

1 **Interacting with the enemy: indirect effects of personality on conspecific aggression**  
2 **in crickets**

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22 **TITLE**

23 Interacting with the enemy: indirect effects of personality on conspecific aggression in  
24 crickets

25 **RUNNING TITLE**

26 Indirect effects of personality on aggression

27 **ABSTRACT**

28 In animal contests, individuals respond plastically to the phenotypes of the opponents  
29 that they confront. These ‘opponent’ – or ‘indirect’ – effects are often repeatable, e.g.,  
30 certain opponents consistently elicit more or less aggressiveness in others.  
31 ‘Personality’ (repeatable among-individual variance in behavior) has been proposed as  
32 an important source of indirect effects. Here, we repeatedly assayed aggressiveness of  
33 wild-caught adult male field crickets *Gryllus campestris* in staged dyadic fights,  
34 measuring aggressiveness of both contestants. Measurements of their personality in  
35 non-social contexts (activity and exploration behavior) enabled us to ask whether  
36 personality caused indirect effects on aggressiveness. Activity, exploration, and  
37 aggressiveness were positively associated into a behavioral syndrome eliciting  
38 aggressiveness in conspecifics, providing direct evidence for the role of personality in  
39 causing indirect effects. Our findings imply that a multivariate view of phenotypes that  
40 includes indirect effects greatly improves our ability to understand the ecology and  
41 evolution of behavior.

42 **KEY WORDS:** animal personality, behavioral syndrome, aggression, indirect effects,  
43 social behavior, crickets

44 **INTRODUCTION**

45 Over the past decades, research has increasingly focused on the adaptive nature of  
46 repeatable among-individual variation in behavior, called ‘animal personality’ in the  
47 behavioral ecology literature (Dall et al. 2004; Sih et al. 2004a,b; Réale et al. 2007,  
48 2010). This field of evolutionary biology thereby increasingly embraces the notion that  
49 natural selection can favor behavioral variation at multiple hierarchical levels, such as  
50 within *and* among individuals (Dingemanse et al. 2010b; Araya-Ajoy and Dingemanse  
51 2014). While the existence of animal personality is increasingly viewed as representing  
52 a ubiquitous feature of natural populations (Bell et al. 2009; Dingemanse et al. 2012b;  
53 Garamszegi et al. 2012), recent focus is shifting from questions regarding adaptive  
54 causation towards those concerned with the consequences of personality variation for  
55 ecological and evolutionary processes (Dall et al. 2012; Wolf and Weissing 2012;  
56 Dingemanse and Araya-Ajoy 2015).

57 Social interactions have recently been implicated as an important factor  
58 shaping behavioral variation at both the within- and among-individual level  
59 (Bergmüller and Taborsky 2010; Montiglio et al. 2013; Dingemanse and Araya-Ajoy  
60 2015; Niemelä and Santostefano 2015). Variation within individuals may partly occur  
61 because individuals modify their behavior as a function of the phenotypes expressed  
62 by conspecifics (so-called ‘indirect effects’; see below). Individuals may alter a  
63 particular behavior either in response to the same trait or a different trait expressed by  
64 conspecifics. An example of the former effect is provided by reciprocal aggression in

65 escalating fights, while an example of the latter is provided by aggression in response  
66 to opponent body size (Moore et al. 1997).

67 Indirect effects can have major consequences for ecological and evolutionary  
68 processes especially when caused by repeatable differences in phenotypes of social  
69 partners that are underpinned by genetic variation. In such cases, the genes of an  
70 individual influence the (behavioral) expression of a trait expressed by an interacting  
71 individual; such effects are termed 'indirect genetic effects' (IGEs) (Wolf et al. 1998,  
72 1999; McGlothlin et al. 2010). IGEs can greatly influence evolutionary processes  
73 particularly when they are correlated with direct genetic effects (DGEs) of an  
74 individual's genes on its own phenotype. In mice, for example, aggressiveness is  
75 heritable (a DGE), aggressiveness elicited in conspecifics is heritable too (an IGE), and  
76 aggressive genotypes elicit greater levels of aggressiveness in conspecifics. Such  
77 positive genetic correlations between DGEs and IGEs are predicted to speed up the  
78 response to directional selection compared to predictions derived from classic  
79 quantitative genetics theory (e.g., McGlothlin et al. 2010). By contrast, negative  
80 genetic correlations between DGEs and IGEs might instead impose evolutionary  
81 constraints. In gulls, for example, genes expressed in females that contribute to early  
82 egg laying (a DGE) actually delay the egg laying date of other females when expressed  
83 in their male partners (an IGE) (Brommer and Rattiste 2008). Obviously, phenotypes  
84 that are relatively repeatable (e.g., body size) and affect fitness in interaction partners  
85 are most likely to induce indirect effects in nature.

86           The consequences of indirect (genetic) effects for evolutionary processes are  
87 even more profound when we fully acknowledge the multivariate nature of most  
88 phenotypes (Moore et al. 1997; McGlothlin and Brodie 2009; Araya-Ajoy and  
89 Dingemanse 2014; Bijma 2014). Behavioral traits are often structured into syndromes  
90 (Garamszegi et al. 2012), i.e., whole suites of behaviors are correlated at the among-  
91 individual level; such syndromes may impose constraints on micro-evolution of  
92 behavior (Dochtermann and Dingemanse 2013). This would, for example, occur when  
93 selection favors multivariate behavioral phenotypes that are not present in the  
94 population due to strong syndrome structure. To date, studies of syndromes have  
95 considered only correlations among ‘direct’ sources of behavioral variation. However,  
96 it is likely that indirect effects are also integrated as parts of syndromes. This would  
97 occur if, for instance, opponent personality is itself an important source of indirect  
98 effects on focal behavior. Indirect effects are expected for behaviors such as  
99 aggressiveness and dominance (Moore et al. 1997, 2002), which are themselves often  
100 correlated with other behaviors such as boldness, exploration and activity (Garamszegi  
101 et al. 2012). Social partner effects should therefore be incorporated as part of  
102 behavioral syndromes to better understand the evolution of correlated traits.

103           Aggressiveness represents a good example of an ‘interacting phenotype’  
104 studied both in quantitative genetics and behavioral ecology (Wilson et al. 2009; Briffa  
105 et al. 2015). Aggressiveness is interesting because of its reciprocal nature and  
106 documented occurrence of opponent (also called ‘social partner’) effects (Wilson et al.  
107 2009, 2011, 2013; Camerlink et al. 2012, 2013; Alemu et al. 2014). Interestingly,  
108 contest theory has traditionally focused on the role of morphological traits such as

109 weapons and body size in shaping the intensity of aggressive interactions (Arnott and  
110 Elwood 2008). In empirical studies, however, morphological traits often explain only a  
111 moderate portion of the variation in contest behavior (Briffa et al. 2015). The  
112 behavioral stress physiology literature implies that behavioral traits may also play an  
113 important role because 'proactive' (active, aggressive, explorative, bold) individuals are  
114 less responsive to intentions signaled by interaction partners compared to 'reactive'  
115 (less active, aggressive, explorative, bold) individuals (Koolhaas et al. 1999; Coppens et  
116 al. 2010). While effects of personality on social responsiveness are predicted by game  
117 theory (Maynard-Smith 1982; Dall et al. 2004), the idea has received little empirical  
118 testing to date (Briffa et al. 2015; Dingemanse and Araya-Ajoy 2015).

119         Crickets are an ideal system to address the occurrence of indirect effects of  
120 personality on the expression of aggressiveness. Males of the European field cricket  
121 (*Gryllus campestris*) are territorial and frequently engage in agonistic contests to  
122 monopolize resources such as burrows or mates (Ritz and Köhler 2007). Wild European  
123 field cricket populations show repeatable among-individual differences in suites of  
124 traits such as activity, aggressiveness, boldness, and exploratory tendency (Rodríguez-  
125 Muñoz et al. 2010; Fisher et al. 2015a,b; Niemelä et al. 2015).

126         Specific study designs are required to quantify variation in individual-level traits  
127 where both direct and indirect effects are targeted (Moore et al. 1997; McGlothlin and  
128 Brodie 2009; Bijma 2014). We focus here on a design where the same individual has  
129 social interactions both as a 'focal' and as an 'opponent', and where each individual  
130 plays each role repeatedly across behavioral trials (Wilson et al. 2009, 2013;

131 Dingemanse and Araya-Ajoy 2015). Such a setup enables the partitioning of the total  
132 phenotypic variance in behavior expressed in focal individuals to variance attributable  
133 to (i) the focal individual's identity (a direct effect), (ii) the opponent's identity (an  
134 indirect effect), and (iii) residual within-individual variance (see Wilson et al. 2009,  
135 2011, 2013 for worked examples). Moreover, because each individual is both used as a  
136 focal individual and as an opponent, we can additionally ask whether individuals that  
137 (on average) have a high value for a particular behavior also elicit this behavior in  
138 conspecifics. Statistically this is measured as the correlation between direct and  
139 indirect effects (i.e., the phenotypic components of the genetic correlation between  
140 DGEs and IGEs detailed above). Application of this design has, for example,  
141 demonstrated that mice with aggressive personalities also elicit aggressiveness in their  
142 interaction partners (Wilson et al. 2009). When other traits are additionally measured,  
143 their effects may readily be incorporated to ask which (combination of) traits exactly  
144 induce social responsiveness in interaction partners (McGlothlin and Brodie 2009;  
145 Bijma 2014).

146         Here, we applied this paradigm to the study of aggressiveness expressed in  
147 pairwise interactions. We asked whether the personality of an individual (measured by  
148 its average level of activity, exploratory behavior, and aggressiveness across repeated  
149 observations) affects the aggressiveness expressed in conspecifics, thus integrating the  
150 indirect effects in a multivariate view of the phenotype. We repeatedly assayed  
151 activity, exploratory tendency, and aggression to determine the presence of  
152 repeatable among-individual variation in these traits ('personality'). To address their  
153 indirect effects, we applied a design where each male cricket fought repeatedly in



154 dyadic interactions, equally often as a focal and as an opponent. We then quantified (i)  
155 whether activity, exploratory behavior, and aggressiveness harbored repeatable  
156 among-individual variation, (ii) whether individuals were also repeatable in the level of  
157 aggressiveness elicited in opponents, (iii) whether an individual's average level of  
158 aggressiveness predicted aggressiveness elicited in opponents or (iv) whether other  
159 key individual-level characteristics (activity, exploratory tendency, body weight) did so  
160 instead. Our second objective was to integrate indirect effects into our description of  
161 behavioral syndrome structure. We therefore tested the relative fit of nine alternative  
162 hypotheses (models) concerning syndrome structure based on literature, with a  
163 structural equation modelling approach (Dochtermann and Jenkins 2007; Dingemanse  
164 et al. 2010a); this allowed us to test explicit predictions on the correlation structure  
165 among behaviors (detailed in Figure 1). Doing so enabled us to determine the nature of  
166 associations between behavioral traits and indirect effects in contest behavior, and  
167 thereby forcefully address the consequences of variation in 'personality' during social  
168 interactions.

## 169 **METHODS**

### 170 **Cricket collection**

171 Crickets were collected from a meadow adjacent to the Max Planck Institute for  
172 Ornithology (Seewiesen), Germany (47°58'35.5"N 11°14'04.5"E), between the 2<sup>nd</sup> week  
173 of May and the 1<sup>st</sup> week of June 2013. The main collection site was a south-west facing  
174 slope within the meadow. From the 4<sup>th</sup> week of April onwards, we searched the field

175 daily for burrows of newly emerged individuals and marked each with a numbered  
176 flag. Our aim was to trap individuals close to adulthood because our interest was in  
177 quantifying adult behavior (detailed below). We attempted to catch crickets in this  
178 stage with funnel traps (detailed in Niemelä et al. 2015) set at the burrow entrance  
179 between 10h00-17h00, for up to 30 minutes and up to 4 times per day per burrow,  
180 until the individual was caught. We succeeded in catching most crickets within the first  
181 two days of attempted catching. Because we monitored burrows on a daily basis, we  
182 could track each individual's developmental stage; all captured adults were no more  
183 than 1 day post-molting, implying that they lacked experience with mating and  
184 fighting. Following capture, crickets were placed in individual containers and  
185 transported to a climate room at the Ludwig Maximilians University of Munich. We  
186 collected a total of 57 males of which we randomly selected 32 at the onset of the  
187 experiments.

## 188 **Housing**

189 All individuals were housed in a climate room (dimensions: 4.5 l x 3.6 w x 2.7 h m) at  
190 26°C ( $\pm 0.5$ ) and 65% ( $\pm 0.5$ ) humidity, under a light:dark photoperiod that wild crickets  
191 experienced at the time of capture (14:10 h). Each individual was housed alone in a  
192 plastic container (10 x 10 x 9 cm) with a sand-covered floor and a flow-through plastic  
193 netted lid that prevented escape but allowed air circulation. Each container included  
194 an artificial, half-cylindrical shelter (6 x 3.5 x 2 cm), a petri dish (3.5 cm diameter) with  
195 food, and another petri dish with water held within a cotton-plugged vial. Individuals

196 were fed with a mix of dry bird food (Aleckwa Delikat, Germany) and fresh food (carrot  
197 and apple) *ad libitum*. Food and water were replaced every 3-4 days. Containers were  
198 checked daily for newly molted adults, enabling us to assign age since final molting for  
199 each individual not yet molted.

## 200 **Experimental protocol**

201 Behavioral trials were conducted between the 26<sup>th</sup> of May and 24<sup>th</sup> of June 2013. Each  
202 individual was repeatedly assayed for each of 3 behaviors (activity, exploration,  
203 aggression; described in detail below), on each of 12 consecutive days. Because  
204 individual identification is required for the aggression test, subjects were marked with  
205 colored tape on the pronotum (red or blue, randomly assigned each time) the day  
206 before a trial. The three tests were always done sequentially and in the same order  
207 (table 1, figure 2); carry-over effects (from one test on the next) could therefore not be  
208 modelled. We chose this set-up because it ensured that all individuals were given the  
209 exact same treatment since this greatly facilitates comparison between individuals  
210 (Dingemanse et al. 2007; Dochtermann 2010).

211 We initially selected 32 individuals randomly from the 57 collected individuals.  
212 These were divided into eight groups of four individuals according to their estimated  
213 age (i.e., days post-molting) to ensure that all the individuals of the same group were  
214 approximately the same age. It has been shown that adult crickets within one week  
215 post-molting do not show fully developed aggressive behavior (Hofmann and  
216 Schildberger 2001); the behavioral assays were therefore conducted when individuals

217 were at least 7 days post-molting. Groups were on average 9.8 days (standard  
218 deviation (SD): 2.6) post-molting at the onset of their behavioral trials. Six individuals  
219 died before the end of the experiment and were replaced, increasing the total number  
220 of tested individuals to 38. For data analysis (see below), we used all data, including  
221 behavioral trials of individuals that died during the experiment and of their  
222 replacements.

223         Within each group of four, each individual was subjected to an aggression test  
224 once on each of 12 consecutive days. From previous work, we know that this inter-test  
225 interval prevents carry-over effects caused by winner or loser effects (Khazraïe and  
226 Campan 1999). To maximize the number of unique pairwise dyads, and sample size per  
227 dyad, each individual fought four times with each of the three interaction partners  
228 within its group. All individuals within a group of four were tested on the same day,  
229 and each of two groups were assayed at the same time (i.e., 8 individuals  
230 simultaneously, one group per shelf). To minimize within-group variation in micro-  
231 environmental effects, each group was assayed at a fixed time of day over the 12  
232 consecutive days (either between 9h00-12h00, 12h00-15h00 or 15h00-18h00). After  
233 each group was tested, walls and dividers of the testing arenas were thoroughly  
234 washed with warm water; sand was raked and mixed in a stock container in order to  
235 minimize the presence of contact pheromones that could otherwise affect future  
236 contests (Judge and Bonanno 2008).

237         All trials were performed on a rack fitted with two shelves, each equipped with  
238 a camera, in the same climate room where the individuals were housed. Arenas were

239 not acoustically isolated from the rest of the room such that all the individuals were  
240 exposed to abundant background noise typical of conditions experienced in the wild.  
241 To prevent any visual disturbances, the shelves were isolated from the climate room  
242 by white curtains; entrance to the room was not permitted during trials. All trials were  
243 recorded using high-resolution digital video cameras (Basler GenICam, Germany) fitted  
244 43 cm above each testing arena. The cameras were connected to a computer outside  
245 of the climate room and managed using the software MediaRecorder (Noldus,  
246 Netherlands). Videos were recorded at 27.81 frames per second and 1600 x 1200  
247 pixels resolution.

248           A small number of trials were excluded from the final dataset: 32 of 382 activity  
249 and 29 of 382 exploration trials (8.4% and 7.6% respectively) due to technical problems  
250 with data recording, and 24 of 191 aggression trials (12.6 %) because males failed to  
251 engage in aggressive interactions. Note that the total number of aggression trials is  
252 approximately half of that of other trials since two individuals are involved in each  
253 aggression test. The final sample size was therefore 350 tests for activity (mean  
254 number per individual: 9.2, SD 3.7), 353 for exploration (mean number per individual:  
255 9.3, SD 3.4) and 167 for aggression (mean number per individual: 8.8, SD 3.8).

## 256 **Behavioral trials and scoring**

257 *Activity:* The general level of activity in a familiar environment was recorded in the  
258 individual's home container. Those home containers were large enough to allow free  
259 movement reflecting its routine activity. Prior to the onset of the behavioral trials,

260 selected individuals were carefully moved to the recording shelves inside their home  
261 container and given 30 minutes to acclimatize. To optimize the automated video  
262 tracking, the lid, water vial, food and petri dishes were removed from the home  
263 containers. Activity was recorded automatically for 60 minutes (Table 1; Figure 2a).

264 *Exploration:* Following the activity test, we gently moved the subject (inside its home  
265 container) to the side of the shelf where it had previously been filmed, after which we  
266 placed two arenas under each camera. These arenas (29.5 l x 15.5 w x 9 h cm) had  
267 floors covered with white sand and an opaque removable divider in the middle. When  
268 the divider was lowered, the two halves (14.5 l x 15.5 w x 9 h cm) of a single arena  
269 could be used as arenas to simultaneously screen, in isolation, exploratory behavior of  
270 each of the two individuals that would later meet in an aggression test. At the onset of  
271 the exploration test, each individual was moved (inside its own shelter) from its home  
272 container to a (randomly allocated) half of the arena. We ensured that all individuals  
273 were inside their shelter at the onset of the exploration trial, after which exploratory  
274 activity was recorded automatically for 60 minutes (Table 1; Figure 2b).

275 Activity and exploration videos were analyzed using Ethovision X 10.0 (Noldus, the  
276 Netherlands). This software package enables tracking of isolated individuals and  
277 extracts the spatial coordinates for each video frame. For both the activity and  
278 exploration tests, we used the Pythagorean equation to calculate the distance moved  
279 from one frame to the next, using the X and Y coordinates of the individual's center  
280 point at each sampling point (User manual of Ethovision X 10.0, Noldus Information  
281 Technology 2013). The frame rate at which the videos were analyzed was adequately

282 set to 6.95 per sec, and we therefore acquired 25020 positions per individual for these  
283 60-min trials. We then summed up all distances to calculate the total distance moved,  
284 in the familiar environment (activity test) viewed as a measure of 'activity'. The total  
285 distance moved in the novel environment (exploration test) is viewed as a measure of  
286 'exploratory behavior' (following Réale et al. 2007).

287 *Aggression:* Following the exploration test, the shelters were removed and the  
288 individuals given a further 10 minutes to acclimatize to their half of the arena. The  
289 divider was then lifted, after which we filmed each dyad engaging in social interactions  
290 for a period of 10 minutes (Table 1, Figure 2c). At the end of each aggression test, each  
291 individual was weighed on an electronic balance (precision: 0,001 g). We then changed  
292 each individual's color tag according to the next day's schedule, and returned it to the  
293 home container in the allotted housing slots within the climate room.

294 We assayed three behaviors that represented an individual's willingness to  
295 engage in aggressive interactions: 'approaching', 'singing', and 'chasing' (defined  
296 below). An interaction was defined as starting when any part of one individual touched  
297 any part of the other (Bertram et al. 2011), and ended if the contact was interrupted  
298 for more than two seconds. Within the 10-minute trial, for each interaction, we scored  
299 for each member of the dyad whether it did or did not perform each of these three  
300 behaviors. We subsequently counted how often a behavior occurred per individual  
301 over the entire trial, resulting in three count variables for each combination of  
302 individual and trial. The total number of interactions observed over all 167 trials was  
303 2589 (mean number of interactions per trial = 15.05, min = 1, max = 52). To acquire the

304 behavioral data, each video was played in slow motion and scored blindly by one of  
305 three observers.

306 Approaching: We scored an individual as 'approaching' during an interaction when it  
307 moved towards the other individual from any angle until they came into contact.  
308 When only one individual was actively approaching the other (i.e. the other cricket sat  
309 still), we assigned the behavior to that individual alone. In cases where both  
310 contestants approached each other at the same time, we assigned the behavior to  
311 both. The mean number of approaches per individual per trial was 9.64 (min = 0, max =  
312 48). Approach behavior has been used by several studies quantifying aggressiveness in  
313 male crickets, sometimes called 'initiating first contact' or 'initiating first aggression'  
314 (Brown et al. 2006, 2007).

315 Singing: Male crickets produce songs by rubbing their wings together when they  
316 encounter rivals and this has been reported as an indication of male aggressive intent  
317 (Alexander 1961). Singing involves wing movements and is visually easily distinguished  
318 from courtship songs and victory songs, the latter produced only by the winner *after*  
319 an interaction. The mean number of times singing recorded per individual per trial was  
320 1.73 (min = 0, max = 19).

321 Chasing: A chase was recorded when one of the contestants (namely, the winner)  
322 actively pursued its opponent immediately after the interaction had ended. We  
323 defined chases as pursuits in the direction of the loser initiated within 2 seconds after  
324 termination of the physical contact. Chasing represents a behavior indicating  
325 aggressiveness (i.e. the motivation to continue the fight after contact has ended), and



326 has also been used as a measure of aggressiveness in several other studies (e.g. Jang et  
327 al. 2008; Bertram et al. 2011; Fitzsimmons and Bertram 2013). Unlike the other two  
328 behaviors, chasing after an interaction is mutually exclusive for the two individuals, i.e.  
329 only the winner ever chases the opponent away. The mean number of chases recorded  
330 per individual per trial was 1.46 (min = 0, max = 22).

### 331 **Statistical analyses**

332 We conducted two sets of statistical analyses. First we partitioned the variation for  
333 each measured trait into within- and among-individual variance using a univariate  
334 mixed-effect modelling framework. Univariate models allow a straightforward  
335 interpretation of fixed effects, and allow for testing repeatabilities of single behaviors.  
336 As a second step, we estimated patterns of among-trait covariance at each estimable  
337 hierarchical level using a multivariate extension of this framework. All models were  
338 fitted using restricted maximum likelihood; dependent variables were mean-centered  
339 and variance standardized to facilitate comparison of variance components across  
340 traits. Throughout, we assumed a Gaussian error distribution, which was confirmed for  
341 all response variables after visual inspection of model residuals.

#### 342 *Univariate mixed-effects models*

343 Sources of variation in traits repeatedly measured in a non-social context (i.e., activity,  
344 exploration, and body weight) were estimated by fitting the focal trait as the response  
345 variable into a univariate mixed-effect model. To statistically control for potential

346 sources of variation not relevant to our biological hypotheses we included the  
347 following fixed effects: test sequence (covariate, range 0-11), time of the day (minutes  
348 from midnight, covariate), shelf (categorical variable with two levels), and location of  
349 the arena within a shelf (categorical variable with four levels). Test sequence was  
350 coded as the progressive order of repeats for the same assay (0 to 11), such that the  
351 fixed effect intercept of the model was estimated for the first test (e.g. Dingemanse et  
352 al. 2012a). Time of day was mean-centered, such that the fixed-effect intercept of the  
353 model was estimated for the phenotype expressed on the average testing time  
354 (following Dingemanse and Dochtermann 2013). We fitted random intercepts for  
355 group (8 levels), date (28 levels), and the focal individual's identity (38 levels). Effects  
356 of variables fitted to control for variation induced by the experimental design (time of  
357 day, shelf, location, group, and date) were unimportant (results not shown) and are  
358 not discussed further, except for the variable test sequence which explained significant  
359 variation.

360         Prior to the analysis of the traits repeatedly measured in a social context (i.e.,  
361 approaching, singing, and chasing), one individual in each trial was randomly assigned  
362 the focal and the other the opponent role. Since we hypothesize that focal behavior  
363 will depend on opponent phenotype, random intercepts for both focal and opponent  
364 identity are included in the univariate model of the focal individual's behavior.  
365 Furthermore, because the same individual is the focal in some aggression trials but the  
366 opponent in others, the model can be expanded to estimate the covariance (or  
367 correlation) between focal and opponent identity effects on focal behaviors. If positive  
368 (vs. negative), individuals that are on average aggressive as focal animals also tend to

369 induce (vs. reduce) aggressiveness in others when acting as the opponent. If opponent  
370 identity effects are present but this correlation is zero, individual-level traits of the  
371 opponent other than its aggressiveness affect the focal individual's aggressiveness.

372           While previous studies applying this design have analyzed sources of variation  
373 in the focal individual's behavior alone (e.g. Wilson et al. 2009), we were able to record  
374 detailed behaviors performed by both individuals in each dyad (i.e. the designation of  
375 focal and opponent identities as described above was arbitrary). In order to fully utilize  
376 all available behavioral data for parameter estimation we therefore modified the  
377 modelling strategy described above as follows. First, we estimated focal and opponent  
378 identity effects (as well as their correlation) for the behavior expressed by the focal  
379 individual. Second, we structured the dataset in a way that enabled us to  
380 simultaneously estimate focal and opponent identity effects (and their correlation) in  
381 the behavior of the opponent (as detailed in the Supplementary Text). Third, because  
382 focal and opponent roles were randomly assigned, (i) focal identity effects in the focal  
383 individual's behavior should equate opponent identity effects in the opponent's  
384 behavior, (ii) opponent identity effects in the focal individual's behavior should equate  
385 focal identity effects in the opponent's behavior, and (iii) the covariance between focal  
386 and opponent identity effects as well as any other fixed-effect (e.g. time of day) and  
387 random-effect (e.g. group, date) estimates should be the same for the focal's and  
388 opponent's aggressive behavior. We thus implemented a bivariate mixed-effect model  
389 with the focal's and opponent's behavior fitted as the two response variables where  
390 parameters that were logically identical were constrained to be identical. This bivariate  
391 implementation thereby enabled us to utilize all data at hand while simultaneously

392 avoiding pseudo-replication and biasing effects of arbitrary assignment of focal versus  
393 opponent roles. Because of the imposed model constraints, the bivariate model  
394 effectively represents a univariate model, and is treated as such in the remaining text.

395 Adjusted individual repeatability was estimated for each trait by calculating the  
396 proportion of the total phenotypic variance not attributable to fixed effects that was  
397 explained by individual identity (Nakagawa and Schielzeth 2010). For behaviors  
398 expressed in a social context, we additionally calculated the proportion of the total  
399 phenotypic variance that was explained by the identity of the opponent and was not  
400 attributable to fixed effects, which we call 'adjusted opponent repeatability' (Wilson et  
401 al. 2009).

#### 402 *Multivariate mixed-effects models*

403 We estimated patterns of trait covariance at each estimable hierarchical level using a  
404 multivariate mixed-effects model. We fitted activity, exploration, and weight as  
405 response variables as well as approach as a proxy for aggressiveness. Though we had  
406 measured various proxies of aggressiveness (detailed above), we used here only  
407 approach because multivariate models including other proxies of aggression did not  
408 converge. However, we were able to confirm with sets of simpler models that all  
409 proxies for aggressiveness were highly correlated (Results not shown), implying that  
410 our decision to use approach as a measure of aggressiveness was justified (for a  
411 further discussion on rationale of this approach, see Araya-Ajoy & Dingemanse 2014).

412 To avoid over-parameterization, we only included fixed (sequence) and random effects  
413 (identity of the focal individual) that explained significant variation in the univariate

414 analyses (see Results). Modelling random intercepts for the focal individual's identity  
415 enabled decomposition of phenotypic covariances (and hence correlations derived  
416 from them) into among-individual and residual within-individual components (Wilson  
417 et al. 2009). Among-individual correlations occur when an individual's average  
418 phenotype over all repeated measures is correlated across traits, called a 'behavioral  
419 syndrome' in the context of behavior (Dingemanse et al. 2012b); within-individual  
420 correlations occur when two phenotypic traits show correlated changes in expression  
421 within the same individual due to the combined effects of integration of within-  
422 individual plasticity and/or correlated measurement error (Dingemanse and  
423 Dochtermann 2013). We further included random intercepts for the opponent's  
424 identity for the response variable approach during the aggression test, which enabled  
425 us to assess whether individuals were repeatable in the aggressiveness elicited in  
426 conspecifics (we will sometimes present this indirect effect as a trait, 'eliciting  
427 aggressiveness', for clarity). Because the same individual repeatedly played focal in  
428 some and opponent in other trials (see above), we also estimated (i) the correlation  
429 between the focal and opponent identity effect on aggressiveness (see above for  
430 biological interpretation) as well as (ii) the correlation between the opponent identity  
431 effect on aggressiveness and the focal identity effect on each of the other traits  
432 (activity, exploration, body weight). This latter type of correlation tests whether  
433 individuals that elicit more aggressiveness in others are themselves more (or less)  
434 active, explorative, or heavier than individuals eliciting less aggressiveness in others.

435 *Significance testing in mixed-effects models*

436 We tested statistical significance of fixed effects using numerator and denominator  
437 degrees of freedom (df) estimated from the algebraic algorithm in ASReml 3.0  
438 (Gilmour et al. 2009). We used likelihood ratio tests (LRTs) to evaluate the statistical  
439 significance of random effects. This  $\chi^2$ -distributed test statistic is calculated as twice  
440 the difference in log-likelihood between a model where a target random effect was  
441 fitted versus not fitted (Shaw 1991). Variances are bound to be positive, therefore  
442 probability (P) of a LRT applied to a variance was calculated assuming an equal mixture  
443 of  $P(\chi^2, df=0)$  and  $P(\chi^2, df=1)$ , i.e.  $df=0.5$  (Self and Liang 1987; Pinheiro and Bates  
444 2000; Visscher 2006). Covariances (correlations) are not bound to be positive, and  
445 their probability was therefore calculated assuming  $P(\chi^2, df=1)$ . LRTs involving one  
446 variance and one covariance were tested assuming an equal mixture of  $P(\chi^2, df=1)$  and  
447  $P(\chi^2, df=2)$ , i.e.  $df=1.5$ . We applied this latter test, for example, to evaluate support for  
448 opponent identity effects in aggressiveness, which requires fitting both a variance (i.e.  
449 the variance attributable to the opponent's identity) and a covariance (i.e. the  
450 covariance between focal and opponent identity effects). All models were  
451 implemented in ASReml 3.0 (Gilmour et al. 2009).

452 *Structural equation modelling*

453 To test our *a priori* hypotheses on behavioral syndrome structure, we applied  
454 structural equation modelling to nine *a priori* conceived scenarios based on the  
455 behavioral syndrome literature (described in Figure 1). We analyzed the among-  
456 individual correlation matrix estimated from the multivariate mixed-effects model

457 using the package 'SEM' in the software R v. 3.1.0. (Team R Core 2012). We then  
458 statistically compared the models using the Akaike information criterion (AIC) (Akaike  
459 1973; Burnham and Anderson 2002), and evaluated the relative support for each  
460 based on AIC differences relative to the model with the lowest AIC ( $\Delta AIC$ ), which  
461 represents the best fitting model; we also calculated each model's weight and relative  
462 likelihood (Anderson 2008).

## 463 **RESULTS**

### 464 **Sources of variation in single traits**

465 Exploration, approach, and weight changed significantly within individuals as a  
466 function of test sequence (Supplementary Table S1). On average individuals became  
467 leaner, less explorative, and less aggressive over the 12-day course of the experiment.  
468 None of the traits were significantly affected by tag color, time of day, testing shelf or  
469 within-shelf test location (Results not shown); neither did the traits vary between  
470 groups or days (Supplementary Table S1).

471 All traits except singing were significantly repeatable (Table 2); adjusted  
472 behavioral repeatabilities were highest for exploration (0.40, SE 0.10) and activity  
473 (0.33, SE 0.09), and substantially lower (though significant) for approaching (0.19 SE  
474 0.06) and chasing (0.09 SE 0.05) during the aggression test. Approaching and chasing  
475 were significantly affected by the identity of the opponent, though singing was not  
476 (Table 2). For approaching and singing, the proportion of variance explained by

477 opponent identity (i.e. adjusted opponent repeatability) was of a similar magnitude  
478 (0.16, 0.06 SE for approaching; 0.12, 0.05 SE for chasing) when compared to adjusted  
479 individual repeatability (0.19 and 0.09, respectively). Opponent identity effects imply  
480 that unidentified individual-level characteristics of opponents affect the behavior of  
481 focal individuals. These effects did not appear to be due to repeatable differences in  
482 approaching, chasing or singing among opponents, since the correlation between focal  
483 and opponent identity effects was rather weak and non-significant for all three  
484 behaviors expressed in the aggression test (Table 2).

485         Body weight was, as expected also repeatable (Table 1). Interestingly,  
486 repeatability was relatively low (0.46, SE 0.11) compared to what might typically be  
487 expected. This indicates that body weight strongly depended on day-to-day changes in  
488 environmental conditions experienced by our crickets.

#### 489 **Among-trait correlations**

##### 490 *Among-individual correlations*

491 The multivariate mixed-effects model provided strong evidence for the existence of  
492 among-individual correlations between most traits (Table 3; Supplementary Table S2,  
493 Supplementary Table S4). Individuals that were on average relatively aggressive  
494 towards conspecifics were also relatively active in a novel environment ('explorative')  
495 and relatively active in a familiar environment, providing strong evidence for the



496 existence of an aggressiveness-activity-exploration syndrome. Furthermore, aggressive  
497 individuals were relatively heavy compared to less aggressive individuals.

498 Our univariate analysis implied that aggressiveness elicited in conspecifics also  
499 harbored among-individual variation (Table 2). Aggressiveness elicited in conspecifics  
500 was not (tightly) related to an individual's average level of aggressiveness, owing to a  
501 non-significant correlation between focal and opponent identity effects (Table 2, see  
502 also Table 3). Our multivariate analysis revealed that individuals that were on average  
503 relatively explorative in novel environments also elicited approach by opponents in the  
504 aggression test (Table 3). This among-individual correlation was relatively strong (0.45,  
505 SE 0.17), implying that personality variation with respect to exploratory tendency  
506 constitutes a major factor explaining why individuals are repeatable in the level of  
507 aggressiveness that they elicit in others.

#### 508 *Residual within-individual correlations*

509 Various traits were also correlated within individuals, implying the existence of  
510 integration of within-individual plasticity across multiple traits and/or correlated  
511 measurement error. Specifically, when individuals increased their body weight across  
512 trials, they also became significantly more explorative in the novel environment  
513 (though not more active in familiar environments) (Table 3, Supplementary Table S4).  
514 In contrast to patterns observed at the among-individual level, most traits were only  
515 weakly and non-significantly associated at the within-individual level.

## 516 Behavioral syndrome structure

517 Based on literature, we constructed and compared nine different *a priori* conceived  
518 models of behavioral syndrome structure (Figure 1). The domain-general syndrome,  
519 which included correlations between activity, exploration, aggression, and aggression  
520 elicited (model 3; Figure 4), best explained the observed data, as inferred from it  
521 having the lowest AIC score (Table S3). The second-best model, hypothesizing a  
522 domain-general structure with only exploration eliciting aggressiveness (model 9) also  
523 had a relatively good fit to the data (model 9:  $\Delta\text{AIC}=2.35$ ; Table S3). The domain  
524 general model nevertheless fitted the data 3.26 times better compared (Akaike weight  
525 of model 3 (0.75) divided by the Akaike weight of model 9 (0.23); Table S3), and is  
526 therefore treated as the sole best-fitting model throughout the remaining text.

527 The four observed behaviors were, notably, also partially independent from  
528 each other, as the syndrome structure implied in our best-supported model explained  
529 relatively little variation in aggressiveness (Figure 4). Exploration behavior represented  
530 an extreme case, where all the among-individual variance was captured by the  
531 hypothesized syndrome structure. We verified that this was not an artefact caused by  
532 the tight among-individual correlation between exploration and activity, as re-analyses  
533 of the data after excluding activity yielded the same results (not shown). These  
534 findings therefore provide strong support for the presence of a domain-general  
535 syndrome that includes indirect effects.

## 536 DISCUSSION

537 This study investigated the occurrence of indirect effects on aggressiveness in male  
538 European field crickets, and asked whether such effects could be driven by personality.  
539 As expected, we found evidence for indirect effects as individuals differed consistently  
540 in the level of aggressiveness that they elicited in others. Direct and indirect effects on  
541 focal behavior were integrated within an overarching behavioral syndrome, existing  
542 largely because more explorative individuals were also more active and aggressive, and  
543 tended to elicit more aggressiveness in conspecifics. This finding constitutes  
544 unambiguous evidence for the notion that an individual's aggressiveness-activity-  
545 exploration type influences aggressiveness of conspecifics. Our study thus implies that  
546 repeatable differences in behavior represent an important component of the social  
547 environment affecting social behaviors of conspecifics. If representing heritable  
548 patterns of variation, these findings imply that personality variation in the  
549 environment will have consequences for evolutionary trajectories of socially expressed  
550 behaviors that would only be appropriately predicted by quantitative genetics theory  
551 developed to model indirect genetic effects (Montiglio et al. 2013; Dingemanse and  
552 Araya-Ajoy 2015).

### 553 **Individual and opponent repeatabilities**

554 All behaviors (except singing during the aggression test) were repeatable and their  
555 estimates were within the range reported for behavioral traits (meta-analysis: Bell et  
556 al. 2009). Values of repeatabilities of non-social behavioral traits (activity, exploration)  
557 were also similar to those documented previously in our and other cricket species

558 (Wilson et al. 2010; Niemelä et al. 2012a,b; Dochtermann and Nelson 2014; Fisher et  
559 al. 2015a,b). In contrast, behaviors expressed in a social context (approaching, singing,  
560 chasing) had relatively low repeatabilities (0.09-0.19) compared to those expressed in  
561 a non-social context (0.33-0.40; Table 2). This was not unexpected since the social  
562 environment experienced varied substantially within-individuals across trials, whereas  
563 the non-social trials represented a more stable environment. As advocated in recent  
564 reviews (Briffa et al. 2015; Dingemanse and Araya-Ajoy 2015), our setup allowed us to  
565 assess repeatability of social behavior as it would typically be expressed in the wild  
566 (i.e., across an ecologically relevant sample of partners) rather than artificially  
567 minimizing social environment effects by using dummies or video playback stimulation.  
568 Few other studies of aggression have explicitly modelled such indirect effects by  
569 considering identity effects of interaction partners; all such studies have focused on  
570 vertebrates. Importantly, the repeatability estimates in these few studies were in the  
571 same range of ours (0.11 to 0.27; Wilson et al. 2009, 2011, 2013). These estimates of  
572 repeatability are lower compared to the ones reported by meta-analyses (Bell et al.  
573 2009), where aggression was one of the most repeatable classes of behaviors (roughly  
574  $R=0.50$ ). Our findings therefore imply that published estimates might often be  
575 substantially biased upwards due to the usage of unnatural standardization  
576 standardized experimental conditions (e.g. size matching) that may lack ecological  
577 relevance.

578           Of course, the magnitude of indirect effects that we detected might be less  
579 substantial under field conditions. In the wild, other environmental effects (kept stable  
580 under laboratory conditions), such as temperature, availability of mates, or predators

581 might affect the amount of among-individual variation in aggressiveness and  
582 aggressiveness elicited in conspecifics. Importantly, studies based on dyadic  
583 interactions might typically overestimate the magnitude of IGEs for species that  
584 naturally interact in larger groups (Hadfield and Wilson 2007; Bijma 2014). This critic  
585 may not be applicable in our case because male-male interactions over ownership of  
586 burrows or mates in our study species are typically dyadic under field conditions too  
587 (personal observation NJD and FS, Fisher et al. 2016). Therefore, our staged dyadic  
588 interactions accurately reflect the natural social environment in which aggressive  
589 interactions take place.

590           The opponent effect provides clear evidence for the notion that individuals  
591 plastically adjusted their aggression level in response to the opponent's phenotypic  
592 traits that vary at the individual level. Among-individual variation in opponents  
593 therefore explains part of the within-individual variation in aggression, which can be  
594 interpreted as a social form of phenotypic plasticity, or 'social responsiveness'  
595 (Webster and Ward 2011; Taborsky and Oliveira 2012; Wolf and McNamara 2013; Wolf  
596 and Krause 2014; Dingemanse and Araya-Ajoy 2015). By modelling the opponent  
597 identity as a random effect, we integrated the influence of all the aspects of an  
598 opponent's individual-level phenotype. However, by also measuring phenotypic traits  
599 of the opponent (including its personality), we were able to tease apart which traits  
600 were driving the opponent effect as we describe below.

601 **Individual-level correlations of behavioral traits across contexts**

602 Among-individual correlations between focal aggression, exploration and activity were  
603 positive, confirming our predictions and revealing a syndrome structure including  
604 behaviors expressed in social and non-social contexts. Positive correlations among  
605 activity and exploration behavior are well-documented in the behavioral syndrome  
606 literature (Sih et al. 2004b; Réale et al. 2007; Garamszegi et al. 2012), and are generally  
607 supported in crickets (e.g. Wilson et al. 2010, but see Dochtermann and Nelson 2014)  
608 The coping style literature predicts the presence of bold, exploratory, and aggressive  
609 individuals called 'proactive' types (Koolhaas et al. 1999; Coppens et al. 2010), which  
610 we indeed found in our study. We note however, that the relationship between  
611 aggressive behavior and other personality traits is often population-specific (Bell 2005;  
612 Dingemanse et al. 2007). Our findings therefore confirm the notion that behaviors  
613 from different functional contexts covary and that domain-general syndromes may be  
614 common among organisms (Sih et al. 2004a,b; Bell 2007; Dochtermann and Jenkins  
615 2007). Such a structure may be important from an evolutionary perspective because  
616 trait correlations are known to affect evolutionary trajectories (Lande and Arnold  
617 1983; Sih et al. 2004b; Sprenger et al. 2012; Dochtermann and Dingemanse 2013).

618         The repeatable part of an individual's weight was correlated with an  
619 individual's average level of exploration and aggressive behavior, implying that  
620 repeatable variation in body condition might have proximately underpinned the  
621 observed behavioral syndrome (see e.g. Royauté et al. 2015b). However, post hoc  
622 analyses showed this not to be the case: we S3conducted a set of SEMs expanding

623 upon our best-supported (domain-general) model. We then compared the relative fit  
624 of models where body weight constituted an additional expression of the latent  
625 variable rather than representing the latent variable itself (versus a null model where  
626 body weight was not associated with behavior) (Figure S1, Table S5). The domain-  
627 general model, where body weight also constituted an expression of the latent  
628 variable, was best supported; this finding implies that weight indeed did not 'drive' the  
629 syndrome structure. Interestingly, body size has often been hypothesized to  
630 represent an important mediator of agonistic behavior in the animal contest literature  
631 (Briffa et al. 2015). Our study implies that social partners responded to personality  
632 rather than other characteristics of conspecifics, such as their body weight. It is of  
633 course possible that crickets instead responded to *relative* differences in body weight.  
634 However, this scenario is not supported by our data because it should result in effects  
635 of the unique combination of focal and partner identity effects (so-called 'pair identity'  
636 effects), which were not present when modelled in our data (Results not shown).

637         Because behaviors are plastic, feedbacks between traits may occur during  
638 interactions and can either increase or decrease the amount of among-individual  
639 variation (Moore et al. 1997; Wolf et al. 1998, 1999; Dingemans and Araya-Ajoy 2015;  
640 Sih et al. 2015). The level of aggression in one individual is often affected by the  
641 aggression displayed by the interacting individual in escalated fights (Huntingford and  
642 Turner 1987), and positive correlations between the focal and opponent effects on this  
643 trait would indicate this type of reciprocal feedback. By contrast, if aggression in part  
644 reflects dominance (e.g., with dominant individuals only displaying aggression towards  
645 subordinates), negative correlations can arise (Wilson et al. 2011). Interestingly, the

646 lack of a significant correlation between being aggressive and eliciting aggression in  
647 others may suggest the absence of feedback loops (either positive or negative) in this  
648 trait. However, our SEM-analysis uncovered a relationship between being aggressive  
649 and eliciting aggressiveness underpinned by the behavioral syndrome (see below),  
650 indicating that positive feedbacks may have been present instead.

651           In contrast to patterns found at the among-individual level, most traits were  
652 only weakly and non-significantly associated at the within-individual level. This  
653 suggests that trait correlations were caused by different proximate mechanisms within  
654 versus among individuals (van Noordwijk and de Jong 1986; Hadfield et al. 2007). This  
655 finding contrasts recent studies reporting within- and among-individual correlations of  
656 similar magnitude (Araya-Ajoy and Dingemanse 2014; Brommer et al. 2014; Royauté et  
657 al. 2015a). The partial difference in correlation structure at the within vs. among  
658 individual level highlights the importance of separating the two in order to make  
659 unbiased inferences about behavioral syndrome structure (Dingemanse et al. 2012b;  
660 Dingemanse and Dochtermann 2013; Brommer et al. 2014).

#### 661 **Integration of indirect effects as part of a multivariate behavioral phenotype**

662 All behavioral traits were linked with aggressiveness elicited in conspecifics as part of a  
663 behavioral syndrome, highlighting the key importance of considering the opponent's  
664 personality in the expression of social behaviors. These findings are supported both by  
665 the multivariate mixed-effect model and the SEM. Because exploration, aggression and  
666 activity were highly correlated, we expected that they would jointly affect the



667 aggression of conspecifics, implying that the syndrome as a whole would explain the  
668 indirect effects on aggression (model 3). This was indeed the case: despite the  
669 ‘penalization’ for a higher number of parameters, the best-supported model was the  
670 one with a domain-general behavioral syndrome (Table S3, Figure 4). A mechanism for  
671 the overarching indirect effect could be that individuals who explore a new  
672 environment are likely to also be involved in more social interactions and this may  
673 result in more attacks received by others within an aggression trial. This interpretation  
674 is warranted by the strong bivariate correlation between indirect effects on aggression  
675 and exploration in the among-individual correlation matrix. Indeed, exploration was  
676 the factor loading most strongly with the latent variable in our structural equation  
677 model. The overarching importance of exploration behavior was implied by the  
678 second-best fitting model, which attracted some support (Table S3). Indeed,  
679 movement variables may represent obvious cues of an individual’s aggressiveness.  
680 Nevertheless, other components of the behavioral syndrome definitely played an  
681 important role in eliciting aggression because a full domain-general model was over  
682 three times better supported. Because indirect effects were integrated as part of the  
683 behavioral syndrome, our findings imply that studies of the evolution of behavior  
684 should explicitly recognize that behavioural phenotypes are multivariate and include  
685 responses of the social environment.

## 686 **Evolutionary implications**

687 Adaptive explanations in the field of behavioral ecology are often based solely on  
688 phenotypic observations, with evolutionary inference relying on the ‘phenotypic  
689 gambit’, assuming correspondence of phenotypic and genetic patterns of (co)variation  
690 (Grafen 1984). Support for this position has been provided by a behavioral meta-  
691 analysis in which phenotypic correlations explained 75% of the variation in genetic  
692 correlations (Dochtermann 2011), although this correspondence should not readily be  
693 assumed for individual cases. Despite this important caveat in our knowledge, the  
694 presence of indirect effects documented here suggests a strong potential for IGEs to  
695 contribute to evolutionary trajectories of aggression and/or traits genetically  
696 correlated with it, which in our case include exploration and activity. Our study thus  
697 implies that behavioral types certainly have the potential to alter the evolutionary  
698 trajectories of behavioral traits by inducing indirect effects in a social context. Further  
699 partitioning the repeatable among individual behavioral variance into additive genetic  
700 and environmental effects would enable us to move from studying causes to revealing  
701 the consequences of personalities. This would ultimately enable us to predict the  
702 evolutionary consequences of the very existence of personality types in natural  
703 populations.

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939

940 **FIGURE LEGENDS**

941 **Figure 1.** Models (1–9) of hypothesized relationships between behaviors (syndrome  
942 structure). Unidirectional arrows represent causal relationships between variables,  
943 bidirectional arrows represent correlations. Solid lines represent relationships present  
944 across the whole set; dashed, dotted, and mixed lines represent relationships  
945 expressed in specific syndrome structures. Model 1: null model of behavioral  
946 independency (Coleman and Wilson 1998). Model 2: coping styles with ‘proactive’ and  
947 ‘reactive’ types (correlations among exploration, activity and aggression) (Koolhaas et  
948 al. 1999; Coppens et al. 2010). Model 3: The dashed arrow is active; a modification of  
949 model 2, where all behaviors including the indirect effects are expression of a general-  
950 domain syndrome. Models 4-6: the dashed arrow alone is active in model 4, the mixed  
951 arrow alone in model 5, both in model 6. Semi-independency of modules (Sih et al.  
952 2004b) for different contexts (social and non-social behaviors). Model 7: the dashed  
953 arrow is active; combination of model 2 with reciprocal feedbacks on aggressiveness  
954 (covariance between aggression and eliciting aggressiveness) (Moore et al. 1997;  
955 Wilson et al. 2009). Models 8-9: the mixed arrow is active only in model 8, the dotted  
956 arrow is active only in model 9; combination of model 2 with ideas from (Verbeek et al.  
957 1996). Exploration, aggression, and activity are part of syndrome and either  
958 exploration (8) or activity (9) has a direct effect on aggressiveness of others.

959 **Figure 2.** Daily experimental timeline, showing a top view of the experimental setups  
960 used to measure a) activity b) exploration and c) aggression.

961 **Figure 3.** Graphical illustration of the indirect effect of exploration behavior on  
962 aggression. We present here correlations between BLUPs derived from the  
963 multivariate mixed-effects model: individual-mean levels of exploration (x axis) and  
964 aggression elicited in the opponents (y axis). Each point represents the BLUP of an  
965 individual (in standard deviation units) with its associated standard error.

966 **Figure 4.** Parameter estimates of the structural equation model that best fitted our  
967 data. For each variable, we report the variance explained by the SEM structure ( $R^2$ ) and  
968 factor loadings with the corresponding SE in parentheses. These represent how  
969 behavioral responses are predicted to change based on changes to the underlying  
970 syndrome structure.

971 **TABLES**

972 **Table 1.** Timetable of the daily experimental procedure (see the Methods for further  
 973 details). Each individual was subjected to this procedure for 12 consecutive days.

<b>Order within day</b>	<b>Action/behavioral test</b>	<b>Location</b>	<b>Duration</b>
1	Food and water removal + acclimation	Home container	30 min
2	Activity	Home container	1 hr
3	Transfer to exploration arena		1 min
4	Exploration	Exploration arena	1 hr
5	Shelter removal + acclimation	Exploration arena	10 min
6	Aggression	Exploration arena w/o divider	10 min



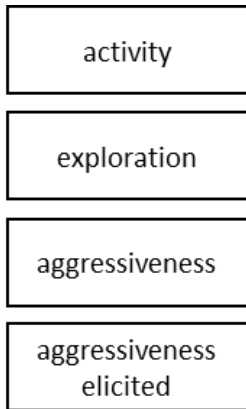
974 **Table 2.** The proportion of total phenotypic variation not attributable to fixed effects explained by focal identity (adjusted individual  
975 repeatability; direct effects) and opponent identity (adjusted opponent repeatability; indirect effects), and the correlation between focal  
976 and opponent identity effects. Positive values for this correlation indicate patterns where individuals that on average express a high value  
977 for the given behavior also elicit an increased expression of this behavior in opponents. Estimates (with SE) are derived from univariate  
978 mixed-effect models with random intercepts for focal and opponent identity as appropriate, and associated  $\chi^2$ -values, degrees of  
979 freedom (df), and values of  $P$  derived from likelihood ratio tests where the full model is compared to one where the random effect of  
980 interested was excluded.

Trait	Individual repeatability (SE)	$\chi^2$ df	<i>P</i>	Opponent repeatability (SE)	$\chi^2$ df	<i>P</i>	Focal - opponent correlation (SE)	$\chi^2$ df	<i>P</i>
Activity	<b>0.33</b> (0.09)	65.60 <sub>0.5</sub>	<b>&lt;0.01</b>	-	-	-	-	-	-
Exploration	<b>0.40</b> (0.10)	92.56 <sub>0.5</sub>	<b>&lt;0.01</b>	-	-	-	-	-	-
Weight	<b>0.46</b> (0.11)	253.37 <sub>0.5</sub>	<b>&lt;0.01</b>	-	-	-	-	-	-
Approach	<b>0.19</b> (0.06)	17.65 <sub>1.5</sub>	<b>&lt;0.01</b>	<b>0.16</b> (0.06)	14.74 <sub>1.5</sub>	<b>&lt;0.01</b>	0.22 (0.26)	0.70 <sub>1</sub>	0.40
Chase	<b>0.09</b> (0.05)	6.49 <sub>1.5</sub>	<b>&lt;0.05</b>	<b>0.12</b> (0.05)	9.72 <sub>1.5</sub>	<b>&lt;0.05</b>	-0.34 (0.39)	0.71 <sub>1</sub>	0.40
Sing	0.07 (0.05)	3.60 <sub>1.5</sub>	0.11	0.05 (0.05)	1.86 <sub>1.5</sub>	0.28	-0.21 (0.56)	0.15 <sub>1</sub>	0.69

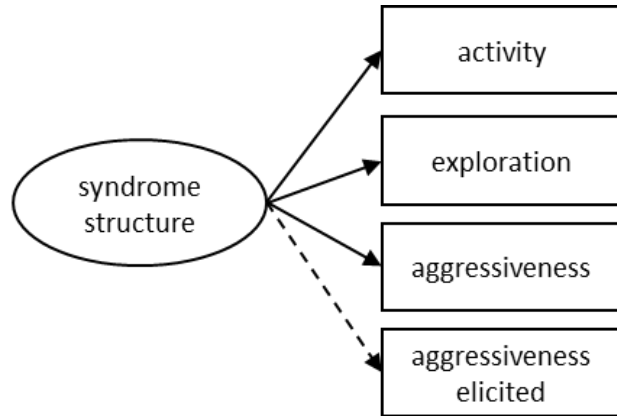
982 **Table 3.** Estimated among- and residual within-individual correlations (with associated SE) for two non-social behaviors (activity and  
983 exploration), two social behaviors (aggressiveness and aggressiveness elicited in opponents), and weight. We present among-individual  
984 correlations on the upper off-diagonals and within-individual correlations on the lower off-diagonals. Correlations printed in bold-face are  
985 significant based on likelihood ratio tests derived from the multivariate model detailed in the main text. Note that the experimental setup  
986 solely allowed for estimating of among-individual correlations with aggressiveness elicited in other individuals.

	<b>Aggressiveness</b>	<b>Activity</b>	<b>Exploration</b>	<b>Weight</b>	<b>Aggressiveness elicited</b>
<b>Aggressiveness</b>	-	<b>0.39</b> (0.18)	<b>0.39</b> (0.17)	<b>0.43</b> (0.16)	0.09 (0.22)
<b>Activity</b>	-0.02 (0.07)	-	<b>0.81</b> (0.08)	<b>0.17</b> (0.15)	0.16 (0.20)
<b>Exploration</b>	0.03 (0.07)	<b>0.17</b> (0.06)	-	0.22 (0.14)	<b>0.45</b> (0.17)
<b>Weight</b>	0.05 (0.06)	0.06 (0.06)	<b>0.13</b> (0.06)	-	-0.31 (0.18)

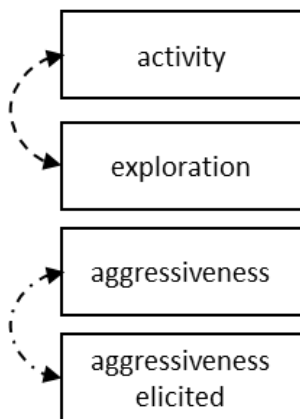
**FIGURES**



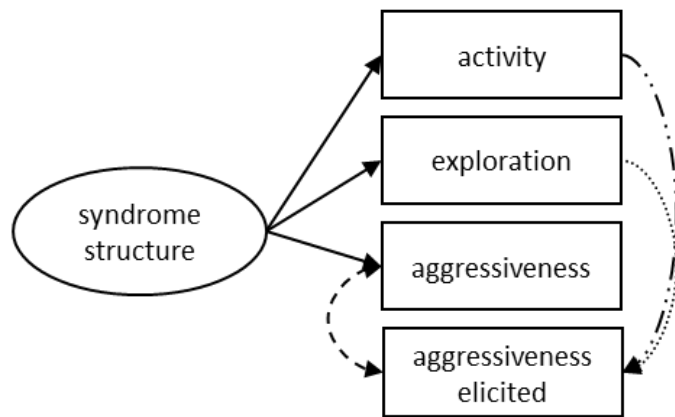
Model 1



Models 2-3



Models 4-6



Models 7-9

**Figure 1.**

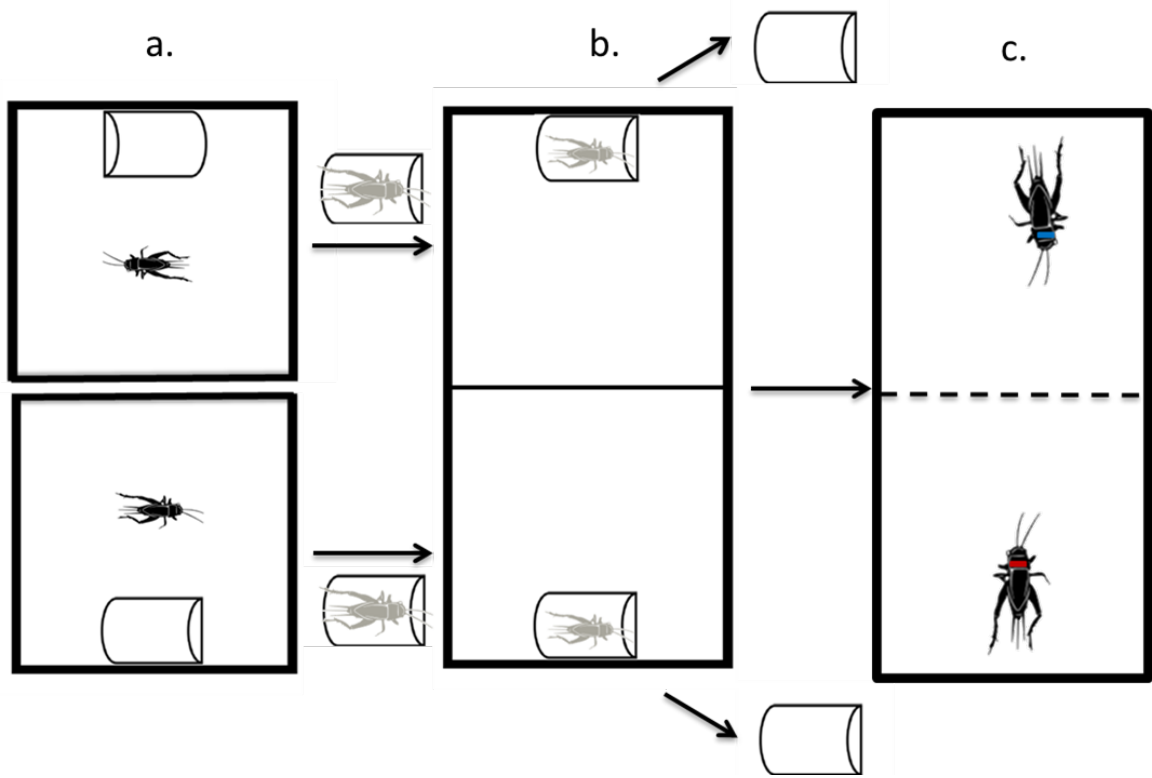


Figure 2.

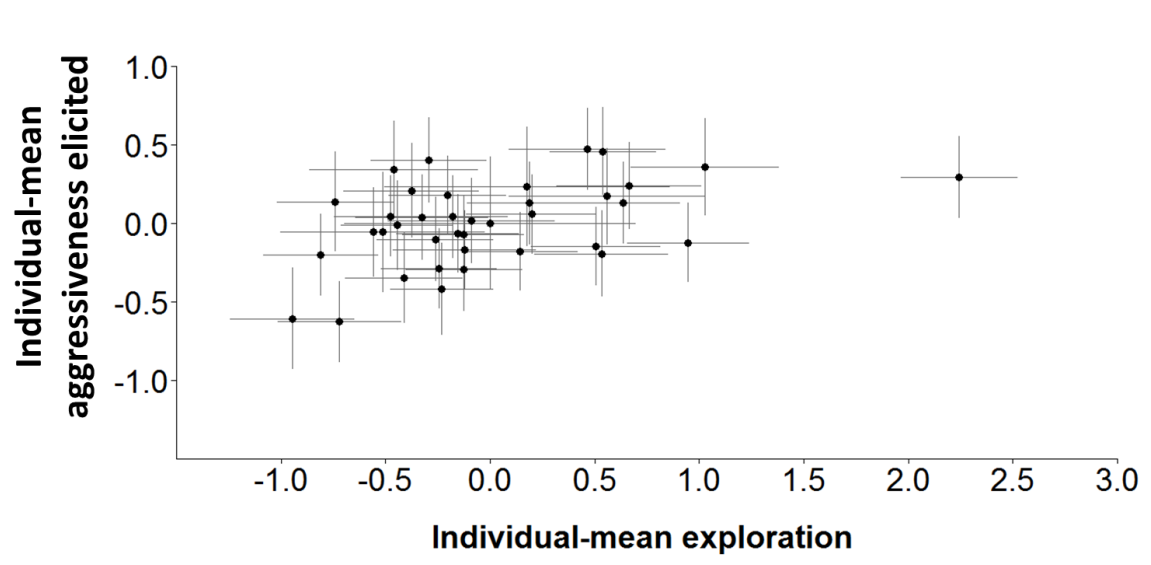
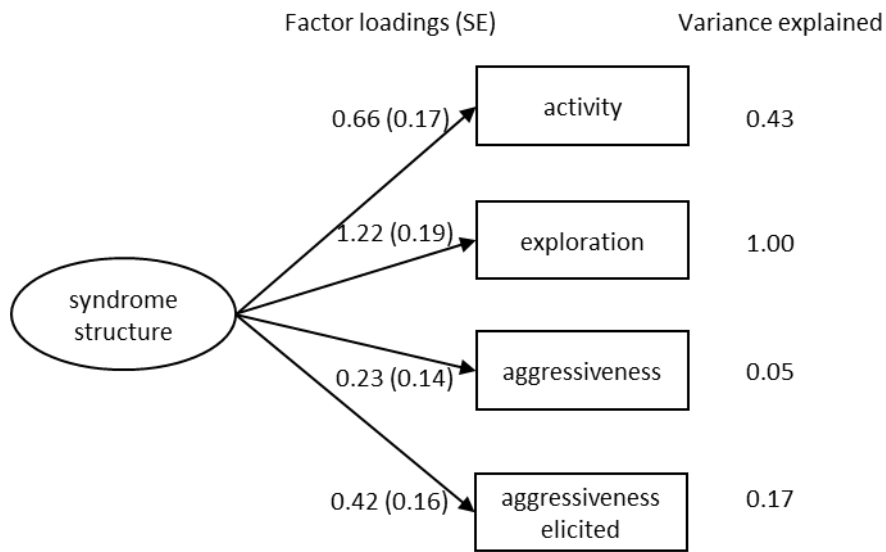


Figure 3.



**Figure 4.**





988 **SUPPLEMENTARY TEXT**

989 As detailed in the main text, analyses of aggressiveness that estimate focal and  
 990 opponent identity effects typically focus on variation in the behavior expressed by the  
 991 (arbitrarily assigned) focal individual alone. Here we detail how we incorporated  
 992 information on the same behavior measured on the opponent in the statistical model  
 993 while avoiding pseudo-replication. We started with the following data structure, where  
 994 each line consisted of information regarding the identity of both individuals, one  
 995 arbitrarily called ‘Individual A’ and the other ‘individual B’, with associated information  
 996 regarding their aggressiveness:

<b>trial ID</b>	<b>Individual A</b>	<b>Individual B</b>	<b>Aggressiveness A</b>	<b>Aggressiveness B</b>
1	14	12	3	6

997 We then rearranged the data in the following way:

<b>trial ID</b>	<b>Focal</b>	<b>Opponent</b>	<b>Dataset</b>	<b>Aggressiveness 1</b>	<b>Aggressiveness 2</b>
1	14	12	1	3	NA
1	12	14	2	NA	6

998 In this re-ordered dataset, the data is printed over two lines, once viewing individual A  
 999 as the ‘focal’ individual in trial 1 (Dataset 1) and once viewing individual B as the ‘focal’  
 1000 individual in trial 1 (Dataset 2). Importantly, the behavior of the individual dubbed

1001 'focal' in Dataset 1 was printed in another column (Aggressiveness 1) than the  
1002 behavior of the individual dubbed 'focal' in Dataset 2 (column Aggressiveness 2). We  
1003 then proceeded to estimate model parameters under the imposed assumption that all  
1004 model parameters (fixed effect coefficients and (co)variance components) are equal  
1005 for the homologous traits as defined in the two non-overlapping data sets (i.e.,  
1006 Aggressiveness 1, Aggressiveness 2). Practically this can be achieved for a pair of  
1007 homologous traits by fitting a bivariate mixed effect model with the following code in  
1008 ASReml, which we have annotated in footnotes below:

```
1009 agg1 agg2 ~mu !r !{Trait.foc Trait.opp !}
```

```
1010 1 2 1
```

```
1011 0
```

```
1012 Trait 0 US !GPZP !=a0a !S2==1 #A
```

```
1013 0.5
```

```
1014 0 0.5
```

```
1015 Trait.foc 2
```

```
1016 4 0 US !GPZPUZPZUZP !=a0ab0c0b0c #B
```

```
1017 0.5
```

1018 0 0.5

1019 0.1 0 0.5

1020 0 0.1 0 0.5

1021 foc

1022 The `!{Trait.tag Trait.opp !}` command enables joining the focal and opponent variance-  
1023 covariance matrix into a single matrix such that covariances between focal and  
1024 opponent identity effects can be estimated.

1025 A – Residual covariance structure. Residual variances are constrained to be positive  
1026 and equal for the two traits. A starting value of 0.5 is supplied. Since no line of data is  
1027 informative for both traits the residual covariance is not estimable and is fixed  
1028 (arbitrarily) to zero.

1029 B – Among-individual covariance structure. There are four random effects in the model  
1030 (focal and opponent effects on two homologous traits) so a 4x4 covariance matrix is  
1031 specified. Variances are constrained to be positive (starting value of 0.5 supplied for  
1032 each), while covariance terms are identifiable between focal and opponent effects  
1033 with each trait (starting value of 0.1 supplied). Covariance parameters for Aggression 1  
1034 are constrained to equal those of Aggression 2. All other cross-trait covariance terms  
1035 are fixed to zero.

1036 **SUPPLEMENTARY TABLES**

1037 **Table S1.** Parameter estimates (with standard errors) of fixed and random effects  
1038 derived from our univariate models fitted to analyze variation in activity, exploration,  
1039 weight, and aggressiveness (approach, chasing, singing). Values printed in bold face  
1040 represent significant effects based either on Wald F tests (fixed effects) or likelihood  
1041 ratio tests (random effects) detailed in the main text. Values of adjusted individual and  
1042 opponent repeatability reported in Table 2 (main text) are derived from the univariate  
1043 models printed here.

1044 **Table S1 continued**

	<b>Activity</b>	<b>Exploration</b>	<b>Weight</b>	<b>Approach</b>	<b>Chase</b>	<b>Sing</b>
<b>Fixed effects</b>	$\beta$ (SE)	$\beta$ (SE)	$\beta$ (SE)	$\beta$ (SE)	$\beta$ (SE)	$\beta$ (SE)
Intercept	-0.01 (0.20)	0.16 (0.22)	-0.21 (0.23)	0.33 (0.12)	0.07 (0.10)	0.06 (0.01)
Sequence	-0.03 (0.02)	<b>-0.04</b> (0.01)	<b>-0.04</b> (0.01)	<b>-0.07</b> (0.01)	-0.02 (0.02)	-0.01 (0.02)
<b>Random effects</b>	$\sigma^2$ (SE)	$\sigma^2$ (SE)	$\sigma^2$ (SE)	$\sigma^2$ (SE)	$\sigma^2$ (SE)	$\sigma^2$ (SE)
Focal	<b>0.32</b> (0.11)	<b>0.40</b> (0.10)	<b>0.46</b> (0.11)	<b>0.19</b> (0.07)	<b>0.09</b> (0.05)	0.08 (0.05)
Opponent	-	-	-	<b>0.16</b> (0.07)	<b>0.12</b> (0.06)	0.05 (0.05)
Group	0.01 (0.07)	0.06 (0.10)	0.15 (0.15)	0.03 (0.09)	0.00 (0.00)	0.00 (0.00)
Date	0.00 (0.02)	0.00 (0.01)	<b>0.18</b> (0.06)	0.00 (0.03)	0.00 (0.00)	0.04 (0.09)
Residual	<b>0.66</b> (0.07)	<b>0.55</b> (0.08)	<b>0.22</b> (0.04)	<b>0.63</b> (0.07)	<b>0.78</b> (0.08)	<b>0.85</b> (0.09)
<b>Covariance</b>				Cov (SE)	Cov (SE)	Cov (SE)
Focal - Opponent	-	-	-	0.04 (0.05)	-0.03 (0.04)	-0.01 (0.03)

1045 **Table S2.** Estimated bivariate among-individual correlations for all traits included in our  
 1046 multivariate model (Table 3): activity, exploration, aggression, aggression elicited in others, and  
 1047 body weight. We print here the  $\chi^2$ -value (associated degrees of freedom = 1 in all cases) and  
 1048 values of  $P$  associated with likelihood ratio tests (LRTs) detailed in the main text (Methods).

Traits	r (SE)	$\chi^2_1$	$P$
activity - exploration	0.81 (0.08)	18.2	<b>&lt;0.001</b>
activity - aggression	0.39 (0.18)	3.88	<b>&lt;0.05</b>
activity - aggression elicited	0.16 (0.20)	0.63	0.426
exploration - aggression	0.39 (0.17)	230.5	<b>&lt;0.001</b>
exploration - aggression elicited	0.45 (0.17)	5.54	<b>&lt;0.05</b>
aggression - aggression elicited	0.09 (0.22)	0.14	0.700
weight - activity	0.17 (0.15)	225.4	<b>&lt;0.001</b>
weight - exploration	0.22 (0.14)	2.27	0.132
weight - aggression	0.43 (0.16)	5.66	<b>&lt;0.05</b>
weight - aggression elicited	-0.31 (0.18)	2.58	0.110

1049 **Table S3.** Relative fit of nine candidate models (detailed in the main text) based on the Akaike's  
 1050 information criterion (AIC). We present each model's AIC-value relative to the model with the  
 1051 lowest AIC-value ( $\Delta$ AIC), its weight, and relative likelihood.

Model	$\Delta$ AIC	Akaike Weight	Relative LL
3 - Full domain behavioural syndrome	0	0.75	1
9 – Behavioural syndrome with aggressiveness elicited only affected by exploration	2.35	0.23	0.31
2 - Behavioural syndrome with aggressiveness elicited independent	8.91	0	0.01
8 – Behavioural syndrome with aggressiveness elicited only affected by activity	9.90	0	0
7 – Behavioural syndrome with aggressiveness elicited linked only to aggressiveness	10.78	0	0
4 – General activity syndrome with social behaviours independent	11.62	0	0
6 – Independent social behaviours and general activity syndromes	13.33	0	0
1 – Behavioural independence	48.94	0	0
5 – Social behaviours syndrome with general activity behaviours independent	50.65	0	0



1052 **Table S4.** Estimated among-individual (a) and residual within-individual (b) variances and covariances/correlations (with SE) for two non-  
 1053 social behaviors (activity and exploration), two social behaviors (aggressiveness and aggressiveness elicited in opponents), and weight.  
 1054 We present variances on the diagonal, and covariances (correlations) on the lower (upper) off-diagonals, respectively.  
 1055 Covariances/correlations printed in bold-face are significant based on likelihood ratio tests derived from the multivariate model detailed  
 1056 in the main text.

a.	Aggressiveness	Activity	Exploration	Weight	Aggressiveness elicited
Aggressiveness	<b>0.21</b> (0.07)	<b>0.39</b> (0.18)	<b>0.39</b> (0.17)	<b>0.43</b> (0.16)	0.09 (0.22)
Activity	<b>0.11</b> (0.06)	<b>0.35</b> (0.06)	<b>0.81</b> (0.08)	<b>0.17</b> (0.15)	0.16 (0.20)
Exploration	<b>0.13</b> (0.06)	<b>0.33</b> (0.08)	<b>0.49</b> (0.10)	0.22 (0.14)	<b>0.45</b> (0.17)
Weight	<b>0.17</b> (0.07)	<b>0.08</b> (0.08)	0.13 (0.09)	<b>0.70</b> (0.13)	-0.31 (0.18)
Aggressiveness elicited	0.02 (0.04)	0.04 (0.05)	<b>0.13</b> (0.06)	-0.11 (0.07)	<b>0.18</b> (0.06)

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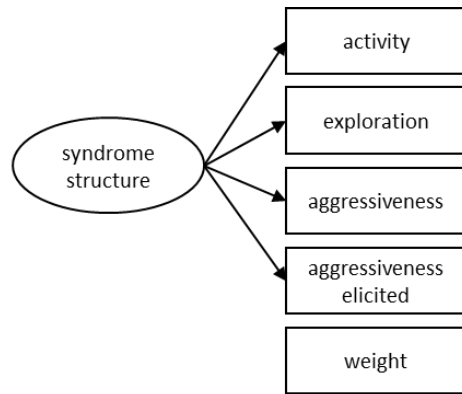
<b>b.</b>	<b>Aggressiveness</b>	<b>Activity</b>	<b>Exploration</b>	<b>Weight</b>
<b>Aggressiveness</b>	<b>0.61</b> (0.06)	-0.02 (0.07)	0.03 (0.07)	0.05 (0.06)
<b>Activity</b>	-0.01 (0.04)	<b>0.65</b> (0.06)	<b>0.17</b> (0.06)	0.06 (0.06)
<b>Exploration</b>	0.02 (0.04)	<b>0.10</b> (0.04)	<b>0.52</b> (0.04)	<b>0.13</b> (0.06)
<b>Weight</b>	0.03 (0.03)	0.03 (0.03)	<b>0.06</b> (0.03)	<b>0.35</b> (0.03)

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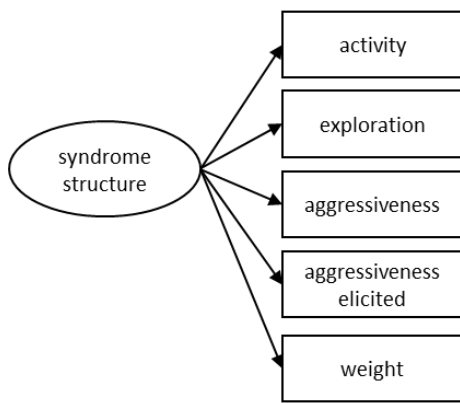
1058 **Table S5.** Relative fit of three a posteriori considered models (detailed in the main text)  
 1059 based on the Akaike's information criterion (AIC). We present each model's AIC-value  
 1060 relative to the model with the lowest AIC-value ( $\Delta$ AIC), its weight, and relative  
 1061 likelihood. Here we investigate how body weight is best incorporated into the best-  
 1062 fitting SEM presented in Table S3. For a visualization of each model's structure, see Fig.  
 1063 S1.

Model	$\Delta$ AIC	Akaike Weight	Relative LL
B – Weight as part of the full syndrome	0	0.73	1
A - Weight independent from full domain behavioural syndrome	1.98	0.27	0.37
C – Weight driving the behavioral syndrome	44.93	0	0

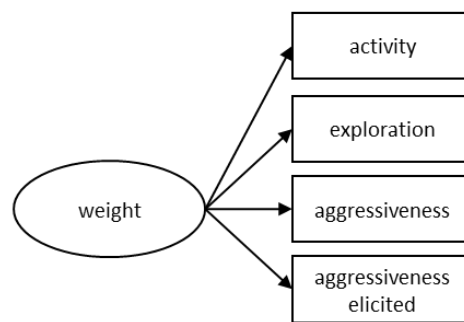
1064 **SUPPLEMENTARY FIGURES**



Model A



Model B



Model C

1065 **Figure S1.** Three a posteriori considered models. The relative fit of these models (Table  
 1066 S5) was considered to investigate how body weight is best incorporated into the best-  
 1067 fitting SEM presented in Table S3. Model A: weight is independent from the behavioral  
 1068 syndrome; model B: weight is part of the syndrome; model C: weight causes the  
 1069 behavioral syndrome.

