of the model with the Year/Month random effect structure was labelled as "flawed science", 134 in a letter to the Editor in Chief of Proceedings of the Royal Society B (see Sherley *et al.*, 2018)
- in a letter to the Editor in Chief of Proceedings of the Royal Society B (see Sherley *et al.*, 2018)
   requesting that they retract the paper (see Coetzee, 2021 FISHERIES/2021/AUG/SWG-
- PEL/57)<sup>1</sup>. Yet, all modelling approaches shown in Figure S1 yield a fisheries effect on chick
- 137 condition at Robben Island that is statistically different from zero at the 5% level (Figure S1).
- This is important because an objective way to link chick condition to population demography has
- been developed and agreed upon by all parties (Sherley *et al.*, 2018; Ross-Gillespie and
- 140 Butterworth 2021b). However, some disagreement about the overall value of this effect (relative
- to the current rate of decline) remains (e.g., Sherley *et al.*, 2020, Butterworth, 2021), and only the
- 142 model using disaggregated data and the Year/Month random effect exceeds a previously agreed
- 143 threshold for a biologically meaningful effect at the population level (-1%) change in population
- 144 growth rate, shown by the dashed line in Figure S1).

# 145146 *Chick survival*

- 147 Criticisms of the disaggregated approach (e.g., Sherley et al., 2018; Sherley, 2020a) have again
- 148 focused on whether this model produces overly precise estimates (OLSPS Marine, 2021, pg. 4).
- 149 There are also questions about whether or not the closure effects were biased by a small
- 150 percentage of the chicks that were not monitored from hatching (see Sherley, 2020b), and
- 151 whether the interaction between island and the closure status should be retained even if the data
- and model selection show no support for an interaction (Bergh, 2020, pg. 4; Butterworth, 2020b,
- 153 pg. 7).
- 154
- 155 Ross-Gillespie and Butterworth (2021a) tried to implement the 2020 panel recommendations for
- 156 the aggregated data approach (Ross-Gillespie and Butterworth, 2021a, Appendix B), but had
- difficulty in estimating an island-dependent variance, so their results for the aggregated approach
- 158 (and those shown in Figure S2) do not explicitly take sample size into account. They indicated
- that this was unlikely to make much difference to inference from their approach as the sample
- 160 sizes were large (see Ross-Gillespie and Butterworth 2021a, pg. 2). Although direct comparisons
- 161 in Table B2 of their Appendix B show that the closure effect at Robben Island is some 3 times
- 162 larger when sample size is taken into account explicitly and would be much more similar to the
- effect they report from the disaggregated approach it would remain not significantly different
- 164 from zero at the 5% level. For the disaggregated approach, Ross-Gillespie and Butterworth
- 165 (2021a) did not attempt to implement the panel recommendation of using model selection
- 166 including a model with Year/Island in the random effect structure. Instead, they reproduce results
- 167 from OLSPS Marine (2021) from a fit to the disaggregated data using Laplace approximation<sup>2</sup>

<sup>&</sup>lt;sup>1</sup> The criticism and application to have the paper retracted are independent of a coding error in Sherley *et al.* (2018) and Sherley (2020a), which was noted by the 2020 panel (see Haddon *et al.*, 2020). That error has been corrected and did not affect the conclusions of that paper or the results presented here.

<sup>&</sup>lt;sup>2</sup> Somewhat ironically, OLSPS Marine (2021) report these as likely to return negatively biased standard errors relative to the Bayesian approach with Gibbs sampling also used there and in e.g. Sherley (2020). The extent to

- and just Year in the random component of the model, and then apply a correction factor of 1.12
- 169 to the confidence intervals in an attempt to convert from a maximum likelihood to a restricted
- maximum likelihood (REML) estimate (see Ross-Gillespie and Butterworth 2021a, pg. 2). Their
  "best implementation" of the aggregated approach yields a closure effect at Dassen Island that
- both exceeds the -1% threshold and differs from zero at the 95% level (Model 1 in Figure S2),
- while their "best implementation" of the aggregated approach does not. Additionally, the closure
- effect at Robben Island is not different from zero in either framework (see Figure 1 of Ross-
- 175 Gillespie and Butterworth 2021a). In contrast, the Bayesian approach (using Gibbs sampling) to
- 176 model the disaggregated data in OLSPS Marine (2021), which is reproduced here as Model 7 in
- 177 Figure S2, also indicates a closure effect at Dassen Island that is significantly different from zero
- 178 at the 5% level (and a point estimate of -0.94%).
- 179



180 181 Figure S2: Fisheries effects point estimates (circles) and approximate 95% confidence intervals (mean  $\pm$ 182  $\times$  1.96 SE) for models examining the impact of fishing within 20 km of Dassen Island (orange) and 183 Robben Island (purple) on African penguin chick survival from 2008 to 2018. Dark red points and error 184 bars are from models without island-specific estimates (models in a frequentist framework with additive 185 effects only). The effect sizes have been converted into the % change in the population growth rate 186 using the conversion factor of 0.124 from Appendix A of Ross-Gillespie and Butterworth (2021a); a 187 negative effect means that fishing being allowed around that island is predicted to have a negative effect 188 on population growth rate. The raw means calculated from the data in Table 1 are shown as plus signs 189 (+) for reference. A = model uses the aggregated data approach (Models 1 to 3) and are implemented in 190 R using REML estimation. Values printed above the x-axis for Models 1 to 3 are AICc-values (smaller 191 AICc values = better relative support for a model). Model 1 reproduces the "best implementation of 192 the aggregated approach" from Table 1 and Figure 1 of Ross-Gillespie and Butterworth (2021a). 193 SHE = model uses the disaggregated data approach and are based on updates to models in Sherley 194 (2020a) implemented in a Bayesian framework using MCMC sampling in JAGS. Values printed above 195 the x-axis for Models 4 to 6 are PSIS–LOO-values (smaller PSIS–LOO values = better relative support 196 for a model). Model 7 directly plots the values from the Bayesian approach using Gibbs sampling in 197 OLSPS Marine (2021). For all models,  $I^*C$  = model contains an Island × Closure interaction in the 198 fixed effects component; I+C = model contains additive Island and Closure effects; C = model contains which the "REML-like" correction applied by Ross-Gillespie and Butterworth (2021a) is appropriate is entirely unclear, but this approach yields a closure effect at Dassen Island that does not differ from zero at the 95% level.

- 199only a main effect of closure in the fixed effects component; Y = Year is used in the random effect200structure; YN = Year/NestID (ID of the nest in which chicks were sampled nested in the year of201sampling) is used in the random effect structure; YI = Year/Island (island of sampling, nested in the202year of sampling) is used in the random effect structure. AICc values and PSIS-LOO values are not203directly comparable to one another.
- 204

205 Figure S2 also shows results for a model fit that implements the random effect structure 206 originally proposed for models using the disaggregated data on chick survival – nest identity 207 (Nest ID) nested within the year of sampling (e.g., specified as Year/NestID in R). This random 208 effect structure was originally selected because chicks in the same nest are provisioned by the 209 same parents, so their probabilities of survival are clearly non-independent of one another 210 (Sherley et al., 2013, 2018). As with Month in the chick condition models above, leaving NestID out of the random effect ignores a source of hierarchical structing in the data that is known to be 211 important (Sherley et al., 2013). This model (Model 4 in Figure S2), fit using a Bayesian 212 213 framework based on Markov Chain Monte Carlo (MCMC) estimation in JAGS (v. 4.3.0; 214 Plummer 2003) is an update of M3 in Table 3 of Sherley (2020a) with the coding error therein 215 corrected (see footnote 1). It implements the Year/NestID random effect and shows closure 216 effects at both Dassen Island and Robben Island that are credibly different from zero at the 5%

- 217 level and both very close to the -1% threshold (-0.95% and -0.98% respectively).
- 218

219 Although it has been claimed that this model structure - with Year/NestID in the random effect -220 should produce estimates that are too precise relative to the aggregated approach and a model 221 with Year/Island in the random effect (see above), Figure S2 shows that this is not necessarily 222 the case. Model 6 in Figure S2 is identical to Model 4, except that it uses Year/Island in the 223 random component of the model rather than Year/NestID. The precisions of the closure effects 224 are very similar between the two models (SE are ~0.016 and ~0.018 respectively). Moreover, 225 objective model selection indicates that Model 4 (Year/NestID) provides a better fit to the data 226 (PSIS-LOO = 16561.8) than Model 6 (Year/Island, PSIS-LOO = 17146.9, Figure S2). In this 227 case, model selection was carried out using Pareto smoothed importance sampling, leave-one-out cross-validation (PSIS-LOO; Vehtari et al., 2019); as with AICc, the lower the value the more 228

- 229 relative support there is for a model.
- 230

Finally, the results in Figure S2 strongly suggest that the differences in the precision of the

- estimates between the approach favoured by Ross-Gillespie and Butterworth (2021a,b; Model 1
- in Figure S2) and Sherley and colleagues (Sherley *et al.*, 2018; Sherley, 2020a; Model 4 in
- Figure S2) are not due to the choice of the random effects structure but actually whether or not
- the interaction effect is retained in the fixed effect structure. Model 1 (closure main effect SE =
- 0.044) and Model 2 (closure main effect SE = 0.018) in Figure S2 are both applied to the
- 237 aggregated data and are identical apart from whether or not the island × closure interaction is
- 238 retained in the model respectively. The  $\triangle$ AICc between the 2 models is 2.5 (in favour of the
- 239 model without the interaction, Figure S2), which suggests no strong evidence to retain it. Model 240 2 is simplified further by draming the Jaland main effect and here AAICa = 5 (7 laws then
- 240 3 is simplified further by dropping the Island main effect and has  $\Delta AICc = 5.67$  lower than 241 Model 1. Moreover, Model 5 (closure main effect SE = 0.039) and Model 4 (closure main effect
- SE = 0.016) are both applied to the disaggregated data and are identical apart from whether or
- 243 not the island × closure interaction is retained in the model respectively. As with the aggregated
- 244 approach, the model without the interaction (Model 4) provides the better relative fit to the data
- based on objective PSIS-LOO model selection. Some subjective arguments for necessarily

- retaining the island × closure interaction have been offered. For example, Bergh (2020) says
- 247 "Figure 5 of FISHERIES/2020/JUL/SWG-PEL/53REV provides estimates under the constraint
- that the closure effect is equal across the two islands (at least in its Figure 5). There seems no
- 249 reason for these estimates to be equal. The separate estimates for the closure effect at these
- 250 islands should be presented" (see also Butterworth, 2020b, pg. 7). To our knowledge, however,
- 251 no objective reason why model selection should be used to select the random component of the
- 252 models (as per the 2019 and 2020 panel recommendations) <u>but not</u> the fixed component has yet
- been put forward. Thus, at the time of writing, Model 4, Figure S2, would appear to be the best
- 254 implementation of the 2019 and 2020 panel recommendations for assessing the impact of the 20
- 255 km fishing closures on African penguin chick survival.
- 256
- 257 Detailed methods on the Bayesian model implementation and model selection are provided on
- 258 page 3 and 15 of Sherley (2020a) and the data and code to run the Bayesian models referenced in
- 259 Figure S2 and to reproduce Figure S2 are available on GitHub
- 260 (<u>https://github.com/rbsherley/IJMS\_AP\_IC</u>).
- 261

Table S1: Annual estimates of the mean survival rates for African penguin chicks monitored at Dassen

- Island and Robben Island between 2008 and 2018 used by OLSPS Marine (2021) and Ross-Gillespie and
- Butterworth (2021) in the aggregated data approach.  $\Phi_{74}$  is the proportion of chicks surviving to 74 days
- post hatching and is based on a lognormal survival function. N = the number of chicks monitored in each
- year at each island. The methods used to derive these annual estimates are given in Appendix D of Ross-Gillespie and Butterworth (2021).

267 268

Year	Island	Closure	$\Phi_{74}$	Ν	Island	Closure	$\Phi_{74}$	Ν
2008	Dassen	Closed	0.7718	246	Robben	Open	0.6395	417
2009	Dassen	Closed	0.6062	229	Robben	Open	0.5479	252
2010	Dassen	Open	0.5314	121	Robben	Open	0.6119	227
2011	Dassen	Open	0.5394	193	Robben	Closed	0.6376	289
2012	Dassen	Open	0.7481	124	Robben	Closed	0.7092	237
2013	Dassen	Open	0.7609	160	Robben	Closed	0.7977	234
2014	Dassen	Closed	0.7677	133	Robben	Open	0.7628	188
2015	Dassen	Closed	0.7864	107	Robben	Open	0.7595	298
2016	Dassen	Closed	0.8361	119	Robben	Open	0.7200	246
2017	Dassen	Open	0.7888	156	Robben	Closed	0.8322	212
2018	Dassen	Open	0.7539	54	Robben	Closed	0.7811	262

271	Supplementary References:
272	Arnqvist, G. 2020. Mixed models offer no freedom from degrees of freedom. Trends in Ecology
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