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Coastal proximity and mental health among urban adults in England: The moderating effect of household income

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ABSTRACT

After adjusting for covariates, self-reported general health in England is higher among populations living closer to the coast, and the association is strongest amongst more deprived groups. We explored whether similar findings were present for mental health using cross-sectional data for urban adults in the Health Survey for England (2008–2012, N \ge 25,963). For urban adults, living \le 1 km from the coast, in comparison to > 50 km, was associated with better mental health as measured by the GHQ12. Stratification by household income revealed this was only amongst the lowest-earning households, and extended to \le 5 km. Our findings support the contention that, for urban adults, coastal settings may help to reduce health inequalities in England.

1. Introduction

1.1. Overview

Poor mental health is among the leading causes of disability worldwide (World Health Organisation, 2018). In England, approximately one in six adults (17 %) surveyed were suffering symptoms of a common mental disorder (CMD), such as anxiety or depression (McManus et al., 2016). However, there is mounting evidence that exposure to natural environments is associated with various benefits for mental health and wellbeing (hereby referred to as 'mental health'; see reviews by Bratman et al., 2012; Frumkin et al., 2017; Hartig et al., 2014; Trostrup et al., 2019). Much of this work reports positive associations between green space and mental health, where measured or tested exposures include neighbourhood vegetation, green exercise, and residential proximity to green space (e.g. Barton and Pretty, 2010; Beyer et al., 2014; Cox et al., 2017b; de Vries et al., 2013; Gascon et al., 2015; McEachan et al., 2016). This may be particularly the case for those in urban areas where nature exposures can be limited (Cox et al., 2017a).

Concurrently, a smaller, yet growing, amount of research suggests that blue spaces (aquatic environments such as coasts, rivers, and lakes) are associated with a range of aspects related to improved mental health. These include: enhanced general health and wellbeing (reviewed by Gascon et al., 2017; see also Wheeler et al., 2012; White et al., 2013a; Völker et al., 2018; Volker and Kistemann, 2011; Wood et al., 2016); increased physical activity levels (White et al., 2014);

improved psychological restoration (White et al., 2010; White et al., 2013b); reduced psychological distress (Nutsford et al., 2016); and lower mortality rates (Crouse et al., 2018). A range of blue space exposures have been explored in these studies including area coverage, presence/absence, visibility and perceived and objective proximity.

There is also evidence that socioeconomic status may act as an effect-modifier, or moderator, of the nature-health relationship (see Hartig et al., 2014; Markevych et al., 2017; Mitchell et al., 2015). For example, several cross-sectional studies find that the association between natural environments and mental health is stronger within more deprived areas, or that health inequality gradients are lessened where green/blue space is more available (e.g. Wheeler et al., 2012; Maas et al., 2006; McEachan et al., 2016; Mitchell and Popham, 2008; Mitchell et al., 2015; van den Berg et al., 2016; Ward Thompson et al., 2012; however, see also Mitchell and Popham, 2007).

Again, however, most of this work has examined socioeconomic deprivation as a moderator of health regarding various measures of green space, with exposure to blue spaces receiving less empirical investigation (Markevych et al., 2017). Indeed, to the best of the authors' knowledge, only two studies have explicitly tested this relationship. First, Wheeler et al. (2012) found that the relationship between living closer to the coast in England and self-reported general health was strongest amongst communities within areas of higher socioeconomic deprivation. More recently, Crouse et al. (2018) examined the association between blue space and mortality in Canada, with results suggesting a similar pattern of effect-modification but lacking statistical power for some outcomes. A further study investigated the moderating

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effect of educational attainment, one aspect of socioeconomic status, on the relationship between both blue spaces and green spaces on various mental and physical health outcomes. They found a significant interaction between blue space and both health outcomes for those with the lowest educational attainment (de Vries et al., 2016).

Thus, although there have been encouraging findings, research examining the links between blue space and mental health remains limited (Gascon et al., 2017). Furthermore, despite growing health inequalities (Barr et al., 2015; Thomson et al., 2018), we currently have a poor understanding of how this relationship might vary between different levels of socioeconomic deprivation (Mitchell et al., 2015).

1.2. The current research

The aim of the present research was to investigate: (1) the association between mental health, as measured using two different indicators of common mental disorders (CMD), and the blue space exposure of coastal proximity, as used in Wheeler et al. (2012) for urban residents; and (2) variations in this association according to household income. The study therefore aimed to directly build on work by Wheeler et al. (2012) through focusing on self-reported mental health (as opposed to general health) as the dependent variable and at household level (instead of area level) deprivation in the form of household income, as the moderating variable. We used the Health Survey for England (HSE), a comprehensive, nationally representative survey which includes various measures of health, health behaviours and socio-demographics (Aresu et al, 2009, 2010, 2011; Boniface et al., 2012; Bridges et al., 2013). Based on the literature previously introduced, we hypothesised that: (a) CMD likelihood would decrease as coastal proximity increased; and (b) this association would be stronger amongst lower income households.

2. Methods

2.1. Sample

Secondary cross-sectional data were utilised from the HSE for English adults for the years 2008–2012 (pooled; aged 16+; adults N = 45,063). All inhabitants of selected households are eligible for interview and full sampling details can be found in (Aresu et al, 2009, 2010, 2011; Boniface et al., 2012; Bridges et al., 2013). Trained interviewers ask respondents a set of core questions related to their health, lifestyle, and background, with additional sections which vary each year.

The GHQ12 was not included in 2011 (n adults 2008-2010 and 2012 = 36,453), or the EQ5D in 2009 (n adults 2008 and 2010-2012 = 40,418), reducing the available samples for our analyses. Both the prevalence of mental health disorders and the relationships between natural environments and health have been found to vary by urbanity (see Alcock et al., 2015; Maas et al., 2006; Mitchell and Popham, 2007; Peen et al., 2010; Wheeler et al., 2012; Wood et al., 2016). Further, access to health services (Bauer et al., 2018; Chukwusa et al., 2019) and characteristics of natural environments are very different between urban and rural areas. We therefore focused only on urban residents, which are those individuals categorised by the trained interviewers as living in an 'Urban' setting, as opposed to 'Rural or isolated dwellings' or 'Town and Fringe' (Bibby and Brindley, 2013). Available urban adults sample sizes were 28,662 (GHQ12) and 31,906 (EQ5D). For respective analyses, we excluded those with missing responses for GHQ12 and the anxious/depression dimension of the EQ5D, therefore the full samples were 26,099 and 28,885 for GHQ12 and anxious/depression respectively.

The richness of the dataset enabled the inclusion of a range of potential confounding factors which may also relate to mental health, including: income, age, sex and the presence of limiting longstanding illnesses. We also included the health risk factors smoking status and body mass index (BMI), as these have received limited attention in previous studies exploring environment-health relationships (Mitchell, 2013).

We categorised responses of "Item not applicable", "No answer/ refused" and "Don't know" as missing and calculated the sample sizes. Missing data categories were excluded where there were < 20 respondents in a category (see Supplemental Table 2). This led to final analysis sample sizes of 25,963 (GHQ12) and 28,723 (EQ5D Anxiety/ depression).

2.2. Coastal proximity

Following previous approaches (Wheeler et al., 2012; White et al., 2013a), coastal proximity was measured in terms of the Euclidean distance (km) from the population density weighted centroid of respondents' Lower-layer Super Output Area (LSOA, as at 2001 Census) to the nearest coastline. There are approximately 32,500 LSOAs in England, each with a mean area of 4 km² and containing an average population of around 1500 (Wheeler et al., 2012). Following Wheeler et al. (2012), we operationalised coastal proximity using five categories: (1) 0-1 km; (2) > 1-5 km; (3) > 5-20 km; (4) > 20-50 km; (5) > 50 km. As with previous studies (Wheeler et al., 2012; White et al., 2013a), we used > 50 km as the reference category to enable us to test if the likelihood of having a CMD decreases with proximity to the coast. This also allowed us to compare 'coastal' (i.e. < 50 km; as used in EU definitions e.g. defining coastal regions (Eurostat, 2018)) respondents with 'inland' (i.e. \geq 50 km) respondents (White et al., 2013b).

2.3. Self-reported mental health

Mental health was measured through two outcomes. The first was the 12-item version of the General Health Questionnaire (GHQ12; Goldberg et al., 1997), available 2008–2010 and 2012, a self-reported measure widely used by health practitioners and researchers to indicate the likelihood or 'caseness' of an individual having a high risk of a CMD. Following established recommendations for the GHQ12, results were dichotomised with scores of four or above widely considered predictive of a high risk of common mental health disorders such as anxiety or depression (Fryers et al., 2004; Katikireddi et al., 2012; Mann et al., 2011; Semlyen et al., 2016). The two outcome categories for this measure were therefore: high likelihood of a CMD (GHQ12 score \geq 4); and low likelihood of a CMD (GHQ12 score < 4).

The second outcome was the anxiety and depression dimension of the EQ-5D-3L (hereafter referred to as the EQ5D; EuroQol Research Foundation, 2018). The EQ5D is a standardised measure of health-related quality of life (EuroQol Research Foundation, 2018), incorporating five dimensions, which has been used by practitioners and researchers (EuroQol, 2018; Hulme et al., 2004; Park et al., 2011) and utilised in studies exploring environmental characteristics (de Oliveira et al., 2013; Kyttä et al., 2011). Although the intended use is as a composite scale, here we use a single dimension - anxiety and depression. This dimension has been associated with anxiety and/or depression measured using a diagnostic scale (Mini International Neuropsychiatric Interview (Supina et al., 2007)), found to align with the GHQ12 (Bohnke and Croudace, 2016) and used for the same purpose in other studies (Semlyen et al., 2016). It should be noted that it was not found to be responsive to changes in anxiety or depression for those clinically diagnosed (Crick et al., 2018) and therefore not necessarily a measure of clinical diagnoses of anxiety and depression. As with the GHQ12, it is a self-completed scale. There are three possible response options, with respondents reporting whether they are not anxious or depressed; moderately anxious or depressed; or extremely anxious or depressed at the time of completion. These responses were dichotomised into the categories 'Not anxious/depressed' and 'At least moderately anxious/depressed' (moderate or extreme anxiety/depression)

to account for the skewness in data and low sample sizes within the extremely anxious depressed category (n = 648; 2 % of total).

2.4. Area level controls

In line with previous research (Wheeler et al., 2012; White et al., 2017; Mitchell and Popham, 2007), we controlled for area level deprivation (English Index of Multiple Deprivation, Noble et al., 2007), as well as green space and freshwater coverage at LSOA level to explore the unique effect of coastal proximity. The English Index of Multiple Deprivation (IMD) consists of area measures of crime, employment, education, and income and has been found to be related to mental health (Bellis et al., 2012) and moderate the coastal-health relationship (Wheeler et al., 2012). Percentage greenspace coverage was based on the generalised land use database (GLUD; Department for Communities and Local Government, 2007) for LSOAs and incorporated all area level green spaces, not including private gardens. Percentage freshwater coverage of the LSOA was derived from the CEH Land Cover Map 2007 (Morton et al., 2011).

2.5. Household level controls

We also included household and individual level covariates which may also relate to mental health. Equivalised household income, which takes into account the number of household members, was used to assess household level deprivation (reference category = highest income quintile). Household income has been found to be related to a range of mental health disorders (Domenech-Abella et al., 2018; Kahn et al., 2000; Sareen et al., 2011). The upper and lower bounds of each quintile vary by year and are given in Supplemental Table 1. Car access was also included at the household level (ref = access).

2.6. Individual level controls

Individual level controls were based on confounders of mental health identified by similar research with large survey datasets (Wheeler et al., 2012; White et al., 2013a; White et al., 2013b; Beyer et al., 2014; Crouse et al., 2017; Stranges et al., 2014). These included: sex (reference = female), age (reference = 16–34 years old), highest qualification level (reference = none/foreign/other), economic status (reference = in employment/student), relationship status (reference = single), year (reference = 2008), presence of limiting long-standing illnesses (reference = no limiting longstanding illness), cigarette smoking status (reference = never smoked cigarettes at all), and weight (body mass index; BMI; reference = normal weight).

2.7. Data linkage

Standard licence versions of HSE data only include large area geographical identifiers to preserve anonymity. In order to allocate higher resolution measures of coastal proximity, green space and freshwater, these three variables at LSOA level were supplied by the authors to the data providers (NatCen Social Research) and linked anonymously to HSE data under agreement from the NHS Health and Social Care Information Centre (now NHS Digital). To prevent identification of any individual LSOA of residence, the three environmental variables were constrained to relatively coarse categories; and LSOA and regional identifiers were removed from the linked data and returned to the authors.

2.8. Analyses

Data were analysed using the "survey" package (version 3.34; Lumley, 2018) in R Studio Version 3.4.2. Generalised linear models (GLM) using a quasi-binomial error structure (appropriate when analysing complex survey data (Lumley, 2018)) and household clusters, to account for multiple respondents within households and provide robust standard errors, were used to identify correlations between coastal proximity and mental health. We were not able to include clustering by LSOA as this had been removed by the data providers for anonymity. The data were weighted using the interview weights provided in the dataset to account for selection, non-response and population biases (Aresu et al, 2009, 2010, 2011; Boniface et al., 2012; Bridges et al., 2013). We calculated the odds ratios (OR) and 95% confidence intervals (CI) of participants having either a high likelihood of a CMD (GHQ12 \geq 4) or of reporting a status of at least moderately anxious/ depressed for this dimension of the EOSD.

We present unadjusted models (nature exposures only) and fullyadjusted models to examine how coastal proximity was associated with mental health before and after adding the controls. A sensitivity analysis was also carried out with > 20 km as a reference category. We then stratified our analysis by household income, whereby we analysed the relationships between coastal proximity and mental health using fully-adjusted (unweighted) GLMs for each household income quintile. This enabled us to observe variations in the relationship between coastal proximity and mental health by household income. We had an *a priori* prediction that the effects would be strongest in the lowest income quintiles, however, we also carry out analyses interacting coastal proximity and household income.

3. Results

3.1. Full model/sample results

Table 1 presents descriptive statistics of the un-stratified mental health models. In the GHQ12 model sample, the proportion of people with a high risk of a common mental disorder (CMD) closely resembled previous national averages (McManus et al., 2016), with approximately 15 % of participants reporting a high likelihood of suffering from a CMD. In comparison, CMD prevalence was slightly higher in the EQ5D model, likely due to the different method of measurement, with approximately 22 % of respondents reporting at least moderate anxiety or depression. CMD prevalence was also greater amongst more deprived areas and lower earning households (Table 1). For the years 2008, 2010 and 2012 where both the GHQ12 and anxiety and depression were present, the correlation was 0.50 (kendall's τ , p < 0.001).

Table 2 displays the unadjusted and adjusted odds ratios (OR) with 95% confidence intervals (CI) of respondents having a high risk of CMD for the full model samples (un-stratified) of both outcomes. Respondents were less likely to report an at risk GHQ12 score of \geq 4 if they lived up to 1 km of the coast compared to > 50 km (OR_{adj} = 0.78, 95 % CI = 0.65 – 0.95).

No significant (p < 0.05) associations were found between coastal proximity and CMD likelihood for either the GHQ12 outcome or anxiety/depression EQ5D dimension in the unadjusted models. Similarly, there were no significant associations between coastal proximity and the anxious/depression dimension of the EQ5D in the adjusted model.

Respondents living in areas of 80–100 % greenspace were less likely to report at risk scores of the GHQ12 and being at least moderately anxious or depressed in the unadjusted models (and 60–80 % with the EQ5D dimension). However, neither of these associations held in the adjusted models.

Freshwater coverage of > 5–100 % was found to be related to the anxious/depression dimension of the EQ5D in both the unadjusted and adjusted models ($OR_{adj} = 0.78$, 95 % CI = 0.63 – 0.96). However, it was not related to GHQ12 in either the unadjusted or adjusted model.

The results from our sensitivity analysis with > 20 km as a reference category were similar, giving us confidence in our results (Supplementary Table 3). As with a reference category of > 50 km, we find significant associations between living ≤ 1 km from the coast and the GHQ12 (≤ 1 km vs. > 20 km OR_{adj} = 0.79, 95 % CI = 0.66–0.94).

| su leter | J | | | | | | | | EQ5D | | | |
|--------------------------------------|--------------------------|-------------------|-------------------------------------|----------------------------|--|----------------------------|-------------------|-------------------|--------------------------|-------------------|---|----------------|
| ol Ne | Full model sample | le | Low risk of CMD^{a} (score | D ^a (score < 4) | High risk of CMD^{a} (score ≥ 4) | ID ^a (score ≥4) | Full model sample | ple | Not anxious or depressed | depressed | At least moderately anxious or depressed | ely anxious or |
| tal Ne | Unweighted N | Weighted % age | Unweighted N | Weighted % age | Unweighted N | Weighted % age | Unweighted N | Weighted % age | Unweighted N | Weighted % age | Unweighted N | Weighted %age |
| LSOA ^b level variables | 25963 | | 21984 | 84.88 | 3979 | 15.12 | 28723 | | 22275 | 78.38 | 6448 | 21.62 |
| 0-1 km | 1532 | 5.75 | 1315 | 86.03 | 217 | 13.97 | 1826 | 6.17 | 1413 | 78.01 | 413 | 21.99 |
| > 1–5 km | 3202 | 11.68 | 2729 | 85.70 | 473 | 14.30 | 3394 | 11.27 | 2635 | 78.88 | 759 | 21.12 |
| > 5–20 km | 3781 | 13.46 | 3186 | 84.77 | 595 | 15.23 | 4400 | 13.93 | 3372 | 77.70 | 1028 | 22.30 |
| > 20-50 km | 7252 | 30.10 | 6171 | 85.07 | 1081 | 14.93 | 7921 | 29.71 | 6216 | 79.13 | 1705 | 20.87 |
| > 50 km (ref) | 10196 | 39.01 | 8583 | 84.36 | 1613 | 15.64 | 11182 | 38.92 | 8639 | 77.97 | 2543 | 22.03 |
| 100 % | 712 | 2.74 | 602 | 84 83 | 110 | 15.17 | 776 | 2.67 | 627 | 81 89 | 149 | 18 11 |
| > 1-5 % | 1607 | 6.09 | 1386 | 86.81 | 221 | 13 19 | 1790 | 610 | 1393 | 78.19 | 397 | 21.81 |
| > 0-1 % | 1459 | 5.42 | 1244 | 85.38 | 215 | 14.62 | 1658 | 5.63 | 1278 | 77.81 | 380 | 22.19 |
| 0 % (ref) | 22185 | 85.75 | 18752 | 84.72 | 3433 | 15.28 | 24499 | 85.59 | 18977 | 78.32 | 5522 | 21.68 |
| Greenspace coverage | | | | | | | | | | | | |
| 80-100 % | 1697 | 6.03 | 1478 | 87.33 | 219 | 12.67 | 1897 | 6.15 | 1516 | 80.20 | 381 | 19.80 |
| 60 - < 80% | 3239 | 11.84 | 2788 | 86.23 | 451 | 13.77 | 3539 | 11.76 | 2817 | 80.21 | 722 | 19.79 |
| 40 - < 60% | 4570 | 17.04 | 3842 | 84.41 | 728 | 15.59 | 5108 | 17.15 | 3912 | 77.26 | 1196 | 22.74 |
| 20 - < 40% | 7630 | 29.15 | 6398 | 84.11 | 1232 | 15.89 | 8545 | 29.45 | 6547 | 77.97 | 1998 | 22.03 |
| 0 - < 20 % (ref) IMD ^c | 8827 | 35.93 | 7478 | 84.88 | 1349 | 15.12 | 9634 | 35.49 | 7483 | 78.34 | 2151 | 21.66 |
| Most deprived | 5549 | 21.65 | 4410 | 80.26 | 1139 | 19.74 | 6219 | 21.97 | 4495 | 73.83 | 1724 | 26.17 |
| 2nd most deprived | 5540 | 21.90 | 4640 | 84.23 | 006 | 15.77 | 6133 | 21.82 | 4674 | 77.70 | 1459 | 22.30 |
| Medium deprived | 5051 | 19.79 | 4298 | 85.01 | 753 | 14.99 | 5566 | 19.67 | 4343 | 78.60 | 1223 | 21.40 |
| 2nd least deprived | 4573 | 17.24 | 3986 | 87.30 | 587 | 12.70 | 5118 | 17.53 | 4086 | 80.53 | 1032 | 19.47 |
| Least deprived (ref) | 5250 | 19.43 | 4650 | 88.49 | 600 | 11.51 | 5687 | 19.00 | 4677 | 82.21 | 1010 | 17.79 |
| Household level variables | | | | | | | | | | | | |
| יבווסות וווכסוווב לתווות | | | 0,00 | | | | ļ | | | | | |
| Lowest | | | 2969 | 76.12 | 953 707 | 23.88 | 4347 | 14.86 15.27 | 2890 2501 | 68.07 76.07 | 1457 | 31.93 |
| Second lowest Middle | 41.CI 2.14 7167 71574 | | 5405 2576 | 05.21 06 77 | 707 | 10./3 | 4002 | 15.61 | 1065 | c0.0/ | 1011 | C6.67 |
| Muuue Sacond hichaet | | | 20100 | 00.10 | 517 517 | 11 00 | 1050 | 17 97 | 1100 | 00.00 00 15 | 17.0 | 17 05 |
| Jecona nagrese Missing data | 4923 19.76 | | 4159 | 84.38 | 764 | 15.62 | 5478 | 19 97 | 4191 | 77 40 | 1 287 | 22.60 |
| Highest (ref) | | | 3898 | 89.85 | 447 | 10.15 | 4743 | 16.80 | 4034 | 85.34 | 709 | 14.66 |
| Car access | | | | | - | | 2 | | - | | | |
| No | 5503 20.79 | | 4277 | 78.04 | 1226 | 21.96 | 6165 | 20.91 | 4173 | 68.92 | 1992 | 31.08 |
| Yes (ref) | _ | | 17707 | 86.68 | 2753 | 13.32 | 22558 | 20.09 | 18102 | 80.88 | 4456 | 19.12 |
| Individual level variables | | | | | | | | | | | | |
| Age categories | 010 | | 00 10 | 67 60 | 107 | 20.21 | 1076 | 0.01 | 0110 | 10 01 | 166 | 26 26 |
| | | | 0017 | 00.00 10 20 | 1016 | 10.07 | 5008 | 0.01 20 F6 | 0117 | 10.44 | 50/ | 20./0 |
| 35 - 74 35 - 54 | 0.44 25.00 0040 35.18 | | 7558 | 83.03 | 1010 | 15.07 16.07 | 80.30 10060 | 25.20 35.25 | 2012Z | 77 15 | 1904 2358 | 20.02 78 CC |
| 16 - 24 (raf) | | | 5010 | 85.40 | 1045 | 14.60 | 7759 | 22.1g | 6331 | 87.78 | 1491 | 17 70 |
| Highest qualification | | | 0100 | | | 00.11 | 10 | 01.00 | 1000 | | 1711 | |
| Higher ed/Degree | 8203 32.73 | | 7130 | 86.83 | 1073 | 13.17 | 9294 | 33.27 | 7596 | 81.89 | 1698 | 18.11 |
| NVQ3/A level | | | 3364 | 85.64 | 580 | 14.36 | 4429 | 17.19 | 3609 | 82.27 | 820 | 17.73 |
| NVQ1/NVQ2/GCSE | 6989 26.97 | | 5888 | 84.66 | 1101 | 15.34 | 7650 | 26.81 | 5858 | 77.46 | 1792 | 22.54 |
| Other/none (ref) | 6827 23.33 | | 5602 | 81.86 | 1225 | 18.14 | 7350 | 22.73 | 5212 | 71.38 | 2138 | 28.62 |
| Working status | | | | | | | | | | | | |

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Table 1 Descriptive statistic

| (continued) | |
|-------------|--|
| able 1 | |

| • | | | | | | | | | | | | |
|--|-------------------|-------------------|--|----------------------------|--|----------------------------|-------------------|-------------------|--------------------------|-------------------|---|---------------|
| | Full model sample | a | Low risk of CMD^{a} (score |) ^a (score < 4) | High risk of CMD^{a} (score ≥ 4) | MD^{a} (score ≥ 4) | Full model sample | ple | Not anxious or depressed | depressed | At least moderately anxious or depressed | ly anxious or |
| . – | Unweighted N | Weighted % age | Unweighted N | Weighted % age | Unweighted N | Weighted % age | Unweighted N | Weighted % age | Unweighted N | Weighted % age | Unweighted N | Weighted %age |
| ILO unemployed ^d 803 | 3.54 | | 576 | 72.26 | 227 | 27.74 | 923 | 3.58 | 601 | 66.91 | 322 | 33.09 |
| ctive | 30.82 | | 7634 | 79.43 | 1898 | 20.57 | 10430 | 30.64 | 7248 | 68.96 | 3182 | 31.04 |
| In work/student (ref) 15628 | | | 13774 | 88.12 | 1854 | 11.88 | 17370 | 65.77 | 14426 | 83.40 | 2944 | 16.60 |
| | | | | | | | | | | | | |
| | | | 9866 | 86.90 | 1511 | 13.10 | 12668 | 48.76 | 10275 | 81.54 | 2393 | 18.46 |
| Female (ref) 14466 Relationshin status | 51.25 | | 11998 | 82.96 | 2468 | 17.04 | 16055 | 51.24 | 12000 | 75.37 | 4055 | 24.63 |
| In a relationship 16207 | 61.27 | | 14132 | 87.25 | 2075 | 12.75 | 17870 | 61.25 | 14427 | 81.06 | 3443 | 18.94 |
| d./ | | | 3530 | 78.33 | 954 | 21.67 | 4980 | 14.58 | 3394 | 67.89 | 1586 | 32.11 |
| divorced | | | | | | | | | | | | |
| Single (ref) 5272 Limiting illness presence | 24.13 | | 4322 | 82.84 | 950 | 17.16 | 5873 | 24.17 | 4454 | 77.91 | 1419 | 22.09 |
| Limiting illness | 22.13 | | 4341 | 67.77 | 2024 | 32.23 | 7046 | 22.09 | 3981 | 56.64 | 3065 | 43.36 |
| Non-limiting longstanding illness | | | 4498 | 89.38 | 514 | 10.62 | 5442 | 17.95 | 4355 | 79.86 | 1087 | 20.14 |
| No longstanding illness (ref) BMI- ^e | | | 13145 | 89.86 | 1441 | 10.14 | 16235 | 59.95 | 13939 | 85.95 | 2296 | 14.05 |
| <i>Obese</i> 5869 | 21.71 | | 4837 | 82.65 | 1032 | 17.35 | 6517 | 21.87 | 4824 | 75.10 | 1693 | 24.90 |
| Overweight 8412 | 31.93 | | 7316 | 86.95 | 1096 | 13.05 | 9237 | 31.65 | 7328 | 79.84 | 1909 | 20.16 |
| Underweight 373 | 1.70 | | 294 | 80.26 | 79 | 19.74 | 399 | 1.63 | 298 | 77.55 | 101 | 22.45 |
| | 12.37 | | 2736 | 81.61 | 650 | 18.39 | 3835 | 12.76 | 2846 | 75.17 | 989 | 24.83 |
| Normal weight (ref) 7923 Smoking status | 32.30 | | 6801 | 85.84 | 1122 | 14.16 | 8735 | 32.10 | 6269 | 80.49 | 1756 | 19.51 |
| Current smoker 5587 | 22.19 | | 4398 | 79.16 | 1189 | 20.84 | 6151 | 22.02 | 4279 | 70.53 | 1872 | 29.47 |
| | | | 6915 | 85.46 | 1184 | 14.54 | 8952 | 29.20 | 6981 | 78.68 | 1971 | 21.32 |
| (ref) | | | 10671 | | 1606 | 12.85 | 13620 | 48.79 | 11015 | 81.74 | 2605 | 18.26 |
| Year | | | | | | | | | | | | |
| 2012 5591 | 21.53 | | 4702 | 84.31 | 889 | 15.69 | 5696 | 19.92 | 4500 | 79.80 | 1196 | 20.20 |
| | I | | I | I | I | I | 5891 | 20.35 | 4277 | 73.78 | 1614 | 26.22 |
| | | | 4956 | 84.71 | 901 | 15.29 | 5865 | 20.76 | 4473 | 76.94 | 1392 | 23.06 |
| 2009 3308 | | | 2728 | 82.68 | 580 | 17.32 | I | I | 1 | 1 | 1 | 1 |
| 2008 (ref 11207 | 42.68 | | 9598 | 85.94 | 1609 | 14.06 | 11271 | 38.96 | 9025 | 80.82 | 2246 | 19.18 |

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Table 2

Unadjusted and adjusted regression models predicting the likelihood of respondents having poor metal health as assessed using the GHQ12 and anxiety/depression component of the EQ5D. Significant results are highlighted in bold type.

| | GHQ12 (≥ | - | | | | | | | | | | |
|----------------------------------|------------|-------------|-------|----------|-------------|---------|----------|-------------|-------|----------|-------------|-------|
| | Unadjustee | d | | Adjusted | | | Unadjust | ted | | Adjusted | | |
| erm | OR | 95 % CI | р | OR | 95 % CI | р | OR | 95 % CI | р | OR | 95 % CI | р |
| SOA level variables ^a | | | | | | | | | | | | |
| oastal proximity | | | | | | | | | | | | |
| 0–1 km | 0.87 | 0.73-1.04 | 0.121 | 0.78 | 0.65-0.95 | 0.011 | 0.99 | 0.85-1.14 | 0.836 | 0.90 | 0.78-1.05 | 0.195 |
| > 1–5 km | 0.90 | 0.80-1.02 | 0.091 | 0.90 | 0.79-1.02 | 0.108 | 0.95 | 0.85-1.05 | 0.330 | 0.93 | 0.83-1.04 | 0.188 |
| > 5–20 km | 0.97 | 0.86-1.09 | 0.592 | 0.99 | 0.87-1.12 | 0.845 | 1.02 | 0.93-1.12 | 0.709 | 1.01 | 0.92-1.11 | 0.818 |
| > 20–50 km | 0.94 | 0.86-1.04 | 0.229 | 0.97 | 0.88-1.07 | 0.507 | 0.93 | 0.86-1.01 | 0.083 | 0.97 | 0.89-1.05 | 0.451 |
| > 50 km (ref) | 0.94 | 0.00-1.04 | 0.229 | 0.97 | 0.00-1.07 | 0.307 | 0.95 | 0.00-1.01 | 0.085 | 0.97 | 0.89-1.05 | 0.431 |
| reshwater coverage | | | | | | | | | | | | |
| > 5–100 % | 0.99 | 0.78 - 1.26 | 0.949 | 0.95 | 0.74-1.21 | 0.656 | 0.79 | 0.65-0.97 | 0.022 | 0.78 | 0.63-0.96 | 0.020 |
| > 1–5 % | 0.87 | 0.74-1.03 | 0.098 | 0.85 | 0.72 - 1.02 | 0.074 | 1.03 | 0.89–1.19 | 0.709 | 1.05 | 0.91 - 1.21 | 0.541 |
| > 0-1 % | 1.03 | 0.86 - 1.24 | 0.758 | 1.03 | 0.85 - 1.26 | 0.768 | 1.09 | 0.95 - 1.25 | 0.239 | 1.12 | 0.96-1.31 | 0.141 |
| 0 % (ref) | | | | | | | | | | | | |
| reenspace coverage | | | | | | | | | | | | |
| 80–100 % | 0.82 | 0.69-0.97 | 0.023 | 0.90 | 0.75-1.08 | 0.267 | 0.86 | 0.75-0.99 | 0.043 | 0.87 | 0.75-1.02 | 0.082 |
| 60 - < 80 % | 0.90 | 0.79-1.02 | 0.103 | 0.97 | 0.85-1.11 | 0.643 | 0.88 | 0.79-0.98 | 0.019 | 0.91 | 0.82-1.02 | 0.122 |
| 40 - < 60 % | 1.04 | 0.93-1.16 | 0.504 | 1.02 | 0.91-1.15 | 0.720 | 1.06 | 0.97-1.16 | 0.218 | 1.02 | 0.92-1.12 | 0.763 |
| | | | | | | | | | 0.213 | 0.99 | 0.92-1.12 | 0.781 |
| 20 - < 40% | 1.06 | 0.96–1.16 | 0.263 | 1.05 | 0.95–1.16 | 0.387 | 1.01 | 0.94–1.10 | 0.741 | 0.99 | 0.91-1.07 | 0.781 |
| 0 - < 20 % (ref) | | | | | | | | | | | | |
| ИD ^ь | | | | | | | | | | | | |
| Most deprived | | | | 1.21 | 1.05-1.38 | 0.008 | | | | 1.04 | 0.92 - 1.17 | 0.518 |
| 2nd most deprived | | | | 1.08 | 0.94-1.24 | 0.257 | | | | 0.99 | 0.89-1.11 | 0.922 |
| Medium deprived | | | | 1.17 | 1.03-1.34 | 0.020 | | | | 1.07 | 0.96-1.20 | 0.230 |
| 2nd least deprived | | | | 1.05 | 0.91-1.21 | 0.509 | | | | 1.00 | 0.90-1.13 | 0.939 |
| Least deprived (ref) | | | | 1.00 | 0.91 1.21 | 0.009 | | | | 1.00 | 0.50 1.10 | 0.909 |
| | | | | | | | | | | | | |
| Iousehold level variables | | | | | | | | | | | | |
| lousehold income quintile | | | | | | | | | | | | |
| Lowest | | | | 1.40 | 1.19-1.64 | < 0.001 | | | | 1.37 | 1.19-1.56 | < 0.0 |
| Second lowest | | | | 1.24 | 1.06-1.44 | 0.007 | | | | 1.19 | 1.05-1.36 | 0.007 |
| Middle | | | | 1.15 | 0.99-1.33 | 0.074 | | | | 1.12 | 0.99 - 1.27 | 0.082 |
| Second highest | | | | 1.13 | 0.97-1.31 | 0.119 | | | | 1.20 | 1.06-1.36 | 0.004 |
| Missing data | | | | 1.21 | 1.04-1.40 | 0.012 | | | | 1.22 | 1.08-1.39 | 0.002 |
| Highest (ref) | | | | | 1101 1110 | 0.012 | | | | | 1100 1107 | 0.001 |
| | | | | | | | | | | | | |
| ar access | | | | | 1 00 1 07 | 0.010 | | | | 1.04 | | |
| No | | | | 1.15 | 1.03-1.27 | 0.010 | | | | 1.24 | 1.14-1.36 | < 0.0 |
| Yes (ref) | | | | | | | | | | | | |
| ndividual level variables | | | | | | | | | | | | |
| ge categories | | | | | | | | | | | | |
| 75+ | | | | 0.41 | 0.34-0.50 | < 0.001 | | | | 0.55 | 0.46-0.65 | < 0.0 |
| 55 - 74 | | | | 0.54 | 0.47-0.63 | < 0.001 | | | | 0.78 | 0.69-0.88 | < 0.0 |
| 35 - 54 | | | | 1.00 | 0.89-1.12 | 0.984 | | | | 1.23 | 1.12-1.35 | < 0.0 |
| 16–34 (ref) | | | | 1.00 | 0.05 1.12 | 0.901 | | | | 1.20 | 1.12 1.00 | ~ 0.0 |
| | | | | | | | | | | | | |
| lighest qualification | | | | | | | | | | | | |
| Higher ed/Degree | | | | 1.19 | 1.05-1.34 | 0.005 | | | | 0.97 | 0.88 - 1.08 | 0.601 |
| NVQ3/A level | | | | 1.12 | 0.98-1.29 | 0.105 | | | | 0.85 | 0.76-0.96 | 0.007 |
| NVQ1/NVQ2/GCSE | | | | 1.03 | 0.92-1.15 | 0.649 | | | | 0.95 | 0.86-1.04 | 0.251 |
| Other/none (ref) | | | | | | | | | | | | |
| Vorking status | | | | | | | | | | | | |
| ILO unemployed ^c | | | | 1.97 | 1.62-2.39 | < 0.001 | | | | 1.72 | 1.46-2.03 | < 0.0 |
| Retired/other inactive | | | | 1.47 | 1.32-1.64 | < 0.001 | | | | | 1.42-1.70 | < 0.0 |
| | | | | 1.4/ | 1.52-1.04 | < 0.001 | | | | 1.55 | 1.42-1.70 | < 0.0 |
| In work/student (ref) | | | | | | | | | | | | |
| ex | | | | | | | | | | | | |
| Male | | | | 0.79 | 0.73-0.86 | < 0.001 | | | | 0.75 | 0.70-0.80 | < 0.0 |
| Female (ref) | | | | | | | | | | | | |
| elationship status | | | | | | | | | | | | |
| In a relationship | | | | 0.80 | 0.71-0.89 | < 0.001 | | | | 0.78 | 0.71-0.86 | < 0.0 |
| Widow/separated./divorce | d | | | 1.12 | 0.98-1.28 | 0.098 | | | | 1.07 | 0.95-1.20 | 0.282 |
| Single (ref) | | | | | | | | | | | | |
| miting illness presence | | | | | | | | | | | | |
| | | | | 4.99 | 2 00 4 60 | < 0.001 | | | | 4.15 | 204 250 | - 0 - |
| Limiting illness | -11 | | | 4.28 | 3.90-4.69 | < 0.001 | | | | 4.15 | 3.84-3.50 | < 0.0 |
| Non-limiting longstanding | | | | 1.19 | 1.06-1.34 | 0.003 | | | | 1.62 | 1.48-1.77 | < 0.0 |
| No longstanding illness (re | Ð | | | | | | | | | | | |
| MI ^d | | | | | | | | | | | | |
| Obese | | | | 1.08 | 0.97-1.21 | 0.140 | | | | 1.07 | 0.98-1.17 | 0.154 |
| Overweight | | | | 0.95 | 0.86-1.06 | 0.373 | | | | 1.03 | 0.94-1.12 | 0.511 |
| 0 | | | | | | | | | | | | |
| Underweight | | | | 1.23 | 0.92-1.65 | 0.161 | | | | 1.04 | 0.79-1.36 | 0.794 |
| Missing data | | | | 1.12 | 0.99–1.26 | 0.071 | | | | 1.03 | 0.92–1.15 | 0.608 |
| Normal weight (ref) | | | | | | | | | | | | |
| noking status | | | | | | | | | | | | |
| - | | | | 1.40 | 1.26-1.54 | < 0.001 | | | | 1.55 | 1.43-1.69 | < 0.0 |
| Current smoker | | | | | | | | | | | | |

Table 2 (continued)

| | GHQ12 (\geq | 4) | | | | | EQ5D: at least moderately anxious/depressed | | | | | | |
|--|--|---------|---|--|------------------------|-------------------------|---|---------|---|---|------------------------|------------------|--|
| | Unadjusted | | | Adjusted | | | Unadjusted | 1 | | Adjusted | | | |
| term | OR | 95 % CI | р | OR | 95 % CI | р | OR | 95 % CI | р | OR | 95 % CI | р | |
| Used to smoke Never smoked (ref) | | | | 1.13 | 1.03–1.24 | 0.011 | | | | 1.09 | 1.01–1.17 | 0.033 | |
| Year 2012 2011 | | | | 1.17 | 1.06-1.30 | 0.003 | | | | 1.12 1.62 | 1.02–1.23 1.48–1.77 | 0.022 < 0.001 | |
| 2010 2009 | | | | 1.11 1.30 | 1.00–1.23 1.15–1.47 | 0.053 < 0.001 | | | | 1.32 | 1.21-1.45 | < 0.001 | |
| 2008 (ref) | | | | | | | | | | | | | |
| Intercept N Households AIC ^e Cox & Snell <i>pseudo-R</i> ² (%) | -1.68 25963 16592 22056.47 0.1 | | | -2.52 25963 16592 19951.43 8.1 | | | -1.25 28723 18419 29987.95 0.1 | | | -2.16 28723 18419 26890.79 10.5 | | | |

^a LSOA = Lower-layer Super Output Area; bIMD = Indices of Multiple Deprivation; cILO = International Labour Organisation; dBMI = Body Mass Index; eAIC = Akaike's Information Criterion.

3.2. Results stratified by household income

We find some significant interactions between coastal proximity and household income for both the GHQ12 measure and the anxiety/depression of the EQ5D (Supplemental Table 4).

Full results for each income quintile are presented in supplementary

materials (Supplemental Tables 5–9), with a summary of the key coastal proximity findings in Fig. 1. As can be seen, living near the coast (\leq 5 km) is associated with lower ORs (than living > 50 km) of poor mental health as measured by both the GHQ12 (0–1 km OR_{adj} = 0.58, 95 % CI = 0.39 – 0.87; > 1–5 km OR_{adj} = 0.76, 95 % CI = 0.59 – 0.98) and the anxiety/depression sub-scale of the EQ5D (0–1 km

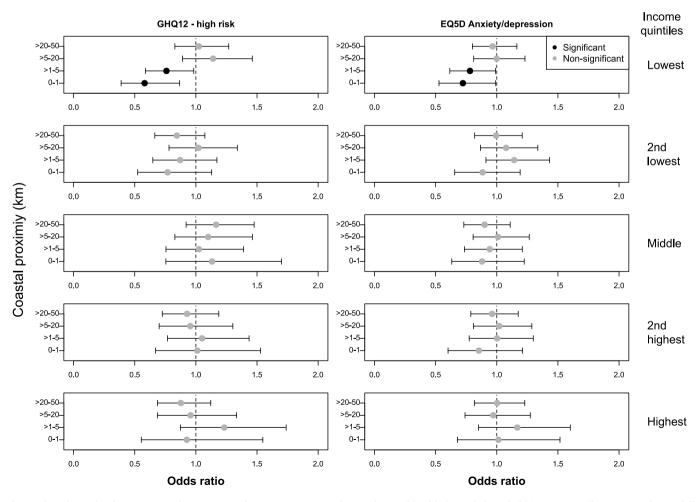


Fig. 1. The relationship between coastal proximity (reference category > 50 km) and mental health for each household income quintile. Note: results are fully adjusted; CMD likelihood presented as odds ratios with 95% confidence intervals. Full model results in Supplemental Tables 5–9.

 $OR_{adj}=0.72,\ 95\ \%\ CI=0.53\ -\ 0.99; > 1-5\ km\ OR_{adj}=0.78,\ 95\ \%\ CI=0.62\ -\ 0.99)$ for individuals in the lowest household income quintile only. There were no other significant associations between coastal proximity and mental health for those in the higher household income quintiles.

4. Discussion

In sum, we have explored the association between two measures of mental health and coastal proximity for urban English adults using four years of pooled data from the Health Survey for England. After adjusting for a range of relevant covariates, those living 0-1 km from the coast had significantly lower odds of being at high risk of a CMD, as measured by the GHQ12 and compared to those living further than 50 km. Coastal proximity was not found to be related to the anxiety/ depression EQ5D dimension.

As predicted, income quintile was a strong predictor of mental health outcomes, and other socioeconomic factors (e.g. employment, relationship and smoking status) were also largely consistent with earlier work (Katikireddi et al., 2016; Stranges et al., 2014). However, we find BMI not to be significantly related to mental health contrasting with research by Stranges et al. (2014) using the HSE. Conversely, we find that those who used to smoke were more likely to have poorer mental health whereas this was not found by Stranges et al. (2014). We also find not having access to a car was significantly related with worse mental health whilst this was not found in earlier HSE years (Riva et al., 2011).

Stratifying by household income revealed that the relationship between coastal proximity and mental health outcomes was present only for those with the lowest household incomes and extended to < 5 km. Specifically, the results imply that people living in urban areas in the lowest household income quintile are less likely to suffer from a common mental disorder (CMD) such as anxiety or depression if they live within 5 km of the coast, compared to those living in urban areas further inland (> 50 km). In particular, living within 1 km of the coast is associated with the strongest reductions in CMD likelihood for people from the most economically deprived households. Respondents from this category reported symptoms consistent with a CMD according to the GHQ12 measure with odds that were 40 % less than those living further than 50 km. This is a greater reduction in comparison to being in a relationship (*vs.* single OR_{adi} = 0.78, 95 % CI = 0.63–0.98).

These findings add to the growing evidence base linking blue spaces, particularly coastal environments, with better health and wellbeing (White et al., 2010; Wheeler et al., 2012; White et al., 2013a; Crouse et al., 2018; Gascon et al., 2015; Gascon et al., 2017; Nutsford et al., 2016; Volker and Kistemann, 2011). This study also highlights the potentially beneficial link between coastal proximity and common mental disorders, which have been highlighted as growing issues in countries such as England (McManus et al., 2016). Given that increasingly many people live by and visit the coast in many countries, and even more of them reside in cities, such research is vital for environmental and social policy (Elliott et al., 2018; Pelling and Blackburn, 2014).

This research also supports previous work which suggests that the positive relationship between living in more natural environments and mental health is stronger within more socioeconomically deprived groups (e.g. Wheeler et al., 2012; Maas et al., 2009; Maas et al., 2006; McEachan et al., 2016; Mitchell and Popham, 2008). It also extends prior research that investigated the interaction with area level deprivation (Wheeler et al., 2012), by demonstrating that household income moderates the association between coastal proximity and health, in this case specifically mental health. This suggests that access to the natural environment may, at least partly, offset the adverse health and wellbeing outcomes associated with low incomes. Indeed, recent work by Elliott et al. (2018) finds that recreational visits to the English coast, particularly walking, are more likely to be made by people from some

lower socioeconomic backgrounds as compared to other natural environments. Subsequently, ensuring coastal environments are accessible to more socioeconomically deprived communities could therefore help to reduce health inequalities (Elliott et al., 2018).

Although not established in this study, it is plausible that there is a causal relationship between coastal living and mental health. Indeed, it could be that exposure to coastal environments improves mental health through a range of potential mechanisms in the same way as has been proposed for green space, such as through reduced stress, improved air quality and immune functioning, and increased opportunities for social contact and physical activity (Hartig et al., 2014; Markevych et al., 2017). In support of this, de Bell et al. (2017) sought to test whether the same mechanisms that have been proposed to explain the relationship between green space and health also applied in blue space visits. Most people identified psychological benefits or social interactions as the most important perceived benefit from their most recent blue space visit. Similarly, higher levels of blue space visibility were associated with lower levels of psychological distress in Wellington, New Zealand, whilst green space visibility was not found to be related (Nutsford et al., 2016) and, in Ireland, a sea view was found to be related to lower depression scores (Dempsey et al., 2018). Earlier work by (Bauman et al., 1999) also suggests that living by the coast is associated with increased opportunities for physical activity.

More recently, White et al. (2014) found that people in England who lived closer to the coast were more likely to visit the coast and, subsequently, achieve their recommended weekly physical activity levels. Combined with the finding that approximately 271 million recreational visits are made each year to coastal environments in England (Elliott et al., 2018), this suggests that the mental health of English coastal urban dwellers (who are more likely to visit the coast) is better than those in urban areas inland because of certain salutary mechanisms, such as physical activity.

There were several unexpected findings in our research. For instance, in contrast to previous research (e.g. de Bell et al., 2017; MacKerron and Mourato, 2013; Völker et al., 2018), we found that whilst living in closer proximity to coastal environments was significantly linked to improved mental health outcomes, living in areas with more freshwater coverage was not related overall according to the GHQ12 measure. However, freshwater coverage was related to the anxiety/depression EQ5D dimension. Freshwater coverage may be specifically related to anxiety and depression, while the GHQ12 measure is slightly broader (Jackson, 2007). Further, we found more people were at least moderately anxious or depressed under the EQ5D measure than were at high risk of a CMD as measured under the GHQ12, suggesting that this measure is perhaps a more sensitive measure of mental health.

Similarly, green space coverage was not consistently related to mental health after adjusting for confounders, as with Nutsford et al. (2016). This is despite growing evidence that living within greener environments is positively connected to general mental health and wellbeing (see Hartig et al., 2014), as well as more specific factors associated with mental health, such as reduced stress levels (Cox et al., 2017c; Van den Berg et al., 2010), reduced rates of antidepressant prescriptions (Taylor et al., 2015), and increased psychological restoration (White et al., 2013b).

As previously discussed, it may therefore be that coastal environments are particularly important for mental health in comparison to green spaces. Similar conclusions were drawn in Hong Kong, where blue space visits were linked to mental health whilst visits to green spaces were not related (Garrett et al., 2019). Our result may also be due to the coarse measures used here to assess green space coverage. While the GLUD data are based on a high resolution cartographic database, it does not capture any measures of quality or accessibility that may be important modifiers of any health benefits of proximity to green space (Wheeler et al., 2015; Markevych et al., 2017). However, a relationship with self-reported health has previously been detected using

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a similar measure elsewhere (Mitchell and Popham, 2007).

4.1. Limitations and future work

Beyond the potential limitations associated with self-reporting health (e.g. Lee and Dugan, 2015), the cross-sectional nature of our study means the results should be interpreted cautiously before making generalisations about a causal relationship between coastal proximity and common mental disorders (Gascon et al., 2017). Future work should therefore examine the potential factors mediating this link, such as physical activity. More longitudinal and experimental research (e.g. White et al., 2013a; White et al., 2015; Annerstedt et al., 2012) is also needed to elucidate a causal relationship and determine whether living by the coast for an extended period remains beneficial for mental health, as well as if these coastal benefits are consistently greater than living in areas with more green space and freshwater coverage. Further, our measure does not capture variations in accessibility and quality which can relate to visit frequency and mental health benefits (Garrett et al., 2019; Wyles et al, 2016, 2017).

We were also not able to account for clustering at the LSOA level which may have resulted in smaller standard errors as we cannot account for some potential non-independence within the data. However, we have included LSOA level controls including additional nature exposures and IMD.

4.2. Conclusion

To summarise, we found that the relationship between coastal proximity and mental health was strongest for those urban adults in more deprived households. This builds on previous research investigating coastal proximity and health inequalities at the community level. Our results therefore add further evidence that the coast might act as a mental health resource, particularly for people living in more socioeconomically deprived circumstances. Ensuring access to these environments may therefore have a role to play in reduction of health inequalities (Allen and Balfour, 2014). At a time of increasing urbanisation, mental health disorders and degradation of coastal and marine environments, such research should be developed and translated to inform relevant environmental, planning and public health policies.

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Appendix A. Supplementary data

Supplementary data to this article can be found online at https://

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